# 8. Assessment of the Flathead Sole Stock in the Gulf of Alaska 

By<br>Carey R. McGilliard ${ }^{1}$ and Wayne Palsson ${ }^{2}$<br>${ }^{1}$ Resource Ecology and Fisheries Management Division<br>${ }^{2}$ Resource Assessment and Conservation Engineering Division<br>Alaska Fisheries Science Center<br>National Marine Fisheries Service<br>National Oceanic and Atmospheric Administration<br>7600 Sand Point Way NE., Seattle, WA 98115-6349

## Executive Summary

## Summary of Changes in Assessment Inputs

(1) 2014-2015 catch data were added to the model and 2013 catch was updated to include October to December catch in that year.
(2) 2014 and 2015 fishery length composition data were added to the model and 2013 fishery length composition data were updated to include October to December length data from that year.
(3) The 2015 bottom trawl survey biomass index was added to the model
(4) Survey length composition data for 2015 were added to the model
(5) Survey conditional age-at-length data for 2015 were added to the model
(6) Effective sample sizes for survey length composition data were changed to the number of hauls for which lengths were collected.
(7) The data sources were weighted using the harmonic mean of effective sample sizes, calculating effective sample sizes following the methods described in McAllister-Ianelli (1997), Appendix 2. Data weighting according to methods in Francis (2011) were used in the 2013 assessment.

## Summary of Changes in Assessment Methodology

No changes were made to the assessment methodology.

## Summary of Results

The key results of the assessment, based on the author's preferred model, are compared to the key results of the accepted 2014 update assessment in the table below.

|  | As estimated or <br> Quantity |  | As estimated or <br> recommended this year for: |  |
| :--- | ---: | ---: | ---: | ---: |
|  | 2015 |  | 2016 | $2016^{*}$ |

*Projections are based on estimated catches of $1,981.8 \mathrm{t}, 2,825 \mathrm{t}$, and $35,187 \mathrm{t}$ used in place of maximum permissible ABC for 2015, 2016, and 2017 respectively. The 2015 projected catch was calculated as the current catch as of October 10, 2015 added to the average October 10 - December 31 GOA catches over the 5 previous years. The 2016 projected catch was calculated as the average catch from 2010-2014. The 2017 projected catch was calculated as the projected maxABC for 2017.

The table below shows apportionment of the 2016 and 2017 ABCs and OFLs among areas, based on the proportion of survey biomass projected for each area in 2016 and 2017 estimated using the survey averaging random effects model developed by survey averaging working group.

| Quantity | Western | Central | West <br> Yakutat | Southeast | Total |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Area |  |  |  |  |  |
| Apportionment | $31.49 \%$ | $57.71 \%$ | $8.37 \%$ | $2.43 \%$ | $100.00 \%$ |
| 2016 ABC (t) | 11,027 | 20,211 | 2,930 | 852 | 35,020 |
| 2017 ABC (t) | 11,080 | 20,307 | 2,944 | 856 | 35,187 |

## Responses to SSC and Plan Team Comments on Assessments in General

GPT comment: The Teams recommend that the random effects survey smoothing model be used as a default for determining current survey biomass and apportionment among areas.

The random effects model was used in the current assessment to estimate 2016 and 2017 survey biomass, proportion of survey biomass expected in each management area in 2016 and 2017, and apportionment of ABCs according to these estimates of survey biomass in each area.

SSC comment: Of the options presented in the Joint Plan Teams minutes <for model numbering>, the SSC agrees that that Option 4 has several advantages and recommends that this Option be advanced next year. Under Option 4, analysts would number their models as follows: "Alpha-numeric model identifiers incorporating two-digit year labels of the form "yy.jx," where the digit after the decimal (" $j$ ") represents a major accepted model change and the alphabetic character ("x") represents a proposed model change (e.g., "12.1c" and "13.4a" might describe two models introduced in 2012 and 2013, respectively)". Differences between major and minor changes would be calculated based on "average difference in spawning biomass" (ADSB: see equation in Team Procedures) or as noted in sub-option c below, some other improvement to the model.
The above system for numbering models will be adopted for the next assessment, as recommended by the SSC.

## Responses to SSC and Plan Team Comments Specific to this Assessment

GPT, November 2013: The Team agreed with the author and recommends that the next assessment should include exploration of natural mortality and survey catchability. This effort might also include how selectivity is treated, and potentially place a prior on natural mortality based on maximum observed age. Additional model development should include estimation of a stock-specific ageing error matrix and exploration of strong patterns exhibited in early recruitment deviations.

Preliminary analyses were explored for the 2015 assessment and will be more fully explored in the future.
SSC, Dec. 2013: The SSC encourages development of a stock-specific aging error matrix and encourages exploration of the extreme patterns in early recruitment deviations.

The extreme patterns in early recruitment deviations were not evident in the 2015 assessment.

## Introduction

Flathead sole (Hippoglossoides elassodon) are distributed from northern California, off Point Reyes, northward along the west coast of North America and throughout the Gulf of Alaska (GOA) and the Eastern Bering Sea (EBS), the Kuril Islands, and possibly the Okhotsk Sea (Hart 1973). They occur primarily on mixed mud and sand bottoms (Norcross et al. 1997, McConnaughey and Smith 2000) in depths < 300 m (Stark and Clausen 1995). The flathead sole distribution overlaps with the similarappearing Bering flounder (Hippoglossoides robustus) in the northern half of the Bering Sea and the Sea of Okhotsk (Hart 1973), but not in the Gulf of Alaska.

Adults exhibit a benthic lifestyle and occupy separate winter spawning and summertime feeding distributions on the EBS shelf and in the GOA. From over-winter grounds near the shelf margins, adults begin a migration onto the mid and outer continental shelf in April or May each year for feeding. The spawning period may range from as early as January but is known to occur in March and April, primarily in deeper waters near the margins of the continental shelf. Eggs are large ( 2.75 to 3.75 mm ) and females have egg counts ranging from about 72,000 ( 20 cm fish) to almost 600,000 ( 38 cm fish). Eggs hatch in 9 to 20 days depending on incubation temperatures within the range of 2.4 to $9.8^{\circ} \mathrm{C}$ and have been found in ichthyoplankton sampling on the southern portion of the BS shelf in April and May (Waldron 1981). Larvae absorb the yolk sac in 6 to 17 days, but the extent of their distribution is unknown. Nearshore sampling indicates that newly settled larvae are in the 40 to 50 mm size range (Norcross et al. 1996). Fifty percent of flathead sole females in the GOA are mature at 8.7 years, or at about 33 cm (Stark 2004). Juveniles less than age 2 have not been found with the adult population and probably remain in shallow nearshore nursery areas.

## Fishery

Flathead sole in the Gulf of Alaska are caught in a directed fishery using bottom trawl gear. Typically 25 or fewer shore-based catcher vessels from 58-125' participate in this fishery, as do 5 catcher-processor vessels ( $90-130$ '). Fishing seasons are driven by seasonal halibut PSC apportionments, with approximately 7 months of fishing occurring between January and November. Catches of flathead sole occur almost entirely in the Western and Central management areas in the gulf (statistical areas 610 and $620+630$, respectively,Table 1). Recruitment to the fishery begins at about age 3.

Historically, catches of flathead sole have exhibited decadal-scale trends (Table 1, Figure 1). From a high of $\sim 2000 \mathrm{t}$ in 1980, annual catches declined steadily to a low of $\sim 150 t$ in 1986 but thereupon increased steadily, reaching a high of $\sim 3100 t$ in 1996. Catches subsequently declined over the next three years, reaching a low of $\sim 900 t$ in 1999, followed by an increasing trend through 2010, when the catch reached its highest level ever ( $3,854 \mathrm{t}$ ). Catch in 2014 was 2,556 t. 2015 closures of the flathead sole fishery are shown in Table 3.

Based on observer data, the majority of the flathead sole catch in the Gulf of Alaska is taken in the Shelikof Strait and on the Albatross Bank near Kodiak Island, as well as near Unimak Island (Stockhausen 2011). Previously, most of the catch is taken in the first and second quarters of the year (Stockhausen 2011).

Annual catches of flathead sole have been well below TACs in recent years (
Table 2), although the population appears to be capable of supporting higher exploitation rates. Limits on flathead sole catches are driven by restrictions on halibut PSC, not by attainment of the TAC (Stockhausen 2011).
See Stockhausen (2011) for a description of the management history of flathead sole.

## Data

The following table specifies the source, type, and years of all data included in the assessment models.

| Source | Type | Years |
| :--- | :--- | :--- |
| Fishery | Catch biomass | 1978-2015 (through October 10, 2015) |
| Fishery | Catch length composition | 1989-1999, 2001-2007, 2009-2015 |
| GOA survey bottom <br> trawl | Survey biomass | Triennial: 1984-1999, Biennial: 2001-2015 |
| GOA survey bottom <br> trawl | Catch length composition | Triennial: 1984-1999, Biennial: 2001-2015 |
| GOA survey bottom <br> trawl | Catch age composition, <br> conditioned on length | Triennial: 1984-1999, Biennial: 2001-2013 |

## Fishery:

## Catch Biomass

The assessment included catch data from 1978 to October 10, 2015 (Figure 1, Table 1). Catches of flathead sole occur almost entirely in the Western and Central management areas in the GOA (statistical areas 610 and $620+630$, respectively, Table 1).

## Catch Size Composition

Fishery length composition data were included in 2cm bins from 6-56cm in 1989-1999, 2001-2007, and 2009-2015; data were omitted in years where there were less than 15 hauls that included measured flathead sole (1982-1988 2000, 2008). The number of hauls were used as the relative effective sample size. Fishery length composition data were voluminous and can be accessed at http://www.afsc.noaa.gov/REFM/Docs/2015/GOA_Flathead_Composition_Data_And_SampleSize.xlsx.

## Survey:

## Biomass and Numerical Abundance

Survey biomass estimates originate from a cooperative bottom trawl survey conducted by the U.S. and Japan in 1984 and 1987 and a U.S. bottom trawl survey conducted by the Alaska Fisheries Science Center Resource Assessment and Conservation Engineering (RACE) Division thereafter. Calculations for final survey biomass and variance estimates are fully described in Wakabayashi et al. (1985). Depths 0-500 meters were fully covered in each survey and occurrence of flathead at depths greater than 500 meters is rare. The survey excluded the eastern region of the Gulf of Alaska (the Yakutat and Southeastern areas) in 2001 (

Table 4 and Table 5). As for previous assessments, the availability of the survey biomass in 2001 was assumed to be 0.9 to account for the biomass in the eastern region of the Gulf. The total survey biomass estimates and CVs that were used in the assessment are listed in (Table 5).
Figure 2 shows maps of survey CPUE in the GOA for the 2011, 2013, and 2015 surveys; survey CPUE in all three years was highest in the Central and Western GOA.

## Survey Size and Age Composition

Sex-specific survey length composition data as well as age frequencies of fish by length (conditional age-at-length) were used in the assessment and can be found at
http://www.afsc.noaa.gov/REFM/Docs/2015/GOA_Flathead_Composition_Data_And_SampleSize.xlsx, along with corresponding sample sizes used in the assessment. There are several advantages to using conditional age-at-length data. The approach preserves information on the relationship between length and age and provides information on variability in length-at-age such that growth parameters and variability in growth can be estimated within the model. In addition, the approach resolves the issue of double-counting individual fish when using both length- and age-composition data (as length-composition data are used to calculate the marginal age compositions). See Stewart (2005) for an additional example of the use of conditional age-at-length data in fishery stock assessments.

## Analytic Approach

## Model Structure

## Tier 3 Model

The assessment was a split sex, age-structured statistical catch-at-age model implemented in Stock Synthesis version 3.24 (SS3) using a maximum likelihood approach. SS3 equations can be found in Methot and Wetzel (2013) and further technical documentation is outlined in Methot (2009). Before 2013 assessments were conducted using an ADMB-based, split-sex, age-structured population dynamics model (Stockhausen 2011). A benchmark assessment was conducted in 2013 in SS3 (McGilliard et al. 2013). Briefly, the current assessment model covers 1955-2015. Age classes included in the model run from age 0 to 29. Age at recruitment was set at 0 years in the model. The oldest age class in the model, age 29, serves as a plus group. Survey catchability was fixed at 1.0.

## Fishery and Survey Selectivity

The fishery and survey selectivity curves were estimated using sex-specific, age-based double-normal functions without a descending limb (instead of a logistic function as previously used). The SS3 modeling framework does not currently include the option of estimating sex-specific, age-based logistic selectivity where both male and female selectivity maintain a logistic shape (as was used in the previous assessment model). Therefore, the double-normal curve without a descending limb was the closest match to the selectivity formulation used in the 2011 model (McGilliard et al. 2013). Length-based, sex-specific, logistic fishery and survey selectivity were implemented as sensitivity analyses in the 2013 assessment model runs (McGilliard et al. 2013). Length-based formulations for fishery and survey selectivity were not used in final model runs because the age-based selectivity curves derived from using length-based curves showed that the oldest fish were not selected, effectively lowering survey catchability and suggesting that the fishery fails to catch the oldest, largest fish. Fits to data were similar for length- and age-based asymptotic survey selectivity curves. Sensitivity analyses assuming dome-shaped fishery or survey selectivity failed to improve model fits to the data.

## Conditional Age-at-Length

A conditional age-at-length approach was used: expected age composition within each length bin was fit to age data conditioned on length (conditional age-at-length) in the objective function, rather than fitting the expected marginal age-composition to age data (which are typically calculated as a function of the conditional age-at-length data and the length-composition data). This approach provides the information necessary to estimate growth curves and variability about mean growth within the assessment model. In addition, the approach allows for all of the length and age-composition information to be used in the assessment without double-counting each sample. The von-Bertalanffy growth curve and variability in the length-at-age relationship were evaluated within the model using the conditional age-at-length approach.

## Data Weighting

In the 2013 assessment, the assumptions about data-weighting were re-evaluated using a more formal approach for assessing variability in mean proportions-at-age and proportions-at-length (Francis, 2011). To account for process error (e.g. variance in selectivities among years), the relative weights for length or age composition data (lambdas) were adjusted according to the method described in Francis (2011), which accounts for correlations in length- and age-composition data (data-weighting method number T3.4 was used). The 2013 assessment used weights calculated using the Francis (2011) method, but the weights for the fishery length-composition data were increased slightly to improve model stability.

In the current assessment, the method described in Francis (2011) was not used because of concerns raised about its use when using conditional age-at-length data. The effective sample size for length composition data was changed to the number of hauls (Volstad and Pennington 1994). The McAllisterIanelli method for weighting among data sources was used in the current assessment (McAllister and Ianelli 1997).

## Ageing Error Matrix

Ageing uncertainty was incorporated into the model using the ageing error matrix calculated from Bering Sea/Aleutian Islands (BSAI) flathead sole ageing data and used in the most recent accepted BSAI flathead sole assessment (McGilliard et al. 2014). SS3 accommodates the specification of ageing error bias and imprecision, while the previous assessment model framework did not. Future assessments should estimate ageing error matrices for GOA flathead sole using GOA age-read data. BSAI and GOA flathead sole are aged by the same individuals using the same techniques and ageing error is expected to be very similar. Assuming perfect age-reading of GOA flathead sole otoliths is thought to be an inferior assumption to using estimates of ageing error from the BSAI flathead sole population. The BSAI data was used in the current assessment, and will be replaced with GOA data when fully analyzed GOA ageing error data are available.

## Recruitment Deviations

Recruitment deviations for the period 1955-1983 were estimated as "early-period" recruits separately from "main-period" recruits (1984-2012) such that the vector of recruits for each period had a sum-tozero constraint, rather than forcing a sum-to-zero constraint across all recruitment deviations.

A bias adjustment factor was specified using the Methot and Taylor (2011) bias adjustment method. Recruitment deviations prior to the start of composition data and in the most recent years in the timeseries are less informed than in the middle of the time-series. This creates a bias in the estimation of recruitment deviations and mean recruitment that is corrected using methods described in Methot and Taylor (2011).

## Model structures considered in this year's assessment

One model is presented as the current, base case 2015 assessment model for GOA flathead sole (2015 Model). The proposed model structure is very similar to the most recent (2013) accepted model for flathead sole except that the effective sample size for all length composition data is now equal to the number of hauls for which lengths were collected for each data source due to correlations within hauls, which was analyzed in Volstad and Pennington (1994). In addition, data were weighted using the McAllister-Ianelli data weighting method, as described above. In addition, the 2013 model is presented with no new updated data (2014 and 2015 data are not included), and the 2015 model with 2013 data (2014 and 2015 data are excluded) are presented for the purpose of comparison.

## Parameters Estimated Outside the Assessment Model

## Natural mortality

Male and female natural mortality were fixed and equal to 0.2 .

## Weight-Length Relationship

The following weight-length relationship used in the previous assessment (McGilliard et al. 2013) is used in the current assessment: $w_{L}=\alpha L^{\beta}$, where $\alpha=4.28 E-06$ and $\beta=3.2298$, length ( $L$ ) was measured in centimeters and weight ( $w$ ) was measured in kilograms.

## Maturity-at-Age

Maturity-at-age $\left(O_{a}\right)$ in the assessment was defined as $O_{a}=1 /\left(1+\gamma e^{\left(a-a_{50}\right)}\right)$, where the slope of the curve was $\gamma=-0.773$ and the age-at- $50 \%$-maturity was $a_{50}=8.74$. These values were used in the previous assessment and were estimated from a histological analysis of 180 samples of GOA flathead sole ovaries collected in the central Gulf of Alaska from January 1999 (Stark, 2004).

## Standard deviation of the Log of Recruitment ( $\sigma_{R}$ )

The standard deviation of the log of recruitment was not defined in previous assessments. Variability of the recruitment deviations that were estimated in previous flathead sole assessments was approximately
$\sigma_{R}=0.6$ and this value is used in the current assessment.

## Catchability

Catchability was assumed equal to 1 , as for previous flathead sole assessments.

## Select selectivity parameters

Selectivity parameter definitions and values for fixed parameters are shown in Table 6.

## Parameters Estimated Inside the Assessment Model

Parameters estimated within the assessment model were the log of unfished recruitment ( $R_{0}$ ), log-scale recruitment deviations, yearly fishing mortality, sex-specific parameters of the von-Bertalanffy growth curve, CV of length-at-age for ages 2 and 29, and selectivity parameters for the fishery and survey. The selectivity parameters are described in greater detail in Table 6.

## Results

## Model Evaluation

## Comparison among models

Figure 3-Figure 4 and Table 7-Table 10 compare the 2015 base case model with (1) a model with the same structure of the base case model, but only including that data that were included in the 2013 model, and (2) the 2013 model. Fits to the survey biomass index and resulting estimates of spawning stock biomass over time are similar among the three model runs in recent years (after 2000; Figure 3, Figure 4). Before 2000, the fits to survey biomass index and estimates of spawning stock biomass for 2015 model with and without new data were higher, indicating that differences in the fits can be attributed to changes to the effective sample sizes and methods for data weighting among data sources. The negative log likelihood component for the survey index improved slightly for the 2015 model run with data only up to 2013 ( $-\operatorname{lnL}=-16.23$ ) as compared to the 2013 model ( $-\operatorname{lnL}=-15.77$; Table 7). Estimations of recruitment deviations and resulting age-0 recruitment are very similar among models, with the exception of the two most recent years, where there is little information to inform estimates (Figure 5, Figure 6). Estimates of growth parameters, unfished recruitment, and survey selectivity were very similar among models (Table 8, Table 10). Estimates of the age at which peak fishery selectivity was reached and the width of the ascending limb of the fishery selectivity curve were smaller for both models run with the 2015 model structure, indicating that changes in estimates of selectivity were due to changes in effective sample size and data weighting methods (Table 9). The 2013 model imposed a constraint on fishery selectivity such that peak female selectivity was reached by age 16 . Without the constraint on peak female selectivity, the model estimated an asymptotic fishery selectivity curve that did not reach a selectivity of 1 (McGilliard et al. 2013). The base case 2015 model and the 2015 model with data up to 2013 estimate peak female selectivity at age 13.08 and 13.25 , respectively, without a constraint (Table 9). The 2015 model was chosen because the approach to specifying effective sample sizes and methods for the relative weighting of data has a scientific basis and avoids issues that have been encountered (and are still being researched) about using the Francis (2011) data weighting methods with a conditional age-at-length approach. In addition, the 2015 model without new data fit the survey biomass index slightly better than the 2013 model and the 2015 model does not require a constraint peak female fishery selectivity.

## The 2015 Base Case Model

The estimated fishery and survey selectivity curves for the 2015 base case model are shown in Figure 7. Although selectivity curves for males and females are similar, it is puzzling that males would be selected at slightly younger ages than females, given that they grow more slowly than females (Figure 8). Future work will explore potential causes for this result. One constraint in the current assessment is that natural mortality is fixed at the same value for both males and females. Furthermore, natural mortality and catchability are both fixed in the assessment.

Fits to fishery and survey selectivity, aggregated over years are shown in Figure 9. These aggregated fits show that the model predicted slightly more females length $40-45 \mathrm{~cm}$ in the fishery than were observed. In addition, the model predicted that more $25-30 \mathrm{~cm}$ females in the survey than were observed and fewer females in the $32-40 \mathrm{~cm}$ range than were observed in the survey. Similarly, the model predicted slightly fewer $30-32 \mathrm{~cm}$ males and in the survey and slightly more $34-40 \mathrm{~cm}$ males in the survey than were observed. Overall, however, model fits to the length composition data, aggregated over years were fairly reasonable. Figure 10 -Figure 12 show fits to yearly fishery and survey length composition data. Fits to fishery length composition data were particularly poor in 1990; fishery selectivity appears to have been quite different in that year. Fits to survey length composition data were poor in 1984, 1987, and 1990. Survey methods in 1984 and 1987 differed from the current protocol and we would expect differences in fits in these years (McGilliard 2013).

Figure 13-Figure 16 show model fits to the mean age at each length and corresponding estimated and observed standard deviations about mean age-at-length and show that the model fits growth data reasonably well. Observed standard deviations are expected to differ from estimated standard deviations about the age-at-length for older ages and larger size bins due to low sample size. Figure 17-Figure 19 show pearson residuals in age-at-length model fits. One very large residual occurs in 1999, but otherwise, the pearson residuals are relatively small.

## Time Series Results

Time series results are shown in Table 13-Table 14 and Figure 20-Figure 21. A time series of number-atage is available at http://www.afsc.noaa.gov/REFM/Docs/2015/GOA_Flathead_TimeSeries_of_NumbersAtAge.xlsx. Age 3 recruitment, age 0 recruitment, and standard deviations of age 0 recruitment are presented in Table 14 for the current and previous assessments. Total biomass for ages $3+$, spawning stock biomass, and standard deviations of spawning stock biomass estimates for the previous and current assessments are presented in Table 13. Figure 20 shows spawning stock biomass estimates and corresponding asymptotic $95 \%$ confidence intervals. Figure 21 is a plot of biomass relative to $\mathrm{B}_{35 \%}$ and F relative to $\mathrm{F}_{35 \%}$ for each year in the time series, along with the OFL and ABC control rules.

## Retrospective Analyses

Spawning stock biomass, age 0 recruits, and recruitment deviations for retrospective analyses extending back 10 years are shown in Figure 22 and Figure 23. A retrospective pattern in spawning stock biomass extending back 8 years is evident, whereby each year of added data lowers the most current estimates by a small amount; runs removing 9 and 10 years of data do not follow this pattern (Figure 22). This retrospective pattern should be explored further in future analyses where alternative values and approaches for modeling catchability, natural mortality, and selectivity are explored.

## Harvest Recommendations

The reference fishing mortality rate for flathead sole is determined by the amount of reliable population information available (Amendment 56 of the Fishery Management Plan for the groundfish fishery of the Bering Sea/Aleutian Islands). Estimates of $\mathrm{F}_{40 \%}, \mathrm{~F}_{35 \%}$, and SPR $_{40 \%}$ were obtained from a spawner-per recruit analysis. Assuming that the average recruitment from the 1983-2012 year classes estimated in this assessment represents a reliable estimate of equilibrium recruitment, then an estimate of $\mathrm{B}_{40 \%}$ is calculated as the product of SPR $_{40 \%}$ times the equilibrium number of recruits. Since reliable estimates of the 2013 spawning biomass (B), $\mathrm{B}_{40 \%}, \mathrm{~F}_{40 \%}$, and $\mathrm{F}_{35 \%}$ exist and $\mathrm{B}>\mathrm{B}_{40 \%}$, the flathead sole reference fishing mortality is defined in Tier 3a. For this tier, $\mathrm{F}_{\text {ABC }}$ is constrained to be $\leq \mathrm{F}_{40 \%}$, and $\mathrm{F}_{\text {OFL }}$ is defined to be $\mathrm{F}_{35 \%}$. The values of these quantities are:

| SSB 2016 | 82,375 |
| :--- | :--- |
| $B_{40 \%}$ | 36,866 |
| $F_{40 \%}$ | 0.32 |
| $\operatorname{maxFabc}$ | 0.32 |
| $B_{35 \%}$ | 32,258 |
| $F_{35 \%}$ | 0.40 |
| $F_{\text {OFL }}$ | 0.40 |

Because the flathead sole stock has not been overfished in recent years and the stock biomass is relatively high, it is not recommended to adjust $\mathrm{F}_{\mathrm{ABC}}$ downward from its upper bound.

A standard set of projections is required for each stock managed under Tiers 1, 2, or 3 of Amendment 56. This set of projections encompasses seven harvest scenarios designed to satisfy the requirements of Amendment 56, the National Environmental Policy Act, and the Magnuson-Stevens Fishery Conservation and Management Act (MSFCMA). For each scenario, the projections begin with the vector of 2015 numbers at age estimated in the assessment. This vector is then projected forward to the beginning of 2016 using the schedules of natural mortality and selectivity described in the assessment and the best available estimate of total (year-end) catch for 2015. In each subsequent year, the fishing mortality rate is prescribed on the basis of the spawning biomass in that year and the respective harvest scenario. In each year, recruitment is drawn from an inverse Gaussian distribution whose parameters consist of maximum likelihood estimates determined from recruitments estimated in the assessment. Spawning biomass is computed in each year based on the time of peak spawning and the maturity and weight schedules described in the assessment. Total catch is assumed to equal the catch associated with the respective harvest scenario in all years. This projection scheme is run 1000 times to obtain distributions of possible future stock sizes, fishing mortality rates, and catches.

Five of the seven standard scenarios will be used in an Environmental Assessment prepared in conjunction with the final SAFE. These five scenarios, which are designed to provide a range of harvest alternatives that are likely to bracket the final TAC for 2016 are as follow ("max $\mathrm{F}_{\text {ABC" }}$ refers to the maximum permissible value of $\mathrm{F}_{A B C}$ under Amendment 56):

Scenario 1: In all future years, $F$ is set equal to $\max F_{A B C}$. (Rationale: Historically, TAC has been constrained by ABC, so this scenario provides a likely upper limit on future TACs.)

Scenario 2: In all future years, $F$ is set equal to a constant fraction of $\max F_{A B C}$, where this fraction is equal to the ratio of the $F_{A B C}$ value for 2016 recommended in the assessment to the $\max _{A B C}$ for 2016. (Rationale: When $F_{A B C}$ is set at a value below $\max F_{A B C}$, it is often set at the value recommended in the stock assessment.)

Scenario 3: In all future years, $F$ is set equal to $50 \%$ of max $F_{A B C}$. (Rationale: This scenario provides a likely lower bound on $F_{A B C}$ that still allows future harvest rates to be adjusted downward when stocks fall below reference levels.)

Scenario 4: In all future years, $F$ is set equal to the 2011-2015 average $F$. (Rationale: For some stocks, TAC can be well below ABC, and recent average $F$ may provide a better indicator of $F_{T A C}$ than $F_{A B C}$.)

Scenario 5: In all future years, $F$ is set equal to zero. (Rationale: In extreme cases, TAC may be set at a level close to zero.) The recommended $F_{A B C}$ and the maximum $F_{A B C}$ are equivalent in this assessment, so scenarios 1 and 2 yield identical results.

The 12-year projections of the mean spawning stock biomass, fishing mortality, and catches for the five scenarios are shown in Table 15-Table 17.

Two other scenarios are needed to satisfy the MSFCMA's requirement to determine whether the flathead sole stock is currently in an overfished condition or is approaching an overfished condition. These two scenarios are as follows (for Tier 3 stocks, the MSY level is defined as B35\%):

Scenario 6: In all future years, $F$ is set equal to $F_{\text {oFL }}$. (Rationale: This scenario determines whether a stock is overfished. If the stock is expected to be above its MSY level in 2016, then the stock is not overfished.)

Scenario 7: In 2016 and 2017, $F$ is set equal to $\max F_{A B C}$, and in all subsequent years, $F$ is set equal to $F_{\text {OFL }}$. (Rationale: This scenario determines whether a stock is approaching an overfished condition. If the stock is expected to be above its MSY level in 2028 under this scenario, then the stock is not approaching an overfished condition.)

The results of these two scenarios indicate that the stock is not overfished and is not approaching an overfished condition. With regard to assessing the current stock level, the expected stock size in the year 2016 of scenario 6 is 82,375 , more than 2 times $B 35 \%$ ( $32,258 \mathrm{t}$ ). Thus the stock is not currently overfished. With regard to whether the stock is approaching an overfished condition, the expected spawning stock size in the year 2028 of scenario $7(34,031 \mathrm{t})$ is greater than $B 35 \%$; thus, the stock is not approaching an overfished condition.

## Area Allocation of Harvests

TAC's for flathead sole in the Gulf of Alaska are divided among four smaller management areas (Western, Central, West Yakutat and Southeast Outside). The area-specific ABC's for flathead sole in the GOA are divided up over the four management areas by applying the fraction of the survey biomass estimated for each area (relative to the total over all areas) in 2016 and 2017 from the survey averaging random effects model to the 2016 and 2017 ABC's. The area-specific allocations for 2016 and 2017 are:

| Quantity | Western | Central | West <br> Yakutat | Southeast | Total |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Area |  |  |  |  |  |
| Apportionment | $31.49 \%$ | $57.71 \%$ | $8.37 \%$ | $2.43 \%$ | $100.00 \%$ |
| 2016 ABC (t) | 11,027 | 20,211 | 2,930 | 852 | 35,020 |
| $2017 \mathrm{ABC}(\mathrm{t})$ | 11,080 | 20,307 | 2,944 | 856 | 35,187 |

## Ecosystem Considerations

## Ecosystem Effects on the Stock

## Prey availability/abundance trends

Based on results from an ecosystem model for the Gulf of Alaska (Aydin et al., 2007), flathead sole in the Gulf of Alaska occupy an intermediate trophic level as both juvenile and adults (Figure 24, Figure 25). Pandalid shrimp and brittle stars were the most important prey for adult flathead sole in the Gulf of Alaska (64\% by weight in sampled stomachs; Yang and Nelson, 2000; Figure 24, Figure 26), while euphausids and mysids constituted the most important prey items for juvenile flathead sole (Figure 25, Figure 27). Other major prey items included polychaetes, mollusks, bivalves and hermit crabs for both juveniles and adults. Commercially important species that were consumed included age-0 Tanner crab (3\%) and age-0 walleye pollock (<0.5\% by weight). Little to no information is available to assess trends in abundance for the major benthic prey species of flathead sole.

## Predator population trends

Important predators on flathead sole include arrowtooth flounder, walleye pollock, Pacific cod, and other groundfish (Figure 24, Figure 25). Pacific cod and Pacific halibut are the major predators on adults, while arrowtooth flounder, sculpins, walleye pollock and Pacific cod are the major predators on juveniles. The
flatfish-directed fishery constitutes the third-largest known source of mortality on flathead sole adults. However, the largest component of mortality on adults is unexplained.

## Fishery Effects on the Ecosystem

Non-target catch in the directed GOA flathead sole fishery are shown in Table 18. Prohibited species catch in the directed GOA flathead sole fishery are shown in Table 19. Historically, the flathead sole fishery has caught a high proportion of the brittlestar, eelpouts, gunnels, polychaetes, and Stichaeidae in some years. In 2014 and 2015, proportion of non-target species caught in the flathead sole fishery ranged from 0 to $32 \%$ ( $32 \%$ of Pandalid shrimp were caught in the flathead sole fishery in 2015). Prohibited species catch in the flathead sole fishery were $0-2 \%$ of the prohibited species catch of each of these species in 2014 and 2015.

## Data Gaps and Research Priorities

The 2013 and 2015 stock assessments incorporated ageing error by using an existing ageing error matrix for BSAI flathead sole. A priority for future assessments is to analyze ageing error data for GOA flathead sole using methods described in Punt et al. (2008) and to incorporate a resulting ageing error matrix into the assessment. Future analyses should explore the relationship between natural mortality and catchability in the model, alternative parameter values, and the effects of these parameters on estimation of selectivity and other parameters. The assessment would benefit from an exploration of ways to better account for scientific uncertainty, especially uncertainty associated with parameters that are currently fixed in the model.

## Literature Cited

Aydin, K., S. Gaichas, I. Ortiz, D. Kinzey, and N. Friday. 2007. A comparison of the Bering Sea, Gulf of Alaska, and Aleutian Islands large marine ecosystems through food web modeling. NOAA Tech Memo. NMFS-AFSC-178. 298 p.
Francis, R. I. C. C. (2011). Data weighting in statistical fisheries stock assessment models. Canadian Journal of Fisheries and Aquatic Sciences, 68, 1124-1138.
Hart, J.L. 1973. Pacific fishes of Canada. Fish. Res. Board Canada, Bull. No. 180. 740 p.
McAllister, M.K. and Ianelli, J.N. 1997. Bayesian stock assessment using catch-age data and the sampling -importance resampling algorithm. Can. J. Fish. Aquat. Sci. 54: 284-300.
McConnaughey, R.A. and K.R. Smith. 2000. Associations between flatfish abundance and surficial sediments in the eastern Bering Sea. Can. J. Fish. Aquat. Sci. 57:2410-2419.
McGilliard, C.R., Palsson, W., Stockhausen, W., and Ianelli, J. 2013. 8. Assessment of the Flathead Sole Stock in the Gulf of Alaska. In Stock Assessment and Fishery Evaluation Document for Groundfish Resources in the Gulf of Alaska as Projected for 2013. pp. 611-756.

Methot, R. D., \& Wetzel, C. R. 2013. Stock synthesis: A biological and statistical framework for fish stock assessment and fishery management. Fisheries Research, 142, 86-99.
Methot, R. D., \& Taylor, I. G. 2011. Adjusting for bias due to variability of estimated recruitments in fishery assessment models. Can. J. Fish. Aquat. Res., 68(10): 1744-1760. Doi: 10.1139/f2011092.

Methot, R. D. (2009). User manual for stock synthesis, model version 3.04b. NOAA Fisheries, Seattle, WA.

Norcross, B.L., B.A. Holladay, S.C. Dressel and M. Frandsen. 1996. Recruitment of juvenile flatfishes in Alaska: habitat preference near Kodiak Island. U. Alaska Coastal Marine Institute, OCS Study MMS 96-0003, Vol. 1.

Norcross, B.L., F-J. Mueter and B.A. Holladay. 1997. Habitat models for juvenile pleuronectids around Kodiak Island, Alaska. Fish. Bull. 95:504-520.

Stark, J.W. 2004. A comparison of the maturation and growth of female flathead sole in the central Gulf of Alaska and south-eastern Bering Sea. J. Fish. Biol. 64:876-889.
Stark, J.W. and D.M. Clausen. 1995. Data report: 1990 Gulf of Alaska bottom trawl survey. NOAA Tech, Memo. NMFS-AFSC-49.

Stockhausen, W.T., M.E. Wilkins and M.H. Martin. 2011. Chapter 8. Assessment of the Flathead Sole Stock in the Gulf of Alaska. In Stock Assessment and Fishery Evaluation Document for Groundfish Resources in the Gulf of Alaska as Projected for 2011. pp. 753-820. North Pacific Fishery Management Council, P.O. Box 103136, Anchorage AK 99510.
Waldron, K.D. 1981. Ichthyoplankton. In D.W. Hood and J.A. Calder (Ed.s), The eastern Bering Sea shelf: Oceanography and resources, Vol. 1, p. 471-493. U.S. Dep. Commer., NOAA, Off. Mar. Poll, Assess., U.S. Gov. Print. Off., Wash. DC.

Yang, M.-S. and M.W. Nelson. 2000. Food habits of the commercially important groundfishes in the Gulf of Alaska in 1990, 1993, and 1996. U.S. Dep. Commer., NOAA Tech. Memo. NMFS-AFSC-112, 174 p.

## Tables

Table 1. Total and regional annual catch of GOA flathead sole through October 10, 2015.

| Year | Total <br> Catch | Western <br> Gulf | Central <br> Gulf | Eastern <br> Gulf |
| :---: | :---: | :---: | :---: | :---: |
| 1978 | 452 |  |  |  |
| 1979 | 165 |  |  |  |
| 1980 | 2,068 |  |  |  |
| 1981 | 1,070 |  |  |  |
| 1982 | 1,368 |  |  |  |
| 1983 | 1,080 |  |  |  |
| 1984 | 549 |  |  |  |
| 1985 | 320 |  |  |  |
| 1986 | 147 |  |  |  |
| 1987 | 151 |  |  |  |
| 1988 | 520 |  |  |  |
| 1989 | 747 |  |  |  |
| 1990 | 1,447 |  |  |  |
| 1991 | 1,237 | 199 | 1,036 | 2.1 |
| 1992 | 2,315 | 355 | 1,947 | 12.7 |
| 1993 | 2,824 | 581 | 2,242 | 0.0 |
| 1994 | 2,525 | 499 | 2,013 | 0.0 |
| 1995 | 2,180 | 589 | 1,563 | 28.0 |
| 1996 | 3,073 | 807 | 2,166 | 100.3 |
| 1997 | 2,441 | 449 | 1,934 | 0.0 |
| 1998 | 1,731 | 556 | 1,168 | 0.0 |
| 1999 | 897 | 186 | 687 | 24.6 |
| 2000 | 1,548 | 259 | 1,274 | 0.0 |
| 2001 | 1,912 | 600 | 1,311 | 0.0 |
| 2002 | 2,146 | 420 | 1,725 | 0.0 |
| 2003 | 2,459 | 525 | 1,934 | 0.1 |
| 2004 | 2,398 | 828 | 1,571 | 0.0 |
| 2005 | 2,552 | 611 | 1,941 |  |
| 2006 | 3,142 | 462 | 2,679 | 0.9 |
| 2007 | 3,130 | 666 | 2,462 | 2.2 |
| 2008 | 3,446 | 297 | 3,149 | 0.0 |
| 2009 | 3,663 | 303 | 3,359 | 1.0 |
| 2010 | 3,854 | 462 | 3,392 | 0.5 |
| 2011 | 2,729 | 393 | 2,336 | 0.3 |
| 2012 | 2,166 | 277 | 1,890 | 0.2 |
| 2013 | 2,817 | 588 | 2,228 | 0.2 |
| 2014 | 2,556 | 219 | 2,336 | 0.9 |
| 2015 | 1,765 | 188 | 1,577 | 0.6 |
|  |  |  |  |  |

Table 2. Historical OFLs, ABCs, TACs, total catch, and percent of catch that was retained.

| Year | OFL | ABC | TAC | Total <br> Catch | $\%$ <br> Retained |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 1995 | 31,557 | 28,790 | 9,740 | 2,180 |  |
| 1996 | 31,557 | 52,270 | 9,740 | 3,073 |  |
| 1997 | 34,010 | 26,110 | 9,040 | 2,441 |  |
| 1998 | 34,010 | 26,110 | 9,040 | 1,731 |  |
| 1999 | 34,010 | 26,010 | 9,040 | 897.32 |  |
| 2000 | 34,210 | 26,270 | 9,060 | 1,548 |  |
| 2001 | 34,210 | 26,270 | 9,060 | 1,912 |  |
| 2002 | 29,530 | 22,690 | 9,280 | 2,146 |  |
| 2003 | 51,560 | 41,390 | 11,150 | 2,459 | 88 |
| 2004 | 64,750 | 51,270 | 10,880 | 2,398 | 80 |
| 2005 | 56,500 | 45,100 | 10,390 | 2,552 | 87 |
| 2006 | 47,003 | 37,820 | 9,077 | 3,142 | 89 |
| 2007 | 48,658 | 39,110 | 9,148 | 3,130 | 89 |
| 2008 | 55,787 | 44,735 | 11,054 | 3,446 | 90 |
| 2009 | 57,911 | 46,464 | 11,181 | 3,663 | 96 |
| 2010 | 59,295 | 47,422 | 10,411 | 3,854 | 95 |
| 2011 | 61,412 | 49,133 | 10,587 | 2,729 | 97 |
| 2012 | 59,380 | 47,407 | 30,319 | 2,166 | 92 |
| 2013 | 61,036 | 48,738 | 30,496 | 2,817 | 87 |
| 2014 | 50,664 | 41,231 | 27,746 | 2,556 | 98 |
| 2015 | 50792 | 41349 | 27756 | 1,765 | 98 |

Table 3. GOA flathead sole fishery closures in 2015

| Sub-Area | Program | Status | Reason | Effective <br> Date |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \hline \text { GOA - Central } \\ & 620 / 630 \end{aligned}$ | All | Bycatch | Regulations | 01-Jan |
| GOA - Western $610$ | All | Bycatch | Regulations | 01-Jan |
| GOA - Central 620/630 | All | Open | Regulations | 20-Jan |
| $\begin{aligned} & \text { GOA - Western } \\ & 610 \end{aligned}$ | All | Open | Regulations | 20-Jan |
| West Yakutat 640 | All | Open | Regulations | 20-Jan |
| West Yakutat 640 | All | Bycatch | Regulations | 01-Jan |
| GOA - Central 620/630 | Catcher Vessel | Bycatch | Chinook <br> Salmon | 03-May |
| $\begin{aligned} & \text { GOA - Western } \\ & 610 \end{aligned}$ | Catcher Vessel | Bycatch | Chinook <br> Salmon | 03-May |
| GOA - Central 620/630 | Catcher Vessel | Open | Regulations | 10-Aug |
| $\begin{aligned} & \text { GOA - Western } \\ & 610 \end{aligned}$ | Catcher Vessel | Open | Regulations | 10-Aug |

Table 4. Survey biomass by area and depth

|  |  | Depth (meters) |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1-100 | 101-200 | 201-300 | 301-500 | 501-700 | 701-1000 | Total |
| CENTRAL GOA |  | 960,073 | 826,531 | 102,448 | 69 | 0 | 0 | 1,889,121 |
|  | 1984 | 64,191 | 85,916 | 8,431 | 0 | 0 | 0 | 158,539 |
|  | 1987 | 64,607 | 38,880 | 9,962 | 36 | 0 | 0 | 113,483 |
|  | 1990 | 100,061 | 52,600 | 8,591 | 5 |  |  | 161,257 |
|  | 1993 | 64,289 | 40,912 | 8,775 | 0 |  |  | 113,976 |
|  | 1996 | 56,342 | 59,964 | 6,422 | 3 |  |  | 122,730 |
|  | 1999 | 95,624 | 40,352 | 3,366 | 14 | 0 | 0 | 139,356 |
|  | 2001 | 44,046 | 37,467 | 3,906 | 11 |  |  | 85,430 |
|  | 2003 | 84,916 | 76,161 | 9,775 | 0 | 0 |  | 170,852 |
|  | 2005 | 61,294 | 75,699 | 5,050 | 0 | 0 | 0 | 142,043 |
|  | 2007 | 72,109 | 95,906 | 9,627 | 0 | 0 | 0 | 177,641 |
|  | 2009 | 60,575 | 62,431 | 5,904 | 0 | 0 | 0 | 128,910 |
|  | 2011 | 66,969 | 50,067 | 11,391 | 0 | 0 |  | 128,428 |
|  | 2013 | 72,923 | 42,847 | 5,293 | 0 | 0 |  | 121,063 |
|  | 2015 | 52,128 | 67,331 | 5,955 | 0 | 0 | 0 | 125,414 |
| EASTERN GOA |  | 131,961 | 159,546 | 5,580 | 370 | 0 | 0 | 297,456 |
|  | 1984 | 21,029 | 24,596 | 74 | 4 | 0 | 0 | 45,703 |
|  | 1987 | 6,060 | 23,835 | 564 | 0 | 0 |  | 30,459 |
|  | 1990 | 11,041 | 11,010 | 991 | 17 |  |  | 23,059 |
|  | 1993 | 4,839 | 10,377 | 1,434 | 193 |  |  | 16,843 |
|  | 1996 | 10,773 | 4,607 | 674 | 6 |  |  | 16,059 |
|  | 1999 | 5,145 | 13,271 | 182 | 0 | 0 | 0 | 18,598 |
|  | 2003 | 7,790 | 11,542 | 56 | 0 | 0 |  | 19,388 |
|  | 2005 | 2,060 | 9,365 | 135 | 151 | 0 | 0 | 11,712 |
|  | 2007 | 9,050 | 16,196 | 154 | 0 | 0 | 0 | 25,400 |
|  | 2009 | 10,111 | 6,150 | 90 | 0 | 0 | 0 | 16,351 |
|  | 2011 | 19,801 | 10,785 | 577 | 0 | 0 |  | 31,162 |
|  | 2013 | 11,007 | 6,887 | 146 | 0 | 0 |  | 18,039 |
|  | 2015 | 13,257 | 10,924 | 503 | 0 | 0 | 0 | 24,684 |
| WESTERN GOA |  | 690,651 | 178,842 | 1,122 | 58 | 8 | 0 | 870,680 |
|  | 1984 | 33,754 | 11,279 | 66 | 1 | 0 | 0 | 45,100 |
|  | 1987 | 20,815 | 12,761 | 27 | 0 | 0 | 0 | 33,603 |
|  | 1990 | 45,913 | 12,696 | 131 | 0 |  |  | 58,740 |
|  | 1993 | 43,944 | 13,854 | 68 | 5 |  |  | 57,871 |
|  | 1996 | 52,543 | 13,974 | 174 | 41 |  |  | 66,732 |
|  | 1999 | 44,578 | 5,018 | 33 | 0 | 8 | 0 | 49,636 |
|  | 2001 | 49,387 | 18,667 | 100 | 11 |  |  | 68,164 |
|  | 2003 | 53,313 | 13,718 | 24 | 0 | 0 |  | 67,055 |
|  | 2005 | 51,541 | 7,805 | 112 | 0 | 0 | 0 | 59,458 |
|  | 2007 | 59,759 | 18,560 | 42 | 0 | 0 | 0 | 78,361 |
|  | 2009 | 68,139 | 11,814 | 163 | 0 | 0 | 0 | 80,115 |
|  | 2011 | 63,066 | 12,866 | 117 | 0 | 0 |  | 76,049 |
|  | 2013 | 52,263 | 9,841 | 28 | 0 | 0 |  | 62,131 |
|  | 2015 | 51,636 | 15,991 | 37 | 0 | 0 | 0 | 67,665 |

Table 5. Survey biomass estimates and CVs used in the assessment as an absolute index of abundance

| Year | Biomass <br> Estimate | CV |
| :---: | :---: | :---: |
| 1984 | 249,341 | 0.12 |
| 1987 | 177,546 | 0.11 |
| 1990 | 243,055 | 0.12 |
| 1993 | 188,690 | 0.13 |
| 1996 | 205,521 | 0.09 |
| 1999 | 207,590 | 0.12 |
| 2001 | 170,660 | 0.12 |
| 2003 | 257,294 | 0.08 |
| 2005 | 213,213 | 0.08 |
| 2007 | 281,402 | 0.08 |
| 2009 | 225,377 | 0.11 |
| 2011 | 235,639 | 0.09 |
| 2013 | 201,233 | 0.09 |
| 2015 | 217,763 | 0.08 |

Table 6. Configuration of fishery and survey age-based, sex-specific double-normal selectivity curves used in the assessment. A numeric value indicates the fixed value of a parameter. The asterisk denotes that the parameter was estimated, but constrained to be below age 16 (as for the accepted 2013 model). A " + " denotes that initial selectivity was fixed at zero for ages 0-2.

| Double-normal selectivity parameters | Fishery | Survey |
| :--- | :---: | :---: |
| Peak: beginning size for the plateau | Estimated* | Estimated |
| Width: width of plateau | 30 | 30 |
| Ascending width (log space) | Estimated | Estimated |
| Descending width (log space) 8 8 <br> Initial: selectivity at smallest length or age <br> bin $0^{+}$ $0^{+}$ <br> Final: selectivity at largest length or age bin 999 999 <br> Male Peak Offset Estimated Estimated <br> Male ascending width offset (log space) Estimated Estimated <br> Male descending width offset (log space) 0 0 <br> Male "Final" offset (transformation <br> required) <br> Male apical selectivity 0 0$\$ 1$0 |  |  |

Table 7. Likelihood components for the base case 2015 model, the base case model with new data removed (data are as for the 2013 model), and the 2013 model. Values for likelihood components for the 2015 base case model cannot be compared directly with the other two models. Only the value for the survey index likelihood component can be compared between the two models using data up to 2013 because effective sample sizes, data weights, and the estimation of the most recent recruitment deviations differ between models.

|  |  | 2015 <br> Model w/ |  |
| :---: | :---: | :---: | :---: |
| Likelihood | 2015 | 2013 | 2013 |
| Component | Model | Data | Model |
| TOTAL | 1,425 | 1,293 | 1,663 |
| Survey | -17.88 | -16.23 | -15.77 |
| Length_comp | 507 | 457 | 182 |
| Age_comp | 941 | 857 | 1,498 |
| Recruitment | -4.694 | -5.062 | -0.996 |

Table 8. Final parameter estimates of growth parameters and unfished recruitment with corresponding standard deviations for the 2015 base case model, the 2015 base case model with data up to 2013, and the 2013 model.

\left.|  |  |  | 2015 Model, 2013 |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Data |  |  |  |  |$\right)$

Table 9. Final fishery selectivity parameters for the 2015 base case model, the 2015 model with data up to 2013, and the 2013 model. "Est" refers to the estimated value and "Std. Dev" is the standard deviation of the estimate.

|  | 2015 Model |  | 2015 Model, 2013 Data |  | 2013 Model |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Double-normal selectivity parameters | Est | Std. <br> Dev. | Est | Std. <br> Dev. | Est | Std. <br> Dev. |
| Peak: beginning size for the plateau | 13.08 | 0.68 | 13.25 | 0.72 | 16.00 | 0.13 |
| Width: width of plateau | 30.00 | NA | 30.00 | NA | 30.00 | NA |
| Ascending width (log space) | 2.93 | 0.17 | 2.92 | 0.18 | 3.53 | 0.11 |
| Descending width (log space) | 8.00 | NA | 8.00 | NA | 8.00 | NA |
| Initial: selectivity at smallest length or age bin | -10 | NA | -10 | NA | -10 | NA |
| Final: selectivity at largest length or age bin | 999 | NA | 999 | NA | 999 | NA |
| Male Peak Offset | -0.94 | 0.49 | -1.05 | 0.51 | -1.68 | 1.77 |
| Male ascending width offset (log space) | -0.10 | 0.15 | -0.15 | 0.17 | -0.23 | 0.46 |
| Male descending width offset (log space) | 0.00 | NA | 0.00 | NA | 0.00 | NA |
| Male "Final" offset (transformation required) | 1.00 | NA | 1.00 | NA | 1.00 | NA |
| Male apical selectivity | 1.00 | NA | 1.00 | NA | 1.00 | NA |

Table 10. Final survey selectivity parameters for the 2015 base case model, the 2015 model with data up to 2013, and the 2013 model. "Est" refers to the estimated value and "Std. Dev" is the standard deviation of the estimate.

|  | 2015 Model |  | 2015 Model,2013 Data |  | 2013 Model |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Double-normal selectivity parameters | Est | Std. <br> Dev. | Est | Std. <br> Dev. | Est | Std. <br> Dev. |
| Peak: beginning size for the plateau (in cm) | 7.22 | 0.24 | 7.31 | 0.25 | 7.12 | 0.28 |
| Width: width of plateau | 30.00 | NA | 30.00 | NA | 30.00 | NA |
| Ascending width (log space) | 2.13 | 0.12 | 2.16 | 0.12 | 2.06 | 0.14 |
| Descending width (log space) | 8.00 | NA | 8.00 | NA | 8.00 | NA |
| bin | -10 | NA | -10 | NA | -10 | NA |
| Final: selectivity at largest length or age bin | 999 | NA | 999 | NA | 999 | NA |
| Male Peak Offset | -0.59 | 0.26 | -0.59 | 0.28 | -0.74 | 0.32 |
| Male ascending width offset (log space) | -0.26 | 0.15 | -0.24 | 0.16 | -0.32 | 0.18 |
| Male descending width offset (log space) | 0.00 | NA | 0.00 | NA | 0.00 | NA |
| Male "Final" offset (transformation required) | 0.00 | NA | 0.00 | NA | 0.00 | NA |
| Male apical selectivity | 1.00 | NA | 1.00 | NA | 1.00 | NA |

Table 11. Estimated yearly fishing mortality rates (rates are apical fishing mortality rates across ages) for the proposed 2015 model.

| Year | Fishing <br> Mortality | Std. <br> Dev. | Year | Fishing <br> Mortality | Std. <br> Dev. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Initial |  |  |  |  |  |
| F | 0.0069 | 0.0004 | 1998 | 0.0153 | 0.0010 |
| 1978 | 0.0052 | 0.0006 | 1999 | 0.0078 | 0.0005 |
| 1979 | 0.0020 | 0.0002 | 2000 | 0.0133 | 0.0008 |
| 1980 | 0.0260 | 0.0027 | 2001 | 0.0164 | 0.0010 |
| 1981 | 0.0141 | 0.0014 | 2002 | 0.0183 | 0.0011 |
| 1982 | 0.0183 | 0.0018 | 2003 | 0.0211 | 0.0013 |
| 1983 | 0.0143 | 0.0013 | 2004 | 0.0208 | 0.0012 |
| 1984 | 0.0069 | 0.0006 | 2005 | 0.0223 | 0.0013 |
| 1985 | 0.0037 | 0.0003 | 2006 | 0.0275 | 0.0016 |
| 1986 | 0.0015 | 0.0001 | 2007 | 0.0273 | 0.0017 |
| 1987 | 0.0014 | 0.0001 | 2008 | 0.0297 | 0.0019 |
| 1988 | 0.0045 | 0.0004 | 2009 | 0.0313 | 0.0020 |
| 1989 | 0.0062 | 0.0005 | 2010 | 0.0326 | 0.0021 |
| 1990 | 0.0119 | 0.0008 | 2011 | 0.0230 | 0.0015 |
| 1991 | 0.0102 | 0.0007 | 2012 | 0.0180 | 0.0011 |
| 1992 | 0.0192 | 0.0012 | 2013 | 0.0231 | 0.0015 |
| 1993 | 0.0237 | 0.0015 | 2014 | 0.0208 | 0.0014 |
| 1994 | 0.0215 | 0.0014 | 2015 | 0.0130 | 0.0009 |
| 1995 | 0.0188 | 0.0012 |  |  |  |
| 1996 | 0.0269 | 0.0017 |  |  |  |
| 1997 | 0.0216 | 0.0014 |  |  |  |

Table 12. Recruitment deviations and standard deviations for the proposed 2015 model.

| Year | Recruitment <br> Deviations | Std. <br> Dev. | Year | Recruitment <br> Deviations | Std. <br> Dev. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1955 | -0.131 | 0.564 | 1985 | -0.242 | 0.377 |
| 1956 | -0.156 | 0.558 | 1986 | -0.228 | 0.332 |
| 1957 | -0.185 | 0.551 | 1987 | -0.123 | 0.300 |
| 1958 | -0.219 | 0.543 | 1988 | -0.195 | 0.322 |
| 1959 | -0.257 | 0.535 | 1989 | 0.225 | 0.209 |
| 1960 | -0.300 | 0.526 | 1990 | -0.341 | 0.271 |
| 1961 | -0.348 | 0.516 | 1991 | -0.149 | 0.244 |
| 1962 | -0.400 | 0.506 | 1992 | 0.327 | 0.172 |
| 1963 | -0.454 | 0.496 | 1993 | -0.165 | 0.219 |
| 1964 | -0.510 | 0.487 | 1994 | -0.067 | 0.198 |
| 1965 | -0.562 | 0.478 | 1995 | -0.265 | 0.217 |
| 1966 | -0.613 | 0.469 | 1996 | -0.479 | 0.242 |
| 1967 | -0.666 | 0.460 | 1997 | 0.212 | 0.152 |
| 1968 | -0.722 | 0.452 | 1998 | -0.019 | 0.185 |
| 1969 | -0.780 | 0.444 | 1999 | 0.401 | 0.149 |
| 1970 | -0.835 | 0.437 | 2000 | -0.238 | 0.241 |
| 1971 | -0.873 | 0.432 | 2001 | 0.007 | 0.171 |
| 1972 | -0.882 | 0.429 | 2002 | -0.038 | 0.171 |
| 1973 | -0.848 | 0.431 | 2003 | 0.300 | 0.147 |
| 1974 | -0.754 | 0.438 | 2004 | -0.006 | 0.193 |
| 1975 | -0.560 | 0.458 | 2005 | 0.285 | 0.156 |
| 1976 | -0.177 | 0.517 | 2006 | -0.122 | 0.205 |
| 1977 | 0.852 | 0.311 | 2007 | -0.010 | 0.188 |
| 1978 | 0.092 | 0.483 | 2008 | -0.263 | 0.215 |
| 1979 | -0.277 | 0.427 | 2009 | 0.000 | 0.200 |
| 1980 | -0.116 | 0.357 | 2010 | 0.482 | 0.186 |
| 1981 | -0.098 | 0.356 | 2011 | 0.455 | 0.243 |
| 1982 | -0.082 | 0.367 | 2012 | 0.316 | 0.298 |
| 1983 | -0.043 | 0.373 |  |  |  |
| 1984 | -0.062 | 0.357 |  |  |  |
|  |  |  |  |  |  |

Table 13. Time series of total and spawning biomass and standard deviation of spawning biomass (Std_Dev) for the previous and proposed 2015 assessments.

| 2013 Assessment |  |  |  | 2015 Assessment |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Total Biomass (age 3+) | Spawning Biomass | Stdev_SPB | Total Biomass (age 3+) | Spawning Biomass | Stdev_SPB |
| 1978 | 269,959 | 51,926 | 5,349 | 277,139 | 58,089 | 6,159 |
| 1979 | 126,738 | 49,361 | 4,913 | 141,975 | 55,470 | 5,688 |
| 1980 | 125,801 | 47,308 | 4,504 | 140,348 | 53,318 | 5,234 |
| 1981 | 135,017 | 44,867 | 4,131 | 150,713 | 50,751 | 4,807 |
| 1982 | 145,957 | 44,019 | 3,806 | 162,748 | 49,778 | 4,424 |
| 1983 | 158,409 | 44,516 | 3,545 | 176,027 | 50,243 | 4,100 |
| 1984 | 169,804 | 47,103 | 3,370 | 187,764 | 52,985 | 3,864 |
| 1985 | 180,069 | 51,879 | 3,304 | 197,571 | 58,136 | 3,750 |
| 1986 | 188,930 | 57,830 | 3,347 | 205,660 | 64,542 | 3,771 |
| 1987 | 195,676 | 63,517 | 3,432 | 212,188 | 70,501 | 3,843 |
| 1988 | 200,541 | 67,904 | 3,477 | 217,348 | 74,843 | 3,856 |
| 1989 | 203,678 | 70,756 | 3,467 | 220,399 | 77,433 | 3,792 |
| 1990 | 204,544 | 72,470 | 3,422 | 221,360 | 78,873 | 3,685 |
| 1991 | 204,089 | 73,083 | 3,361 | 221,208 | 79,357 | 3,570 |
| 1992 | 202,641 | 72,992 | 3,293 | 219,546 | 79,543 | 3,462 |
| 1993 | 203,362 | 72,348 | 3,221 | 220,077 | 78,828 | 3,361 |
| 1994 | 202,816 | 71,365 | 3,147 | 218,587 | 77,623 | 3,260 |
| 1995 | 202,782 | 70,378 | 3,072 | 216,623 | 76,576 | 3,157 |
| 1996 | 206,051 | 69,971 | 3,000 | 217,713 | 75,944 | 3,059 |
| 1997 | 209,034 | 69,659 | 2,945 | 218,763 | 75,244 | 2,970 |
| 1998 | 211,821 | 70,224 | 2,907 | 218,974 | 75,216 | 2,894 |
| 1999 | 213,612 | 71,498 | 2,884 | 218,544 | 75,734 | 2,829 |
| 2000 | 213,109 | 73,417 | 2,873 | 216,628 | 76,770 | 2,773 |
| 2001 | 215,414 | 74,985 | 2,877 | 216,872 | 77,424 | 2,727 |
| 2002 | 217,217 | 75,985 | 2,880 | 216,677 | 77,572 | 2,683 |
| 2003 | 222,411 | 76,306 | 2,868 | 219,537 | 77,148 | 2,632 |
| 2004 | 225,341 | 76,200 | 2,839 | 221,399 | 76,372 | 2,575 |
| 2005 | 228,763 | 76,389 | 2,813 | 223,115 | 75,936 | 2,528 |
| 2006 | 231,545 | 77,226 | 2,818 | 224,310 | 76,121 | 2,513 |
| 2007 | 235,092 | 78,381 | 2,871 | 226,919 | 76,661 | 2,539 |
| 2008 | 237,259 | 79,679 | 2,959 | 228,563 | 77,474 | 2,595 |
| 2009 | 240,735 | 80,631 | 3,067 | 231,727 | 78,025 | 2,667 |
| 2010 | 241,844 | 81,282 | 3,197 | 233,324 | 78,367 | 2,754 |
| 2011 | 241,226 | 81,824 | 3,365 | 233,972 | 78,739 | 2,866 |
| 2012 | 238,297 | 82,867 | 3,570 | 232,367 | 79,826 | 3,006 |
| 2013 | 236,745 | 83,899 | 3,812 | 231,266 | 81,114 | 3,166 |
| 2014 | 252,361 | 84,058 | 0 | 233,760 | 81,718 | 3,334 |
| 2015 |  |  |  | 238,766 | 82,006 | 3,510 |
| 2016 |  |  |  | 265,088 | 82,375 | 0 |

Table 14. Time series of recruitment at ages 3 and 0 and standard deviation of age 0 recruits for the previous and proposed 2015 assessments.

|  | 2013 Assessment |  |  | 2015 Assessment |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Recruits <br> (Age 3) | Recruits <br> (Age 0) | Std. dev | Recruits <br> (Age 3) | Recruits <br> (Age 0) | Std. dev |
| 1978 | 100,774 | 358,535 | 165,729 | 106,393 | 368,484 | 177,506 |
| 1979 | 150,505 | 251,699 | 105,587 | 155,476 | 253,772 | 108,446 |
| 1980 | 365,437 | 289,323 | 103,307 | 433,871 | 297,149 | 105,949 |
| 1981 | 196,748 | 309,033 | 106,072 | 202,218 | 301,679 | 106,894 |
| 1982 | 138,123 | 292,494 | 102,141 | 139,268 | 305,519 | 112,206 |
| 1983 | 158,769 | 263,263 | 92,987 | 163,073 | 316,566 | 116,231 |
| 1984 | 169,587 | 265,842 | 91,170 | 165,559 | 309,669 | 111,617 |
| 1985 | 160,516 | 265,419 | 85,684 | 167,669 | 257,747 | 97,245 |
| 1986 | 144,479 | 231,152 | 71,352 | 173,734 | 260,632 | 86,860 |
| 1987 | 145,895 | 251,384 | 73,740 | 169,949 | 288,331 | 86,107 |
| 1988 | 145,664 | 272,819 | 82,111 | 141,454 | 267,445 | 87,185 |
| 1989 | 126,856 | 404,447 | 78,027 | 143,037 | 405,666 | 82,826 |
| 1990 | 137,958 | 208,573 | 57,071 | 158,238 | 229,639 | 62,810 |
| 1991 | 149,718 | 308,667 | 71,517 | 146,773 | 277,449 | 68,434 |
| 1992 | 221,949 | 477,306 | 75,450 | 222,629 | 445,008 | 74,920 |
| 1993 | 114,458 | 252,968 | 53,346 | 126,024 | 271,093 | 59,833 |
| 1994 | 169,384 | 343,279 | 56,150 | 152,259 | 298,176 | 59,033 |
| 1995 | 261,922 | 240,966 | 46,048 | 244,213 | 243,695 | 53,035 |
| 1996 | 138,818 | 172,439 | 39,645 | 148,772 | 196,069 | 48,182 |
| 1997 | 188,371 | 447,516 | 56,870 | 163,633 | 390,474 | 58,701 |
| 1998 | 132,230 | 330,472 | 53,743 | 133,736 | 310,045 | 58,382 |
| 1999 | 94,628 | 509,094 | 68,171 | 107,601 | 471,911 | 69,994 |
| 2000 | 245,590 | 208,070 | 54,735 | 214,291 | 249,058 | 60,890 |
| 2001 | 181,356 | 380,791 | 56,138 | 170,152 | 317,992 | 54,702 |
| 2002 | 279,375 | 330,704 | 53,962 | 258,981 | 304,039 | 53,283 |
| 2003 | 114,181 | 413,028 | 62,922 | 136,681 | 426,427 | 63,497 |
| 2004 | 208,962 | 320,105 | 59,460 | 174,510 | 314,108 | 62,031 |
| 2005 | 181,476 | 436,627 | 67,398 | 166,852 | 420,247 | 66,478 |
| 2006 | 226,652 | 254,330 | 52,683 | 234,017 | 279,669 | 58,466 |
| 2007 | 175,656 | 284,855 | 60,856 | 172,376 | 312,754 | 60,275 |
| 2008 | 239,595 | 248,675 | 59,976 | 230,623 | 242,828 | 53,751 |
| 2009 | 139,560 | 362,494 | 97,638 | 153,476 | 315,972 | 65,533 |
| 2010 | 156,309 | 536,437 | 178,348 | 171,631 | 511,681 | 98,931 |
| 2011 | 136,455 | 355,967 | 176,865 | 133,257 | 504,307 | 126,418 |
| 2012 | 198,917 | 362,445 | 15,778 | 173,400 | 445,553 | 137,487 |
| 2013 | 294,376 | 362,445 |  | 280,805 | 371,808 | 13,501 |
| 2014 |  |  |  | 276,755 | 371,808 | 13,501 |
| 2015 |  |  |  | 244,513 | 371,808 |  |
| Average | 177,535 | 322,324 |  | 183,103 | 329,639 |  |

Table 15. Projected spawning biomass for the seven harvest scenarios listed in the "Harvest Recommendations" section.

| Year | Scenario 1 | Scenario 2 | Scenario 3 | Scenario 4 | Scenario 5 | Scenario 6 | Scenario 7 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2015 | 82,007 | 82,007 | 82,007 | 82,007 | 82,007 | 82,007 | 82,007 |
| 2016 | 82,375 | 82,375 | 82,375 | 82,375 | 82,375 | 82,375 | 82,375 |
| 2017 | 82,690 | 82,690 | 82,690 | 82,690 | 82,690 | 63,484 | 67,189 |
| 2018 | 68,562 | 68,562 | 84,160 | 81,445 | 85,342 | 52,732 | 58,027 |
| 2019 | 60,299 | 60,299 | 86,617 | 81,584 | 88,867 | 47,401 | 50,806 |
| 2020 | 55,403 | 55,403 | 89,388 | 82,353 | 92,613 | 44,726 | 46,895 |
| 2021 | 51,753 | 51,753 | 91,564 | 82,761 | 95,691 | 42,625 | 43,986 |
| 2022 | 48,341 | 48,341 | 92,640 | 82,277 | 97,599 | 40,292 | 41,128 |
| 2023 | 45,159 | 45,159 | 92,694 | 80,998 | 98,401 | 37,927 | 38,425 |
| 2024 | 42,527 | 42,527 | 92,097 | 79,324 | 98,446 | 36,078 | 36,337 |
| 2025 | 40,596 | 40,596 | 91,198 | 77,611 | 98,075 | 34,951 | 35,070 |
| 2026 | 39,311 | 39,311 | 90,220 | 76,048 | 97,514 | 34,371 | 34,418 |
| 2027 | 38,512 | 38,512 | 89,277 | 74,705 | 96,892 | 34,114 | 34,129 |
| 2028 | 38,043 | 38,043 | 88,427 | 73,594 | 96,288 | 34,029 | 34,031 |

Table 16. Projected fishing mortality rates for the seven harvest scenarios listed in the "Harvest Recommendations" section.

| Year | Scenario 1 | Scenario 2 | Scenario 3 | Scenario 4 | Scenario 5 | Scenario 6 | Scenario 7 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2015 | 0.02 | 0.02 | 0.02 | 0.02 | 0.02 | 0.02 | 0.02 |
| 2016 | 0.02 | 0.02 | 0.02 | 0.02 | 0.02 | 0.40 | 0.32 |
| 2017 | 0.32 | 0.32 | 0.02 | 0.07 | 0.00 | 0.40 | 0.32 |
| 2018 | 0.32 | 0.32 | 0.02 | 0.07 | 0.00 | 0.40 | 0.40 |
| 2019 | 0.32 | 0.32 | 0.02 | 0.07 | 0.00 | 0.40 | 0.40 |
| 2020 | 0.32 | 0.32 | 0.02 | 0.07 | 0.00 | 0.40 | 0.40 |
| 2021 | 0.32 | 0.32 | 0.02 | 0.07 | 0.00 | 0.40 | 0.40 |
| 2022 | 0.32 | 0.32 | 0.02 | 0.07 | 0.00 | 0.40 | 0.40 |
| 2023 | 0.32 | 0.32 | 0.02 | 0.07 | 0.00 | 0.39 | 0.39 |
| 2024 | 0.32 | 0.32 | 0.02 | 0.07 | 0.00 | 0.38 | 0.38 |
| 2025 | 0.31 | 0.31 | 0.02 | 0.07 | 0.00 | 0.37 | 0.37 |
| 2026 | 0.31 | 0.31 | 0.02 | 0.07 | 0.00 | 0.37 | 0.37 |
| 2027 | 0.31 | 0.31 | 0.02 | 0.07 | 0.00 | 0.36 | 0.36 |
| 2028 | 0.31 | 0.31 | 0.02 | 0.07 | 0.00 | 0.36 | 0.36 |

Table 17. Projected catches for the seven harvest scenarios listed in the "Harvest Recommendations" section.

| Year | Scenario 1 | Scenario 2 | Scenario 3 | Scenario 4 | Scenario 5 | Scenario 6 | Scenario 7 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2015 | 1,982 | 1,982 | 1,982 | 1,982 | 1,982 | 1,982 | 1,982 |
| 2016 | 2,825 | 2,825 | 2,825 | 2,825 | 2,825 | 42,840 | 35,020 |
| 2017 | 35,187 | 35,187 | 2,447 | 8,088 | 0 | 32,951 | 28,489 |
| 2018 | 29,050 | 29,050 | 2,483 | 7,937 | 0 | 27,306 | 30,086 |
| 2019 | 25,392 | 25,392 | 2,547 | 7,915 | 0 | 24,417 | 26,203 |
| 2020 | 23,197 | 23,197 | 2,627 | 7,978 | 0 | 22,895 | 24,031 |
| 2021 | 21,606 | 21,606 | 2,702 | 8,040 | 0 | 21,735 | 22,447 |
| 2022 | 20,161 | 20,161 | 2,747 | 8,024 | 0 | 20,507 | 20,949 |
| 2023 | 18,809 | 18,809 | 2,752 | 7,906 | 0 | 19,021 | 19,354 |
| 2024 | 17,677 | 17,677 | 2,729 | 7,725 | 0 | 17,645 | 17,838 |
| 2025 | 16,797 | 16,797 | 2,695 | 7,536 | 0 | 16,775 | 16,867 |
| 2026 | 16,183 | 16,183 | 2,661 | 7,372 | 0 | 16,339 | 16,376 |
| 2027 | 15,800 | 15,800 | 2,630 | 7,237 | 0 | 16,155 | 16,165 |
| 2028 | 15,570 | 15,570 | 2,604 | 7,128 | 0 | 16,105 | 16,105 |

Table 18. Non-target catch in the directed GOA flathead sole fishery as a proportion of total weight of bycatch of each species. Conditional highlighting from white (lowest numbers) to green (highest numbers) is applied. No seabird bycatch was recorded in the GOA flathead sole fishery.

| Non-Target Species | 2003 | 2004 | 2005 | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 | 2013 | 2014 | 2015 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Benthic urochordata | 0.00 | 0.03 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.10 | 0.07 | 0.06 | 0.00 | 0.00 |
| Bivalves | 0.03 | 0.34 | 0.15 | 0.09 | 0.00 | 0.11 | 0.00 | 0.01 | 0.05 | 0.00 | 0.03 | 0.00 | 0.00 |
| Brittle star unidentified | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.84 | 0.01 | 0.00 | 0.00 | 0.11 | 0.00 |
| Capelin | 0.00 | 0.00 | 0.00 | 0.00 |  | 0.00 | 0.00 | 0.00 | 0.00 | 0.29 | 0.00 | 0.00 | 0.00 |
| Corals Bryozoans Unidentified | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 | 0.03 | 0.00 | 0.03 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| Dark Rockfish |  |  |  |  |  | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| Eelpouts | 0.52 | 0.07 | 0.04 | 0.18 | 0.12 | 0.02 | 0.00 | 0.94 | 0.24 | 0.06 | 0.00 | 0.00 | 0.10 |
| Eulachon | 0.07 | 0.00 | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.01 | 0.00 | 0.08 |
| Giant Grenadier | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| Greenlings | 0.01 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.01 | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 |
| Ratail Grenadier Unidentified | 0.01 | 0.01 | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.03 | 0.00 | 0.00 | 0.00 |
| Gunnels | 0.00 |  |  | 1.00 |  | 0.24 |  |  |  | 0.00 | 0.00 |  | 0.00 |
| Hermit crab unidentified | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 | 0.06 | 0.00 | 0.00 | 0.05 | 0.00 | 0.00 | 0.00 | 0.02 |
| Invertebrate unidentified | 0.05 | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 |
| Large Sculpins | 0.00 | 0.02 | 0.00 | 0.00 | 0.00 |  |  |  |  |  |  |  |  |
| Bigmouth Sculpin |  |  |  |  |  | 0.01 | 0.01 | 0.00 | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 |
| Great Sculpin |  |  |  |  |  | 0.00 | 0.01 | 0.01 | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 |
| Plain Sculpin |  |  |  |  |  | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| Warty Sculpin |  |  |  |  |  | 0.41 | 0.00 | 0.00 |  |  |  | 0.00 | 0.00 |
| Yellow Irish Lord |  |  |  |  |  | 0.01 | 0.00 | 0.01 | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 |
| Misc crabs | 0.08 | 0.17 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.02 | 0.01 | 0.00 | 0.00 | 0.00 | 0.00 |
| Misc fish | 0.03 | 0.00 | 0.00 | 0.00 | 0.00 | 0.01 | 0.01 | 0.05 | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 |
| Other osmerids | 0.01 | 0.00 | 0.00 | 0.08 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| Other Sculpins | 0.01 | 0.00 | 0.01 | 0.02 | 0.00 | 0.00 | 0.00 | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 | 0.01 |
| Pandalid shrimp | 0.38 | 0.01 | 0.08 | 0.10 | 0.00 | 0.02 | 0.01 | 0.17 | 0.02 | 0.07 | 0.02 | 0.00 | 0.32 |
| Polychaete unidentified | 0.00 |  | 0.03 |  | 0.00 | 0.00 | 0.00 | 0.00 | 0.78 |  | 0.00 |  | 0.17 |
| Scypho jellies | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| Sea anemone unidentified | 0.04 | 0.02 | 0.00 | 0.00 | 0.00 | 0.02 | 0.00 | 0.04 | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 |
| Sea pens whips | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.01 | 0.01 | 0.00 | 0.01 | 0.00 | 0.00 |
| Sea star | 0.01 | 0.01 | 0.03 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 | 0.00 |
| Snails | 0.14 | 0.02 | 0.07 | 0.02 | 0.00 | 0.05 | 0.01 | 0.03 | 0.07 | 0.03 | 0.02 | 0.00 | 0.01 |
| Sponge unidentified | 0.12 | 0.21 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| Stichaeidae | 0.51 | 0.02 | 0.75 | 0.55 | 0.00 | 0.08 | 0.01 | 0.20 | 0.01 | 0.00 | 0.03 | 0.00 | 0.20 |
| urchins dollars cucumbers | 0.05 | 0.01 | 0.04 | 0.00 | 0.00 | 0.00 | 0.00 | 0.03 | 0.03 | 0.00 | 0.01 | 0.00 | 0.00 |

Table 19. Proportion of prohibited species catch caught in the GOA flathead sole fishery in 2015

|  | 2015 |  | 2014 |  |
| :---: | :---: | :---: | :---: | :---: |
| Species Group Name | PSCNQ <br> Estimate | Halibut <br> Mortality | PSCNQ <br> Estimate | Halibut <br> Mortality |
| Bairdi Tanner Crab | 0.017 | -- | 0.000 | -- |
| Blue King Crab |  | -- |  | -- |
| Chinook Salmon Golden (Brown) King | 0.000 | -- | 0.075 | -- |
| Crab | 0.000 | -- | 0.000 | -- |
| Halibut | 0.001 | 0.002 | 0.001 | 0.001 |
| Herring | 0.000 | -- | 0.000 | -- |
| Non-Chinook Salmon Opilio Tanner (Snow) | 0.000 | -- | 0.000 | -- |
| Crab | 0.000 | -- | 0.000 | -- |
| Red King Crab | 0.000 | -- | 0.000 | -- |

Figures


Figure 1. Catch biomass in metric tons 1978-2015 (as of October 10, 2015).


Figure 2. GOA trawl survey catch per unit effort (CPUE) for flathead sole for the 2011-2015 surveys. Purple lines denote CPUE values and pink dots denote hauls were no flathead sole were caught.


Figure 3. Survey biomass index (black dots), asymptotic 95\% confidence intervals (vertical black lines), and estimated survey biomass for the proposed 2015 model, the 2015 model without 2014-2015 data, and the 2013 accepted model (solid lines).


Figure 4. Time series of spawning biomass and $95 \%$ asymptotic confidence intervals for the proposed 2015 model, the 2015 model without 2014-2015 data, and the 2013 accepted model.


Figure 5. Recruitment deviations for years 1978-2012 and 95\% asymptotic confidence intervals for the proposed 2015 model, the 2015 model without 2014-2015 data, and the 2013 accepted model.


Figure 6. Time series of age-0 recruits for the proposed 2015 model, the 2015 model without 2014-2015 data, and the 2013 accepted model.


Figure 7. Selectivity curves for the fishery (blue lines) and the survey (red lines), and for females (solid lines) and males (dashed lines) for the proposed 2015 model.


Figure 8. Estimated length-at-age relationship with 95\% asymptotic confidence intervals for males (blue) and females (red). The blue dashed line and red solid line show the mean relationship and dotted lines show confidence intervals.

## length comps, whole catch, aggregated across time by fleet



Figure 9. Observed (grey shaded area, black lines) and expected (red lines) proportions-at-length, aggregated over years for the fishery and survey and for females (upper half of plots) and males (lower half of plots) for the proposed 2015 model.


Figure 10. Observed (grey filled area and black line) and expected (lines) fishery length compositions for the proposed 2015 model (1 of 2).
length comps, whole catch, Fishery


Figure 11. As for Figure 10, but for years (2 of 2).


Figure 12. Observed (grey filled area and black line) and expected (lines) survey length compositions for the proposed 2015 model (1 of 2).

## Conditional AAL plot, whole catch, Survey



Figure 13. Observed and expected mean age-at-length for both females and males with $90 \%$ intervals about observed age-at-length (left panels) and observed and expected standard deviation in age-at-length (right panels) for the proposed 2015 model for years 1990-1996.

## Conditional AAL plot, whole catch, Survey



Figure 14. Observed and expected mean age-at-length for both females and males with $90 \%$ intervals about observed age-at-length (left panels) and observed and expected standard deviation in age-at-length (right panels) for the proposed 2015 model for years 1999-2003 (1 of 3).

Conditional AAL plot, whole catch, Survey


Figure 15. Observed and expected mean age-at-length for both females and males with $90 \%$ intervals about observed age-at-length (left panels) and observed and expected standard deviation in age-at-length (right panels) for the proposed 2015 model for years 2005-2009 (2 of 3).

## Conditional AAL plot, whole catch, Survey



Figure 16. Observed and expected mean age-at-length for both females and males with $90 \%$ intervals about observed age-at-length (left panels) and observed and expected standard deviation in age-at-length (right panels) for the proposed 2015 model for years 2011-2013 (3 of 3).


Figure 17. Pearson residuals associated with fits to the length-at-age relationship within the model for females (red, top panel) and males (blue, bottom panel) for the survey (1 of 3).


Figure 18. Pearson residuals associated with fits to the length-at-age relationship within the model for females (red, top panel) and males (blue, bottom panel) for the survey (2 of 3).


Figure 19. Pearson residuals associated with fits to the length-at-age relationship within the model for females (red, top panel) and males (blue, bottom panel) for the survey (3 of 3 ).

## Spawning biomass (mt) with $\sim 95 \%$ asymptotic intervals



Figure 20. Time series of estimated spawning stock biomass (mt) over time (solid blue line and circles) and asymptotic $95 \%$ confidence intervals (blue dashed lines) for the current base case model.


Figure 21. Spawning stock biomass relative to $B_{35 \%}$ and fishing mortality (F) relative to $F_{35 \%}$ from 19782017 (solid black line), the OFL control rule (dotted red line), the maxABC control rule (solid red line), $B_{35 \%}$ (vertical grey line), and $F_{35 \%}$ (horizontal grey line). The 2016 and 2017 spawning biomass and fishing mortality rates are as predicted by Alternatives 1 and 2 in the harvest projections.


Figure 22. Spawning stock biomass and corresponding 95\% asymptotic confidence intervals for base case model runs with 0 to 10 years of the most recent data removed.


Figure 23. Recruitment deviations (top panel) and age-0 recruits (bottom panel) for base case model runs with 0 to 10 years of the most recent data removed.


Figure 24. Gulf of Alaska food web from the GOA ecosystem model (Aydin et al., 2007) highlighting adult flathead sole links to predators (blue boxes and lines) and prey (green boxes and lines). Box size reflects relative standing stock biomass.


Figure 25. Gulf of Alaska food web from the GOA ecosystem model (Aydin et al., 2007) highlighting juvenile flathead sole links to predators (blue boxes and lines) and prey (green boxes and lines). Box size reflects relative standing stock biomass.


Figure 26. Diet composition for Gulf of Alaska adult flathead sole from the GOA ecosystem model (Aydin et al., 2007).


Figure 27. Diet composition for Gulf of Alaska juvenile flathead sole from the GOA ecosystem model (Aydin et al., 2007).


Figure 28. Decomposition of natural mortality for Gulf of Alaska adult flathead sole from the GOA ecosystem model (Aydin et al., 2007).

## GOA FH. Sole_Juv mortality



Figure 29. Decomposition of natural mortality for Gulf of Alaska juvenile flathead sole from the GOA ecosystem model (Aydin et al., 2007).

## Appendix 8A: Non-Commercial Catches of GOA Flathead Sole

Table A1. NMFS data sources

| Year | Annual <br> Longline <br> Survey | Salmon EFP 13-01 | Shelikof <br> Acoustic Survey | Shelikof and Chirikof EIT | Shumagin and Sanak EIT | Shumigans <br> Acoustic Survey | Structure of Gulf of Alaska Forage Fish Communities | Western Gulf of Alaska Pollock Acoustic Cooperative Survey |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1990 | 80.785 |  |  |  |  |  |  |  |
| 1991 | 53.619 |  |  |  |  |  |  |  |
| 1992 | 67.202 |  |  |  |  |  |  |  |
| 1993 | 56.48 |  |  |  |  |  |  |  |
| 1994 | 40.037 |  |  |  |  |  |  |  |
| 1995 | 82.214 |  |  |  |  |  |  |  |
| 1996 | 48.615 |  |  |  |  |  |  |  |
| 1997 | 46.469 |  |  |  |  |  |  |  |
| 1998 | 35.032 |  |  |  |  |  |  |  |
| 1999 | 33.602 |  |  |  |  |  |  |  |
| 2000 | 12.155 |  |  |  |  |  |  |  |
| 2001 | 17.159 |  |  |  |  |  |  |  |
| 2002 | 24.309 |  |  |  |  |  |  |  |
| 2003 | 15.73 |  |  |  |  |  |  |  |
| 2004 | 20.019 |  |  |  |  |  |  |  |
| 2005 | 7.15 |  |  |  |  |  |  |  |
| 2006 | 40.036 |  |  |  |  |  |  |  |
| 2007 | 29.313 |  |  |  |  |  |  |  |
| 2008 | 37.891 |  |  |  |  |  |  |  |
| 2009 | 54.334 |  |  |  |  |  |  |  |
| 2010 | 81.5 |  | 4.492 |  |  | 201.01 | 7.808 | 15.6 |
| 2011 | 38.606 |  |  |  |  |  |  |  |
| 2012 | 18.55 |  |  | 7.22 | 2.76 |  |  |  |
| 2013 | 56.478 | 380 |  |  |  |  |  |  |
| 2014 | 62.913 | 180 |  |  |  |  |  |  |

Table A2. ADF\&G data sources

| Year | Large-Mesh <br> Trawl Survey | Sablefish <br> Longline <br> Survey | Scallop <br> Dredge <br> Survey | Small-Mesh <br> Trawl Survey |
| :---: | :---: | :---: | :---: | :---: |
| 1998 | 2465.29 | 3.8 | 0.22 |  |
| 1999 | 4842.57 | 5.6 | 0.45 |  |
| 2000 | 2723.03 | 1 |  | 2427.75 |
| 2001 | 6394.27 | 2.6 |  |  |
| 2002 | 2277.08 | 1.4 | 0.09 |  |
| 2003 | 5496.63 | 2.4 |  | 2565.67 |
| 2004 | 3864.43 | 1.1 |  | 3299.13 |
| 2005 | 6450.74 |  | 7.47 | 3157.94 |
| 2006 | 2617.47 | 7.864 | 7.47 | 2797.83 |
| 2007 | 3856.18 |  | 1.05 | 385.44 |
| 2008 | 2099.94 |  | 0.3 |  |
| 2009 | 5154.93 |  | 10.41 |  |
| 2010 | 84389.475 |  | 1.49 | 12008.01 |
| 2011 | 84023.542 |  | 52.078 | 9154.2 |
| 2012 | 92629.38 |  | 5.95 | 7976.89 |
| 2013 | 78993.8 |  | 14.4 | 4789.321 |
| 2014 | 72746.41 |  |  | 6175.3 |

Table A3. IPHC data

| Year | IPHC Annual <br> Longline <br> Survey |
| :---: | :---: |
| 2010 | 4 |
| 2011 | 1 |
| 2012 | 29 |
| 2014 | 20 |

