12. Assessment of the Pacific ocean perch stock in the Bering Sea/Aleutian Islands

by

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Executive Summary

The last full assessment for Pacific ocean perch (POP) was presented to the Plan Team in 2012. The following changes were made to POP assessment relative to the November 2012 SAFE:

Summary of Changes in Assessment Inputs

Changes in the Input Data

- 1) The harvest time series were updated through October 11, 2014.
- 2) The survey biomass estimates and age composition data from the U.S.-Japan cooperative survey in 1980,1983, and 1986 were removed from the assessment.
- 3) The 2014 AI survey biomass estimate and length composition was included in the assessment.
- 4) The 2012 AI survey age compositon was included in the assessment.
- 5) The 2013 fishery age composition were included in the assessment.
- 6) The 2012 fishery length composition was included in the assessment

Changes in the Assessment Methodology

- 1) Several fishery selectivity models were evaluated, with the recommended model using a bicubic spline to model to estimate fishery selectivity as a function of year and age.
- 2) The multinomial input sample sizes for the age and length composition data were obtained by an iterative reweighting procedure that ensures that the standard deviation of the normalized residuals for each composition data type is 1.
- 3) The length-at-age, weights-at-age, and age-to-length conversion matrix were updated based on data from the NMFS AI trawl survey beginning in 1991.

Summary of Results

A summary of the 2014 assessment recommended ABC's relative to the 2013 recommendations is shown below. BSAI Pacific ocean perch are not overfished or approaching an overfished condition. The recommended 2015 ABC and OFL are 34,988 t and 42,588 t, which are 11% and 13% increases, respectively, from the maximum ABC and OFL specified last year for 2015 of 31,641 t and 37,817 t. The 2014 AI survey biomass is large and consistent with the survey biomass estimates in 2010 and 2012, and the size composition data continue to show relatively strong cohorts from 1994 to 2000. A summary of the recommended ABCs and OFLs from this assessment relative the ABC and OFL specified last year is shown below:

| | As estima | nted or | As estimated or | | |
|--------------------------------------|----------------|----------------|-----------------|----------------|--|
| | specified last | t year for: | recommended t | his year for: | |
| Quantity | 2014 | 2015 | 2015 | 2016 | |
| M (natural mortality rate) | 0.062 | 0.062 | 0.062 | 0.062 | |
| Tier | 3a | 3a | 3a | 3a | |
| Projected total (age 3+) biomass (t) | 639,505 | 620,270 | 577,967 | 561,090 | |
| Female spawning biomass (t) | | | | | |
| Projected | 257,878 | 243,400 | 234,426 | 223,744 | |
| $B_{100\%}$ | 459,436 | 459,436 | 423,008 | 423,008 | |
| $B_{40\%}$ | 183,774 | 183,774 | 169,203 | 169,203 | |
| $B_{35\%}$ | 160,803 | 160,803 | 148,053 | 148,053 | |
| F_{OFL} | 0.076 | 0.076 | 0.109 | 0.109 | |
| $maxF_{ABC}$ | 0.063 | 0.063 | 0.089 | 0.089 | |
| F_{ABC} | 0.063 | 0.063 | 0.089 | 0.089 | |
| OFL (t) | 39,585 | 37,817 | 42,558 | 40,809 | |
| maxABC (t) | 33,122 | 31,641 | 34,988 | 33,550 | |
| ABC (t) | 33,122 | 31,641 | 34,988 | 33,550 | |
| | As determined | last year for: | As determined | this year for: | |
| Status | 2012 | 2013 | 2013 | 2014 | |
| Overfishing | No | n/a | No | n/a | |
| Overfished | n/a | | n/a | No | |
| Approaching overfished | n/a | | n/a | No | |

Summaries for the Plan Team

The ABC for BSAI Pacific ocean perch is currently apportioned among four areas: the western, central, and eastern Aleutian Islands, and eastern Bering Sea. The current method of determining area apportionments uses a weighted average of the three most recent trawl survey biomass estimates in each of these areas is used to apportion the ABC. Weights of 4, 6, and 9 are used, with higher weights being applied to the more recent surveys. It is also of interest to estimate the area proportions using the random effects model. The survey averaging workgroup is evaluating the use of the random effects model to smooth survey time series for computing area apportionments, and its use for computing area proportion might logically be delayed until after the workgroup has completed their evaluations. However, a comparison between the two methods may be useful. The following table gives the projected OFLs and apportioned ABCs for 2015 and 2016 from the two methods, and the recent OFLs, ABCs, TACs, and catches.

| | BSAI | Western AI | Central AI | Eastern AI | EBS | Total |
|---------------------------|--------|------------|------------|------------|-------|--------|
| Apportionment | | | | | | |
| (weighted average) | | 28.5% | 23.6% | 24.6% | 23.3% | 100% |
| Apportionment | | | | | | |
| (RE model) | | 29.1% | 22.1% | 23.8% | 25.1% | 100% |
| OFL (2013) | 41,900 | | | | | |
| ABC (2013) | | 10,200 | 6,980 | 9,790 | 8,130 | 35,100 |
| TAC (2013) | | 10,200 | 6,980 | 9,790 | 8,130 | 35,100 |
| Catch (2013) | | 10,065 | 6,747 | 9,530 | 5,050 | 31,393 |
| | | | | | | |
| OFL (2014) | 39,585 | | | | | |
| ABC (2014) | | 9,598 | 6,594 | 9,246 | 7,684 | 33,122 |
| TAC (2014) | | 9,598 | 6,594 | 9,246 | 7,684 | 33,122 |
| Catch (2014) ¹ | | 9,485 | 6,438 | 8,124 | 1,842 | 25,889 |
| | | | | | | |
| OFL (2015) | 42,558 | | | | | |
| ABC (2015, | | | | | | |
| weighted average) | | 9,981 | 8,240 | 8,623 | 8,143 | 34,988 |
| ABC (2015, RE | | | | | | |
| model) | | 10,182 | 7,723 | 8,312 | 8,771 | 34,988 |
| | | | | | | |
| OFL (2016) | 40,809 | | | | | |
| ABC (2016, | | | | | | |
| weighted average) | | 9,571 | 7,902 | 8,269 | 7,809 | 33,550 |
| ABC (2016, RE | | | | | | |
| model) | | 9,763 | 7,406 | 7,970 | 8,411 | 33,550 |

Catch through October 11, 2014

Responses to SSC and Plan Team Comments in General

The SSC requests that all assessment authors of AI species evaluate AI survey information to ensure that the same standardized survey time series is used. (SSC, December 2012)

Model runs in this assessment exclude the cooperative surveys conducted in the 1980s.

"The Teams recommended that each stock assessment model incorporate the best possible estimate of the current year's removals. The Teams plan to inventory how their respective authors address and calculate total current year removals. Following analysis of this inventory, the Teams will provide advice to authors on the appropriate methodology for calculating current year removals to ensure consistency across assessments and FMPs." (Plan Team, September 2013)

The estimates of current year catch are inferred by expanding the catch through September, 2014, by the recent pattern of the proportion of the remaining ABC that is caught by the end of the year.

"For assessments involving age-structured models, this year's CIE review of BSAI and GOA rockfish assessments included three main recommendations for future research: Authors should consider: (1) development of alternative survey estimators, (2) evaluating selectivity and fits to the plus group, and (3) re-evaluating natural mortality rates. The SSC recommends that authors address the CIE review during full assessment updates scheduled in 2014." (SSC, December 2013)

Selectivity curves and natural mortality rates are evaluated in this assessment. The development of alternative survey estimators (i.e., model-based standardization of survey catch data) affects all NPFMC assessments that use survey data. Potential methodologies have been discussed in a limited number of meetings in 2014 among AFSC scientists, and between AFSC scientists and NWFSC scientists, who are in the process of developing more refined standardization methods. Continuation of these meetings will hopefully result in progress on this task.

"During public testimony, it was proposed that assessment authors should consider projecting the reference points for the future two years (e.g., 2014 and 2015) on the phase diagrams. It was suggested that this forecast would be useful to the public. The SSC agrees. The SSC appreciated this suggestion and asks the assessment authors to do so in the next assessment." (SSC December 2013)

These projections were added to the phase-plane plots.

Responses to SSC and Plan Team Comments Specific to this Assessment

The SSC offers the following advice to assessment authors:

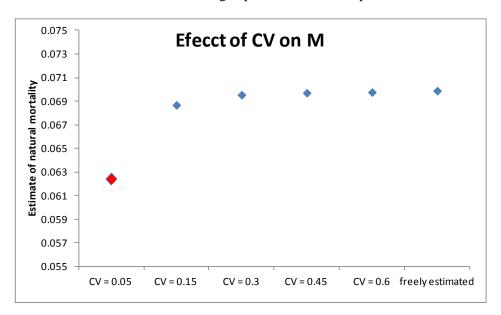
- Explore alternative selectivity patterns
- Evaluate alternative selectivity time periods
- Provide model sensitivity to Q and M
- Explore lack of fit to the plus age group
- Fit to the maturity data should be evaluated for potential bias from excess data consisting of 100% and 0% maturity because the logistic model cannot predict 0 and 1. (SSC, December 2012)

Several methods for modeling fishery selectivity over age and year were considered in this assessment, with a bicubic spline (used in the preferred model) providing improved fits to the survey and fishery age composition plus group.

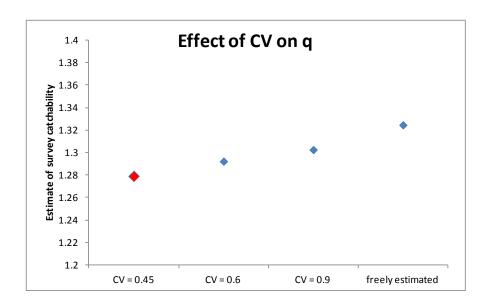
Although predicted proportion mature obtained from the logistic model asymptotically approaches 1 as

age increases, the difference between the asymptote of 1 and the predicted value for older ages is sufficiently small such that any effect on the model fit is not a concern. For example, the predicted proportion mature at age 25 is 0.99999905.

A series of models runs were conducted to address how sensitive the model estimates of natural mortality (M) and survey catchability (q) are to the prior distributions. Because M and q are inversely correlated, the runs for M were conducted by holding q fixed at the value estimated in the preferred model (and vice versa). The current model uses a prior distribution for M with a mean and CV each set to 0.05, and produces and estimate of 0.062 (shown in red in the graph below). Increasing the CV, or allowing M to be freely estimated, produces estimates ranging between 0.069 and 0.070. This value of M is consistent with a recent meta-analysis relating natural mortality estimates to maximum age (Then et al. 2014). Future assessments will consider increasing expected value of the prior distribution for M from 0.05 to 0.07.



The prior for q has a mean of 1 and a CV of 0.45, and a similar plot showing the effect of increasing the CV of the prior is shown below. The estimates of q ranges from 1.28 estimated in the current model to 1.32 when q is freely estimated (while holding M fixed at 0.0062). For this stock, estimates of q are more sensitive to the addition of age and length composition data and survey biomass estimates, which motivate obtaining information on q outside the assessment model in order to develop more informative priors.



Introduction

Pacific ocean perch (*Sebastes alutus*) inhabit the outer continental shelf and upper slope regions of the North Pacific Ocean and Bering Sea. Pacific ocean perch, and four other associated species of rockfish (northern rockfish, *S. polyspinis*; rougheye rockfish, *S. aleutianus*; shortraker rockfish, *S. borealis*; and sharpchin rockfish, *S. zacentrus*) were managed as a complex in the two distinct areas from 1979 to 1990. Known as the POP complex, these five species were managed as a single entity with a single TAC (total allowable catch). In 1991, the North Pacific Fishery Management Council separated POP from the other red rockfish in order to provide protection from possible overfishing. Of the five species in the former POP complex, *S. alutus* has historically been the most abundant rockfish in this region and has contributed most to the commercial rockfish catch.

Information on Stock Structure

A variety of types of research can be used to infer stock structure of POP, including age and length compositions, growth patterns and other life-history information, and genetic studies. Spatial differences in age or length compositions can be used to infer differences in recruitment patterns that may correspond to population structure. In Queen Charlotte Sound, British Columbia, Gunderson (1972) found substantial differences in the mean lengths of POP in fishery hauls taken at similar depths which were related to differences in growth rates and concluded that POP likely form aggregations with distinct biological characteristics. In a subsequent study, Gunderson (1977) found differences in size and age composition between Moresby Gully and two other gullies in Queen Charlotte Sound. Westrheim (1970, 1973) recognized "British Columbia" and "Gulf of Alaska" POP stocks off the western coast of Canada based upon spatial differences in length frequencies, age frequencies, and growth patterns observed from a trawl survey. In a study that has influenced management off Alaska, Chikuni (1975) recognized distinct POP stocks in four areas – eastern Pacific (British Columbia), Gulf of Alaska, Aleutian Islands, and Bering Sea. However, Chikuni (1975) states that the eastern Bering Sea (EBS) stock likely receives larvae from both the Gulf of Alaska (GOA) and Aleutian Islands (AI) stock, and the AI stock likely receives larvae from the GOA stock.

An alternative approach to evaluating stock structure involves examination of rockfish life-history stages directly. Stock differentiation occurs from separation at key life-history stages. Because many rockfish species are not thought to exhibit large-scale movements as adults, movement to new areas and boundaries of discrete stocks may depend largely upon the pelagic larval and juvenile life-history stages. Simulation modeling of ocean currents in the Alaska region suggest that larval dispersal may occur over very broad areas, and may be dependent on month of parturition (Stockhausen and Herman 2007).

Analysis of field samples of rockfish larvae are hindered by difficulties in indentifying species. Analyses of archived *Sebastes* larvae was undertaken by Dr. Art Kendall revealed that species identification based on morphological characteristics is difficult because of overlapping characteristics among species, as few rockfish species in the north Pacific have published descriptions of the complete larval developmental series. However, all of the larvae examined could be assigned to four morphs identified by Kendall (1991), where each morph is associated with one or more species. Rockfish identification can be aided by studies that combine genetic and morphometric techniques and information has been developed to identify individual species based on allozymes (Seeb and Kendall 1991) and mitochondrial DNA (Gharrett et al. 2001, Rocha-Olivares 1998). The Ocean Carrying Capacity (OCC) field program, conducted by the Auke Bay laboratory, uses surface trawls to collect juvenile salmon and incidentally collects juvenile rockfish. These juvenile rockfish are large enough (approximately 25 mm and larger) to allow extraction of a tissue sample for genetic analysis without impeding morphometric studies. In 2002, species identifications were made for an initial sample of 55 juveniles with both morphometric and genetic techniques. The two techniques showed initial agreement on 39 of the 55 specimens, and the

genetic results motivated re-evaluation of some of the morphological species identifications. Forty of the specimens were identified as POP, and showed considerably more morphological variation for this species than previously documented.

Because stocks are, by definition, reproductively isolated population units, it is expected that different stocks would show differences in genetic material due to random drift or natural selection. Thus, analysis of genetic material from North Pacific rockfish is currently an active area of research.

Seeb and Gunderson (1988) used protein electrophoresis to infer genetic differences based upon differences in allozymes from POP collected from Washington to the Aleutian Islands. Discrete genetic stock groups were not observed, but instead gradual genetic variation occurred that was consistent with the isolation by distance model. The study included several samples in Queen Charlotte Sound where Gunderson (1972, 1977) found differences in size compositions and growth characteristics. Seeb and Gunderson (1988) concluded that the gene flow with Queen Charlotte Sound is sufficient to prevent genetic differentiation, but adult migrations were insufficient to prevent localized differences in length and age compositions. More recent studies of POP using microsatellite DNA revealed population structure at small spatial scales, consistent with the work of Gunderson (1972, 1977). These findings suggest that adult POP do not migrate far from their natal grounds and larvae are entrained by currents in localized retention areas (Withler et al. 2001).

Interpretations of stock structure are influenced by the technique used to assess genetic analysis differentiation, as illustrated by the differing conclusions produced from the POP allozyme work of Seeb and Gunderson (1988) and the microsatellite work of Withler et al. (2001). Note that these two techniques assess components of the genome that diverge on very different time scales and that, in this case, microsatellites are much more sensitive to genetic isolation. Protein electrophoresis examines DNA variation only indirectly via allozyme frequencies, and does not recognize situations where differences in DNA may result in identical allozymes (Park and Moran 1994). In addition, many microsatellite loci may be selectively neutral or near-neutral, whereas allozymes are central metabolic pathway enzymes and do not have quite the latitude to produce viable mutations. The mutation rate of microsatellite alleles can be orders of magnitude higher than allozyme locus mutation rates. Most current studies on rockfish genetic population structure involve direct examination of either mitochondrial DNA (mtDNA) or microsatellite DNA.

Dr. Anthony Gharrett of the Juneau Center of Fisheries and Ocean Sciences has examined the mtDNA and microsatellite variation for POP samples collected in the GOA and BSAI. The POP mtDNA analysis was performed on 124 fish collected from six regions ranging from southeast Alaska to the Bering Sea slope and central Aleutian Islands. No population structure was observed, as most fish (102) were characterized by a common haplotype. Preliminary results from an analysis of 10 microsatellite loci from the six regions resulted in 7 loci with significant heterogeneity in the distribution of allele frequencies. Additionally, the sample in each region was statistically distinct from those in adjacent regions, suggesting population structure on a relatively fine spatial scale consistent with the results on Gunderson (1972, 1977) and Wither et al. (2001). Ongoing genetic research with POP is focusing on increasing the sample sizes and collection sites for the microsatellite analysis in order to further refine our perception of stock structure.

Fishery

POP were highly sought by Japanese and Soviet fisheries and supported a major trawl fishery throughout the 1960s. Catches in the eastern Bering Sea peaked at 47,000 (metric tons, t) in 1961; the peak catch in the Aleutian Islands region occurred in 1965 at 109,100 t. These stocks were not productive enough to support such large removals. Catches continued to decline throughout the 1960s and 1970s, reaching

their lowest levels in the mid 1980s. With the gradual phase-out of the foreign fishery in the 200-mile U.S. Exclusive Economic Zone (EEZ), a small joint-venture fishery developed but was soon replaced by a domestic fishery by 1990. In 1990 the domestic fishery recorded the highest POP removals since 1977. The OFLs, ABCs, TACs, and catches by management complex from 1977 to 2000 (when POP were managed as separate stocks in the EBS and AI) are shown in Table 1. Note that in some years, POP were managed in the "POP complex" management group, which also included rougheye rockfish, shortraker rockfish, northern rockfish, and sharpchin rockfish. In 2002 POP were managed as a single stock across the BSAI (with the ABC subdivided between the EBS and AI subareas, and the BSAI OFLs, ABCs, TACs, and catches for this period is shown in Table 2. The ABCs, TACs, and catches from 1988 to 2012 are shown in Table 2. The catches of POP from 1977 by fishery type (i.e., foreign, joint venture, or domestic) is shown in Table 3.

Estimates of retained and discarded POP from the fishery have been available since 1990 (Table 4). From 1990-2009, the eastern Bering Sea region generally showed a higher discard rate than in the Aleutian Islands region, with the average rates 33% and 14%, respectively. From 2010-2013, the eastern Bering Sea discard rate was less than 7% but increased to 41% in 2014, although this may be an artifact of only including the catches through Oct 11, 2014. In contrast, the Aleutian Islands discard rates from 2010-2014 were less than 3%.

Initial age-structured assessments for BSAI POP modeled separate selectivity curves for the foreign and domestic fisheries (Ianelli and Ito 1992), although examination of the distribution of observer catch reveals interannual changes in the depth and areas in which POP are observed to be caught within the foreign and domestic periods. For example, POP are predominately taken in depths between 200 m and 300 m, although during the late 1970s-early-1980s a relatively large portion of POP were observed to be captured at depths greater than 300 m (Table 5, Figure 1). Additionally, from 1999 through the early 2000s the proportion caught between 100 m and 200 m increased from ~ 20% in the early to mid 1990s to ~ 30%, and since the mid 2000s the proportion caught between 200 m and 300 m has increased. The area of capture has changed as well; during the late 1970s POP were predominately captured in the western Aleutians, whereas from the early 1980s to the mid-1990s POP were captured predominately in the eastern Aleutians. Establishment of area-specific TACs in the mid-1990s redistributed the POP catch such that about 50% of the current catch is now taken in the western Aleutians (Table 6, Figure 1). Note that the extent to which the patterns of observed catch can be used as a proxy for patterns in total catch is dependent upon the degree to which the observer sampling represents the true fishery. In particular, the proportions of total POP caught that were actually sampled by observers were very low in the foreign fishery, due to low sampling ratio prior to 1984 (Megrey and Wespestad 1990).

Non-commercial catches are shown in Appendix A.

Data

Fishery Data

Catch per unit effort (CPUE) data from Japanese trawl fisheries indicate that POP stock abundance has declined to very low levels in the Aleutian Islands region (Ito 1986). By 1977, CPUE values had dropped by more than 90-95% from those of the early 1960s. Japanese CPUE data after 1977, however, is not a good index of stock abundance because most of the fishing effort has been directed to species other than POP. Standardizing and partitioning total groundfish effort into effort directed solely toward POP is difficult. Increased quota restrictions, effort shifts to different target species, and rapid improvements in fishing technology undoubtedly affect our estimates of effective fishing effort. Consequently, we included CPUE data primarily to evaluate its consistency with other sources of information. We used nominal CPUE data for class 8 trawlers in the eastern Bering Sea and Aleutian Islands regions from 1968-

1979. During this time period these vessels were known to target on POP (Ito 1982).

Length measurements and otoliths read from the EBS and AI management areas (Tables 7 and 8) were combined to create fishery age/size composition matrices. Years that were not selected for age or length composition were rejected due to low samples sizes of fish measured (years 1973-1976, 1985-1986), and/or otoliths read (years 1984-86). In 1982, the method for aging otoliths at the Alaska Fisheries Science Center changed from surface reading to the break and burn method (Betty Goetz, Alaska Fisheries Science Center, pers. comm.), as the latter method is considered more accurate for older fish (Tagart 1984). The time at which the otoliths collected from 1977 to 1982 were read is not known for many vessels and cruises. However, the information available suggests that otoliths from 1977 to 1980 were read prior to 1981, whereas otoliths from 1981 and 1982 were read after 1982. Thus, fishery otoliths from 1977 to 1980 were not used because they were believed to be read by surface ageing and thought to be biased.

Beginning in 1998, samples of otoliths from the fishery catch have been read almost annually or biennially, and show relatively strong year classes from 1984-1988. Fishery age compositions from 2005-2013 indicate several strong recent year classes from 1995-2000 (Figure 2).

Survey Data

Cooperative U.S. – Japan trawl surveys were conducted in the AI 1980, 1983, and 1986, and have been used in previous BSAI POP assessments. However, differences exist in gear design and vessels used between these surveys and the NMFS surveys beginning in 1991 (Skip Zenger, National Marine Fisheries Service, personal communication). For example, the Japanese nets used in the cooperative surveys varied between years and included large roller gear, in contrast to the poly-nor'eastern nets used in the current surveys (Ronholt et al 1994). Given the difficulty of documenting the methodologies for these surveys, and standardizing these surveys with the NMFS surveys, this assessment model is conducted with only the NMFS surveys.

The Aleutian Islands survey biomass estimates were used as an index of abundance for the BSAI POP stock. Since 2000 the survey has occurred biennially, although the 2008 survey was canceled due to a lack of funding. Note that there is wide variability among survey estimates from the portion of the southern Bering Sea portion of the survey (from 165 ° W to 170 ° W), as the post-1991 coefficients of variation (CVs) range from 0.41 to 0.63 (Table 9). The biomass estimates in this region increased from 1,501 t in 1991 to 18,217 t in 1994, and have since ranged between 12,099 t (1997) and 87,794 t (2010). The estimated biomass of Pacific ocean perch in the Aleutian Islands management area region (170° W to 170° E) appears to be less variable, with CVs ranging from 0.12 to 0.24. The biomass estimates for the AI area have ranged between a low of 342,785 t in 1991 and 887,559 t in 2014, with the estimates from the 2010, 2012, and 2014 survey ranging between 864,000 t (2012) and 888,000 t (2014).

From 1991-2010, the Aleutian Island surveys have indicated higher abundances in the western Aleutian Islands than in other subareas. However, in the 2012 survey the biomass estimate for the western AI was 263,661 t, a decrease from the 2010 estimate of 395,933 t, whereas the estimates for the eastern AI increased from 266,607 in the 2010 survey to 366,413 in the 2012 survey (Table 10). The estimated biomass from the 2014 survey for the western AI and eastern AI are 338,000 t and 234,000 t, respectively, which are more similar to the 2010 survey than the 2012 survey. The 2014 estimate for the central AI is 316,000 t, a 35% increase from the 2012 estimate. The total AI 2014 survey biomass estimate was 970,968 t with a CV of 0.19, which is a 8% increase from the 2012 estimate of 902,398 t (CV=0.17). Maps of survey CPUE are shown in Figure 3, and indicate relatively high abundance throughout much of the Aleutian Islands.

Age composition data exists for each Aleutian Islands survey, and the length measurements and otoliths

read are shown in Table 11. The survey age compositions from 1991-2000 indicate relatively strong year classes in 1977, 1984, and 1988. Recent age composition data from 2004 -2012 indicate relatively strong year classes from 1996 to 2000 (Figure 4).

The biennial EBS slope survey was initiated in 2002. The most recent slope survey prior to 2002, excluding some preliminary tows in 2000 intended for evaluating survey gear, was in 1991. Previous slope survey results have not been used in the BSAI model due to high CVs, relatively small population sizes compared to the AI biomass estimates, and lack of recent surveys. However, the biomass estimates in the EBS slope survey have been increasing, ranging from 76,665 t in 2002 to 231,383 in the 2012 survey, with CVs ranging from 0.33 in 2012 to 0.53 in 2002 (Table 9). The slope survey was canceled in 2008 and 2014 due to lack of funding. The slope survey results are not used in this assessment, but the feasibility of incorporating this time series will be evaluated in future years given the increased length of the time series and increased levels of biomass.

Comparison of Fishery and Survey Catches by Depth and Age

A comparison of fishery and survey catches can indicate whether fishery selectivity is suspected of being time-varying and/or dome-shaped. For example, some of the variation in depths and areas fished mentioned above could be related to changes in the stock distribution, which would suggest that availability to the stock is more constant than implied by examining only the catch data. This issue was examined by comparing the catch-weighted mean depth in the fishery (from 1991 – 2013) to that in tows from the Aleutian Islands trawl survey from 1991 to 2014. The survey tows show a relatively constant mean depth of capture across years, whereas the fishery depth of capture show higher of interannual variation (Figure 5).

The plus group for the POP assessment model is 40 years, and of interest is the relative age composition of the old fish within the plus group. Fishery and survey data were binned across years in each of five periods from 1990 to 2011, the age composition of ages 40 to 70+ are shown in Figure 6. Overall, survey age composition generally exceeds the fishery age composition, and this pattern seems to be more pronounced in the earlier time periods. For example, in the 1990-1991 period, the survey age composition exceeded the fishery age composition for each age from 41 – 52 (except age 53); a similar pattern is seen in the 1997-1998 period for ages greater than 58. The pattern can be seen more clearly in the histogram of differences between survey and fishery age proportions (Figure 7); positive differences indicate that the survey proportion exceeded the fishery proportion for a given age. Of the ages with a non-zero difference, the proportion of ages with a positive difference ranged between 0.78 and 0.82 for the four earliest periods, and decreased to 0.6 in the 2009-2011 period. Overall, these data suggest that the some domeshaped fishery selectivity has likely occurred since 1991, and that it may be diminishing in recent years.

Biological data

A large number of samples are collected from the surveys for age determination, length-weight relationships, sex ratio information, and for estimating the length distribution of the population. The age compositions were determined by constructing age-length keys for each year and using them to convert the observed length frequencies from each year. Because the survey age data were based on the break and burn method of ageing POP, they were treated as unbiased but measured with error. Kimura and Lyons (1991) reported that the percent agreement between readers varies from 60% for age 3 fish to 13% for age 25 fish data. The information on percent agreement was used to derive the variability of observed age around the "true" age, assuming a normal distribution. The mean number of fish at age available to the survey or fishery is multiplied by the aging error matrix to produce the observed survey or fishery age compositions.

Aging methods have improved since the start of the time series. Historically, POP age determinations were done using scales and surface readings from otoliths. These gave estimates of natural mortality of about 0.15 and longevity of about 30 years (Gunderson 1977). Based on the now accepted break and burn method of age determination using otoliths, Chilton and Beamish (1982) determined the maximum age of *S. alutus* to be 90 years. Using similar information, Archibald et al. (1981) concluded that natural mortality for POP should be on the order of 0.05.

Aleutian Islands survey data from 1991 through 2012 were used to estimate growth curves. The resulting von Bertalannfy growth parameters were $L_{inf} = 41.53$ cm, k = 0.14, and $t_0 = -1.202$. Growth information from the Aleutian Islands was used to convert estimated numbers-at-age within the model to estimated numbers-at-length.

A conversion matrix was created to convert modeled number at ages to modeled number at length bin, and consists of the proportion of each age that is expected in each length bin. This matrix was created by fitting a power relationship to the observed standard deviation in length at each age (obtained from the aged fish from the 1991-2012 surveys), and the predicted relationship was used to produce variation around the predicted size at age from the von Bertalanffy relationship. The resulting CVs of length at age of the transition matrix decrease from 0.13 at age 3 to 0.07 at age 40.

The estimated length(cm)-weight(g) relationship for Aleutian Islands POP was estimated with survey information from the same years, with the length-weight parameters estimated as $a = 9.56 \times 10^{-6}$ and b = 3.11, where weight $= a*(\text{length})^b$. The Aleutian Islands length-weight relationship was used to produce estimated weights at age.

| The following table summ | narizes the | data available | for the | BSALPOP | model: |
|-----------------------------|-----------------|----------------|---------|----------------|--------|
| The following table sulfill | idi izes tiie i | data available | ioi uic | DULLIOI | mouci. |

| Component | BSAI |
|---------------------------|--|
| Fishery catch | 1960-2014 |
| Fishery age composition | 1981-82, 1990, 1998, 2000-2009, 2011, 2013 |
| Fishery size composition | 1964-72, 1983-1984, 1987-1989, 1991-1997, 1999, 2010, 2012 |
| Fishery CPUE | 1968-79 |
| Survey age composition | 1991, 1994, 1997, 2000, 2002, 2004, 2006, 2010, 2012 |
| Survey length composition | 2014 |
| Survey biomass estimates | 1991, 1994, 1997, 2000, 2002, 2004, 2006, 2010, 2012, 2014 |

Analytic Approach

Model Structure

An age-structured population dynamics model, implemented in the software program AD Model Builder, was used to obtain estimates of recruitment, numbers at age, and catch at age. Population size in numbers at age *a* in year *t* was modeled as

$$N_{t,a} = N_{t-1,a-1}e^{-Z_{t-1,a-1}}$$
 $3 < a < A, 1960 < t \le T$

where Z is the sum of the instantaneous fishing mortality rate ($F_{t,a}$) and the natural mortality rate (M), A is the maximum number of age groups modeled in the population, and T is the terminal year of the analysis (defined as 2014).

The numbers at age A are a "pooled" group consisting of fish of age A and older, and are estimated as

$$N_{t,A} = N_{t-1,A-1}e^{-Z_{t-1,A-1}} + N_{t-1,A}e^{-Z_{t-1,A}}$$

The plus group was set to 40+, following a sensitivity analysis conducted in the 2012 stock assessment (Spencer and Ianelli 2012)

The numbers at age in the first year of the model are estimated as

$$N_a = R_0 e^{-M(a-3)}$$

where R_0 is the number of age 3 recruits for an unfished population, thus producing an age structure in equilibrium with an unfished stock. Previous assessments have estimated non-equilibrium numbers at age in the first year of the model (as a function of cohort-dependent deviations from average recruitment), although this formulation tended to put most of abundance in the first year in a single cohort. It is generally thought that little fishing for rockfish occurred prior to 1960, so an equilibrium unfished agestructure seems reasonable.

The total numbers of age 3 fish from 1960 to 2011 are estimated as parameters in the model, and are modeled with a lognormal distribution

$$N_{t,3} = e^{\mu_R + \nu_t}$$

where v_t is a time-variant deviation with a log-scale recruitment standard deviation of σ_r . Little information exists to determine the year-class strength for the three most recent cohorts (2012-2014), which were set to the estimated mean recruitment (based upon the log-scale mean, and the value of σ_r).

The fishing mortality rate for a specific age and time $(F_{t,a})$ is modeled as the product of a $s_{a,t}^f$ and a year-specific fully-selected fishing mortality rate f. The fully selected mortality rate is modeled as the product of a mean (μ_f) and a year-specific deviation (ε_t) , thus $F_{t,a}$ is

$$F_{t,a} = S_{a,t}^f f_t = S_{a,t}^f e^{(\mu_f + \varepsilon_t)}$$

The mean number-at-age for each year was computed as

$$\overline{N}_{t,a} = N_{t,a} (1 - e^{-Z_{t,a}}) / Z_{t,a}$$

Catch biomass-at-age was computed as the product of mean numbers at age, instantaneous fishing mortality, and weight at age. The predicted trawl survey biomass \hat{B}_t^{nvl} was computed as

$$\hat{B}_{t}^{twl} = q^{twl} \sum_{a} (\overline{N}_{t,a} s_{a}^{twl} W_{a})$$

where W_a is the population weight-at-age, s_a^{twl} is the survey selectivity, and q^{twl} is the trawl survey catchability. A CPUE index from 1968 to 1979 is also included in the assessment and is computed as

$$\hat{I}_{t}^{cpue} = q^{cpue} \sum_{a} (\overline{N}_{t,a} s_{a,t}^{f} W_{a})$$

where q^{cpue} is the scaling factor for the CPUE index.

To facilitate parameter estimation, prior distributions were used for the survey catchability and the natural mortality rate M. A lognormal distribution was also used for the natural mortality rate M, with the mean set to 0.05 and the CV set to 0.05. The standard deviation of log recruits, σ_r , was fixed at 0.75, a value consistent with the root mean squared error (RMSE; defined below) of recruitment deviations. Similar, the prior distribution for *qsurv* followed a lognormal distribution with a mean of 1.0 and a coefficient of variation (CV) of 0.05.

Several quantities were computed in order to compare the variance of the residuals to the assumed input variances. The RMSE should be comparable to the assumed coefficient of variation of a data series. This quantity was computed for the AI trawl survey and the estimated recruitments, and for lognormal distribution is defined as

$$RMSE = \sqrt{\frac{\sum_{n} (\ln(y) - \ln(\hat{y}))^{2}}{n}}$$

where y and \hat{y} are the observed and estimated values, respectively, of a series length n. The standardized deviation of normalized residuals (SDNR) are closely related to the RMSE. Values of SDNR approximately 1 indicate that the model is fitting a data component as well as would be expected for a given specified input variance. The normalized residuals for a given year i of the AI trawl survey data was computed as

$$\delta_i = \frac{\ln(B_i) - \ln(\hat{B}_i)}{\sigma_i}$$

where σ_i is the input sampling standard deviation of the estimated survey biomass. For age or length composition data assumed to follow a multinomial distribution, the normalized residuals for age/length group a in year i were computed as

$$\delta_{i,a} = \frac{(p_{i,a} - \hat{p}_{i,a})}{\sqrt{\hat{p}_{i,a}(1 - \hat{p}_{i,a})/n_i}}$$

where p and \hat{p} are the observed and estimated proportion, respectively, and n is the input assumed sample size for the multinomial distribution. The effective sample size was also computed for the age and length compositions modeled with a multinomial distribution, and for a given year i was computed as

$$E_{i} = \frac{\sum_{a} \hat{p}_{a} (1 - \hat{p}_{a})}{\sum_{a} (\hat{p}_{a} - p_{a})}.$$

An effective sample size that is nearly equal to the input sample size can be interpreted as having a model fit that is consistent with the input sample size.

Parameterization of fishery selectivity

Three models were evaluated that differed in the parameterization for fishery selectivity at age $(s^f_{a,t})$.

Model 1) *Logistic curve varying between 4-year blocks* (used in previous assessments):

A time-varying fishery selectivity curve is used to account for the interannual changes in terms of depth

and management area fished (Tables 5 and 6). Fishery selectivity is modeled with a logistic equation in which deviations are allowed in the parameters specifying the age (a_{50}) and slope (slp) at 50% selection such that the fishing selectivity $s_{a,t}^f$ for age a and year t is modeled as

$$s_a^f = \frac{1}{1 + e^{-(\phi_{asc} + \gamma_t)(a - (a_{50\%} + \eta_t))}}$$

where η_t and γ_t are time-varying deviations that sum to zero and are constrained by adding a lognormal prior to the likelihood function with mean of zero and a CV of 0.1. Deviations in ϕ_{asc} and $a_{50\%}$ allowed between 4-year blocks (i.e., 1964-67, 1968-71, etc.). Little information exists to estimate the selectivity for the 1960-1963 period, so years were combined with those from 1964-1967.

Model 2) Time-invariant double logistic

$$s_a^f = \frac{1}{1 + e^{-\phi_{asc}(a - a_{50\%})}} \frac{1}{1 + e^{-\phi_{des}(a - d_{50\%})}}$$

where fishing selectivity is the product of two logistic curve, and allows for dome-shaped selectivity when the descending slope parameter (ϕ_{des}) is negative.

Model 3) Bicubic spline

A mathematical definition of a spline is a smooth function that is used for either interpolating between fixed points (referred to as "knots" or "nodes") or smoothing a dataset. Splines are of interest when the underlying process for which the spline represents is a smooth, nonlinear function. Splines are constructed from separate piecewise functions that are joined at the knots, and smoothness is ensured by requiring that at each knot, the two functions joined have equal function values, first derivatives, and second derivatives. These conditions can only be met by using polynomial splines of order 3 or higher, and cubic splines are often used because they limit unnecessary bending between the knots. Splines are implemented in non-parametric modeling such as generalized additive models, and been examined in ecological modeling as an approach for modeling time-varying parameters (Thorson et al. 2013). In stock assessment modeling, non-parameteric selectivity curves (a category that includes splines) performed well in an evaluation of various approaches for modeling fishery selectivity (Thorson and Taylor 2013).

The bicubic splines was implemented with the "vcubic_spline_function" function in AD Modelbuilder and models selectivity varying across time and age, and was developed from code provided in Press et al. (1992). Four year knots and 5 age knots were used, for a total of 20 selectivity parameters.

Briefly, the bicubic spline function requires the user to specify a number of age and year nodes that form a grid in the year-age matrix of time-varying selectivity (with equal grid spacing), and values at these nodes are the log-scale fishery selectivity and estimated as parameters. Fishery selectivity at ages and years between the nodes are interpolated with a bicubic spline. The smoothness of the surface is controlled by the number of nodes, and also by a series of penalties estimated within the model. The bicubic spline function was original developed by Dr. Steve Martell for the Integrated Statistical Catch at Age (iSCAM) model, which included penalties for: 1) smoothness across the ages (modeled with the sum of second differences); 2) the slope of the

rate of decline when selectivity decreases with age (modeled with the sum of first differences); and 3) the smoothness across years (modeled with the sum of second differences). In addition to these penalties, an additional penalty on the interannual variability across years (modeled with the first difference) was used in this assessment to address situations in which the selectivity across years was relatively smooth but also non-constant (as would occur with a trend).

Sample sizes for age and length composition data

In previous assessments, the sample sizes for the age and length composition data were set to the square root of the number of fish lengthed or otoliths read. This procedure has resulted in the SDNR for the age and length compositions differing substantially from 1, indicating a mismatch between the precision of the model fit and the assumed input variance.

In this assessment, the sample sizes for the composition data are obtained from an iteratively reweighted procedure using the SDNR (method TA1.2 in Francis 2011). An initial model run in which the sample sizes are specified as in the 2012 assessment is conducted, and a weight that is the inverse of the variance of the normalized residuals for each composition dataset is obtained. The sample sizes for the next model run are the original sample sizes multiplied by the estimated weights, which then produced a new set of weights, and process is iterated until the weights converge. The reweighting was applied to Model 3, and the final sample sizes were applied to Models 1 and 2.

Parameters Estimated Outside the Assessment Model

The parameters estimated independently include the age error matrix, the age-length conversion matrix, and individual weight at age. The calculations for these quantities are described above.

Parameters Estimated Inside the Assessment Model

Parameter estimation is facilitated by comparing the model output to several observed quantities, such as the age and length composition of the survey and fishery catch, the survey biomass, and the catch biomass. The general approach is to assume that deviations between model estimates and observed quantities are attributable to observation error and can be described with statistical distributions. Each data component provides a contribution to a total log-likelihood function, and parameter values that minimize the negative log-likelihood are selected.

The likelihood of the initial recruitments were modeled with a lognormal distribution, yielding the following negative log-likelihood (excluding some constant terms)

$$\lambda_{1} \left[\sum_{t=1}^{n} \frac{\left(v_{t} + \sigma_{r}^{2} / 2 \right)^{2}}{2\sigma_{r}^{2}} + n \ln(\sigma_{r}) \right]$$

where *n* is the number of year where recruitment is estimated. The adjustment of adding $\sigma^2/2$ to the deviation was made in order to produce deviations from the mean, rather than the median, recruitment. If σ_r is fixed, the term $n \ln (\sigma_r)$ adds a constant value to the negative log-likelihood.

The likelihoods of the fishery and survey age and length compositions were modeled with a multinomial distribution. The negative log of the multinomial function (excluding constant terms) for the fishery length composition data, with the addition of a term that scales the likelihood, is

$$-n_{f,t,l} \sum_{s,t,l} (p_{f,t,l} \ln(\hat{p}_{f,t,l}) - p_{f,t,l} \ln(p_{f,t,l}))$$

where n is the reweighted sample size, and $p_{f,t,l}$ and $\hat{p}_{f,t,l}$ are the observed and estimated proportion at length in the fishery by year and length. The likelihood for the age and length proportions in the survey, $p_{surv,t,a}$ and $p_{surv,t,l}$, respectively, follow similar equations.

The negative log-likelihood of the survey biomass was modeled with a lognormal distribution:

$$\lambda_2 \sum_{t} (\ln(obs_biom_t) - \ln(pred_biom_t))^2 / 2cv_t^2$$

where obs_biom_t is the observed survey biomass at time t, cv_t is the coefficient of variation of the survey biomass in year t, and λ_2 is a weighting factor. The negative log-likelihood of the CPUE index is computed in a similar manner, and is weighted by λ_3 . The negative log-likelihood of the catch biomass was modeled with a lognormal distribution:

$$\lambda_4 \sum_t (\ln(obs_cat_t) - \ln(pred_cat_t))^2$$

where obs_cat_t and $pred_cat_t$ are the observed and predicted catch. Because the catch biomass is generally thought to be observed with higher precision that other variables, λ_4 is given a very high weight so as to fit the catch biomass nearly exactly. This can be accomplished by varying the F levels, and the deviations in F are not included in the overall likelihood function.

A maturity ogive was fit within the assessment model to samples collected in 2010 from fishery and survey vessels (n=280; TenBrink and Spencer 2013) and in 2004 by fishery observers (n=165). The samples were analyzed using histological methods. Parameters of the logistic equation were estimated by maximizing the bionomial likelihood within the assessment model. The number of fish sampled and number of mature fish by age for each collection were the input data, thus weighting the two collections by sample size. Due to the low number of young fish, high weights were applied to age 3 and 4 fish in order to preclude the logistic equation from predicting a high proportion of mature fish at age 0. The estimated age at 50% maturity is 9.1 years.

The overall negative log-likelihood function, excluding the priors on *M* and survey catchability, the penalties on time-varying fishery selectivity parameters, and the maturity ogive parameters, is

$$\lambda_{1} \left[\sum_{t=1}^{n} \frac{\left(v_{t} + \sigma_{r}^{2} / 2 \right)^{2}}{2\sigma_{r}^{2}} + n \ln(\sigma_{r}) \right] +$$

$$\lambda_{2} \sum_{t} \left(\ln(obs_biom_{t}) - \ln(pred_biom_{t}) \right)^{2} / 2cv_{t}^{2} +$$

$$\lambda_{3} \sum_{t} \left(\ln(obs_cpue_{t}) - \ln(pred_cpue_{t}) \right)^{2} / 2cv_{CPUE}^{2} +$$

$$- n_{f,t,l} \sum_{s,t,l} (p_{f,t,l} \ln(\hat{p}_{f,t,l}) - p_{f,t,l} \ln(p_{f,t,l})) +$$

$$- n_{f,t,a} \sum_{s,t,l} (p_{f,t,a} \ln(\hat{p}_{f,t,a}) - p_{f,t,a} \ln(p_{f,t,a})) +$$

$$- n_{surv,t,a} \sum_{s,t,a} (p_{surv,t,a} \ln(\hat{p}_{surv,t,a}) - p_{surv,t,a} \ln(p_{surv,t,a})) +$$

$$- n_{surv,t,l} \sum_{s,t,a} (p_{surv,t,l} \ln(\hat{p}_{surv,t,l}) - p_{surv,t,l} \ln(p_{surv,t,l})) +$$

$$\lambda_{4} \sum_{t} \left(\ln(obs_cat_{t}) - \ln(pred_cat_{t}) \right)^{2}$$

For the model run in this analysis, λ_1 , λ_2 , λ_3 , and λ_4 were assigned weights of 1,1, 0.5, and 500, reflecting a strong emphasis on fitting the catch data and a de-emphasis of the CPUE index. The negative log-likelihood function was minimized by varying the following parameters (using the bicubic fishery selectivity):

| Parameter type | Number |
|-----------------------------------|--------|
| 1) Fishing mortality mean | 1 |
| 2) Fishing mortality deviations | 55 |
| 3) Recruitment mean | 1 |
| 4) Recruitment deviations | 52 |
| 5) Unfished recruitment | 1 |
| 6) Biomass survey catchability | 1 |
| 7) CPUE index catchability | 1 |
| 8) Fishery selectivity parameters | 20 |
| 10) Survey selectivity parameters | 2 |
| 11) Natural mortality rate | 1 |
| 12) Maturity parameters | 2 |
| Total parameters | 137 |

Finally, a Monte Carlo Markov Chain (MCMC) algorithm was used to obtain estimates of parameter uncertainty (Gelman et al. 1995). One million MCMC simulations were conducted, with every 1,000th sample saved for the sample from the posterior distribution after excluding the first 50,000 simulations. Ninety-five percent confidence intervals were produced as the values corresponding to the 5th and 95th percentiles of the MCMC evaluation. For this assessment, confidence intervals on total biomass, spawning biomass, and recruitment strength are presented.

Results

Model Evaluation

Several attributes of the model fits are shown in Table 12. Models 0 and 0.1 are presented for to demonstrate intermediate steps between the 2012 model and the recommended 2014 model (i.e., a "bridging" analysis). Model 0 has the updated data through 2014, Model 0.1 excludes the cooperative survey biomass estimates and age/size composition data, and each uses the age and length composition sample weights as produced for the 2012 assessment. The sample sizes for the composition are identical in Models 1-3, and were produced by applying iterative reweighting to Model 3 (bicubic spline selectivity). Akaike Information Criterion (AIC) and the Bayesian Information Criterion (BIC) were used to evaluate model selection. Each of these metrics penalize the negative log-likelihood by multiple of the number of parameters; in AIC, this multiple is 2 whereas in BIC is the natural log of the number of data points.

For all the models, the number of parameters is "nominal" number of parameters, which overestimates the number of independent parameters because of the use of penalties and prior distributions in the models. Deviance Information Criterion (DIC) could be used, but will often select the models with higher number of parameters (Martell and Stewart 2014). For these reasons, model selection considered additional information such as the root mean squared errors and negative log-likelihoods in the fits to the data, and the residual patterns in fitting the composition data.

Model 3 (bicubic spline selectivity) had the lowest AIC and BIC, and produced better fits to the fishery length composition and survey age composition data than Models 1 and 2 (as revealed by both the negative log-likelihoods and the RMSE). Model 2 has the lowest number of parameters, and was considered to evaluate whether the lack of fit to the age-plus group in the fishery and survey data is more related to the selectivity shape its variation over time. Both appear to be important, as the improvement in Model 3 is achieved by allowing the dome-shaped selectivity curves to vary over time. Additionally, Model 3 achieves to the best fit the time series of survey biomass estimates, which show a strong trend that is informative on the dynamics of the stock.

The estimated spawning biomass for the models is shown in Figure 8, with the bridging Models 0 and 0.1 shown in blue. Relative to the bridging models, Model 3 estimates higher biomass in the early 1980s (resulting from the dome-shaped selectivity in the 1960s and 1970s), and relative similar levels of biomass in the recent years (resulting from the shift from dome-shaped to near asymptotic selectivity).

Model 3 was selected as the preferred model, and the results below were obtained from this model.

Time series results

In this assessment, spawning biomass is defined as the biomass estimate of mature females age 3 and older. Total biomass is defined as the biomass estimate of POP age 3 and older. Recruitment is defined as the number of age 3 POP.

A retrospective analysis was conducted to evaluate the effect of recent data on estimated spawning stock biomass. For the current assessment model, a series of model runs were conducted in which the end year of the model was varied from 2012 to 2002, and this was accomplished by sequentially dropping age and length composition data, survey biomass estimates, and catch estimates from the input data files.

The plot of retrospective estimates of spawning biomass is shown in Figure 9. The largest changes in estimated survey biomass occurred with end years 2004, 2006, 2010, and 2012, when survey biomass estimates and survey age composition data are added to the model. The 2014 survey contains similar

information as in the 2012 survey and thus has relatively little effect on the retrospective pattern.

The change in estimated spawning biomass from the 2009 to 2010 end years was particularly large, as the 2010 survey biomass estimate was substantially increased from the 2006 estimate. A series of exploratory models runs conducted in the 2010 assessment revealed that a combination of the high survey biomass and new observations of strong 1994-2000 year classes observed in both the fishery and survey age and length composition data lowered the estimates of survey catchability and increased estimated biomass.

Mohn's rho can be used to evaluate the severity of any retrospective pattern, and compares an estimated quantity (in this case, spawning stock biomass) in the terminal year of each retrospective model run with the estimated quantity in the same year of the model using the full data set . The absence of any retrospective pattern would result in a Mohn's rho of 0, and would result from either identical estimates in the model runs, or from positive deviations from the reference model being offset by negative deviations. The Mohn's rho for this retrospective runs was -3.43.

The practice of estimating survey catchability within the model likely contributes to the retrospective pattern. The increasing trend in survey biomass cannot be explained by simply increasing the estimates of recruitments in recent years, but instead require an increase in the recruitment estimates for many year classes. This increases the scale of the biomass estimates, and thus lowers the survey catchability. The estimates of survey catchability is greater than 2 for the retrospective runs with 2004 and 2005 end years, and gradually decrease to the estimate of 1.28 obtained with data through 2014. The sensitivity of estimated survey catchability to the sequential addition of survey biomass estimates and age/length composition data motivates the estimation of survey catchability outside the model.

Prior and Posterior Distributions

Posterior distributions for M, q, total 2014 biomass, and median recruitment, based upon the MCMC integrations, are shown in Figure 10. The posterior distribution for M shows little overlap with the prior distribution. However, a sensitivity analysis on the estimation on M (presented in the Executive Summary) indicates that the estimate of M obtained without a prior would be 0.07 (with fixing survey catchabilitity at its current value), and this value of M is also supported by a recent meta-analysis relating estimates of natural mortality to maximum age (Then at al. 2014). In future assessment, increasing the expected value of the prior distribution for M from 0.05 to 0.07 will be considered.

Biomass Trends

The estimated survey biomass index begins with 1,304,450 t in 1960, declines to 238,512 t in 1981, and increases to 751,086 t in 2010, and declines to 671,125 t in 2012 (Figure 11). The survey point estimates are used in a relative sense rather than in an absolute sense, with a survey catchability (q) estimated at 1.28 rather than fixed at 1.0, which is to 16% increase from the value of 1.10 in the 2012 assessment. The model estimate of survey biomass still does not match the high 2010 - 2014 survey biomass estimates very well. Because the AI survey biomass estimates are taken as an index for the entire BSAI area, one might expect that q would be below 1.0 to the extent that the total BSAI biomass is higher than the Aleutian Islands biomass. One factor that may cause an increase in survey catchability is the expansion of survey trawl estimates to untrawlable areas (Kreiger and Sigler 1996). The fit to the CPUE index is shown in Figure 12.

The total biomass showed a similar trend as the survey biomass, with the 2014 total biomass estimated as 597,506 t. The estimated time series of total biomass and spawning biomass, with 95% credibility bounds obtained from MCMC integration, are shown in Figure 13. Total biomass, spawning biomass, and recruitment are given in Table 13.

Age/size compositions

The fits to the fishery age and length composition is shown in Figures 14-15. The observed proportion in the binned age 25+ group for years 1981 and 1982 is higher than the estimated proportion, although the fits improve for the remainder of the fishery age compositions. The observed proportion in the binned length group of 39+ cm for 1964 and 1965 was lower than the estimated proportion, reflecting the modeling of the initial numbers at age as an equilibrium population. However, by 1966 reasonable fits were observed for the binned length group in the fishery length composition (Figure 15). Some of the lack of fit in the mid- to late-1980s is attributable to the low sample size of lengths observed from a reduced fishery. Good fits are obtained for most age groups in the 1991-2012 surveys (Figure 16). The model provides a reasonable fit to the 2014 survey length composition (Figure 17).

Fishing and Survey Selectivity

The estimated age at 50% selection for the survey is 6.02 (Figure 18). The estimated fishery selectivity by age and year is shown in Figure 19, and shows pattern consistent with the empirical data in fishery catch examined above. Strong dome-shaped selectivity is estimated in the early 1960s to allow fish of age 20 older from this period to survive the large fully-selected fishing rates in the 1960s and early 1970s and be available for capture in the fishery and survey in the early 1980s (by which time they have entered the 40+ group). The model estimates that dome-shaped selectivity has decreased over time (with the exception of dome-shaped selectivity for ages 35 and above since 2010).

Fishing Mortality

The estimates of instantaneous fishing mortality for POP range from highs during the 1970's to low levels in the 1980's (Figure 20). Fishing mortality rates since the early 1980's, however, have moderated considerably due to the phase out of the foreign fleets and quota limitations imposed by the North Pacific Fishery Management Council. Note that because of the change in the fishery selectivity over time, the fully-selected rates are not completely comparable over time with respect to the degree to which the stock has been harvested. Nonetheless, the average fully-selected fishing mortality from 1965 to 1980 was 0.57, whereas the average from 1981 to 2011 was 0.04.

The plot of estimated fishing mortality rates and spawning stock biomass relative to the harvest control rules (Figure 21) indicate that BSAI POP would be considered overfished (using current definitions) during much of the period from the mid-1960s to the mid-1980s, although it should be noted the current definitions of $B_{35\%}$ are based on the estimated recruitment of the post-1977 year classes and the average fishery selectivity from the most recent 5 years.

Recruitment

Year-class strength varies widely for BSAI POP (Figure 22; Table 12). The relationship between spawning stock and recruitment also displays a high degree of variability (Figure 23). The 1957 and 1962 year classes are particularly large and sustained the heavy fishing in the 1960s. The rebuilding of the stock in the 1980s and 1990s was based upon recruitments for the 1981, 1984, 1986, and 1988 year classes. Recruitment appears to be lower in early 1990s, but cohorts from 1994 to 2000 generally show relatively strong recruitment (with the exception the 1997 and 1999 year classes), which is consistent with the increasing trend of biomass and the fishery and survey age compositions shown in Figures 14 and 16.

Harvest recommendations

Amendment 56 reference points

The reference fishing mortality rate for Pacific ocean perch is determined by the amount of reliable population information available (Amendment 56 of the Fishery Management Plan for the groundfish

fishery of the Bering Sea/Aleutian Islands). Estimates of $F_{0.40}$, $F_{0.35}$, and $SPR_{0.40}$ were obtained from a spawner-per-recruit analysis. Assuming that the average recruitment from the 1977-2011 year classes estimated in this assessment represents a reliable estimate of equilibrium recruitment, then an estimate of $B_{0.40}$ is calculated as the product of $SPR_{0.40}$ * equilibrium recruits, and this quantity is 169,203 t. The year 2015 estimated spawning stock biomass is 234,426 t.

Specification of OFL and maximum permissible ABC

Since reliable estimates of the 2015 spawning biomass (B), $B_{0.40}$, $F_{0.40}$, and $F_{0.35}$ exist and $B>B_{0.40}$ (234,426 t > 169,203 t), POP reference fishing mortality have been classified in tier 3a. For this tier, F_{ABC} maximum permissible F_{ABC} is $F_{0.40}$, and F_{OFL} is equal to $F_{0.35}$. The values of $F_{0.40}$ and $F_{0.35}$ are 0.089 and 0.109, respectively.

The 2015 ABC associated with the $F_{0.40}$ level of 0.089 is 34,988 t.

The estimated catch level for year 2015 associated with the overfishing level of F = 0.074 is 42,558 t. A summary of these values is below.

| 2015 SSB estimate (B) | = | 234,426 t |
|-----------------------|---|-----------|
| $B_{0.40}$ | = | 169,203 t |
| $F_{ABC} = F_{0.40}$ | = | 0.089 |
| $F_{OFL} = F_{0.35}$ | = | 0.109 |
| MaxPermABC | = | 34,988 t |
| OFL | = | 42,558 t |

ABC recommendation

We recommend the maximum permissible ABC 34,988 t.

Projections

A standard set of projections is conducted for each stock managed under Tiers 1, 2, or 3 of Amendment 56. This set of projections encompasses seven harvest scenarios designed to satisfy the requirements of Amendment 56, the National Environmental Policy Act, and the Magnuson-Stevens Fishery Conservation and Management Act (MSFCMA).

For each scenario, the projections begin with the vector of 2014 numbers at age estimated in the assessment. This vector is then projected forward to the beginning of 2015 using the schedules of natural mortality and selectivity described in the assessment and the best available estimate of total (year-end) catch for 2014. In each subsequent year, the fishing mortality rate is prescribed on the basis of the spawning biomass in that year and the respective harvest scenario. In each year, recruitment is drawn from an inverse Gaussian distribution whose parameters consist of maximum likelihood estimates determined from recruitments estimated in the assessment. Spawning biomass is computed in each year based on the time of peak spawning and the maturity and weight schedules described in the assessment. Total catch is assumed to equal the catch associated with the respective harvest scenario in all years. This projection scheme is run 1000 times to obtain distributions of possible future stock sizes, fishing mortality rates, and catches.

Five of the seven standard scenarios will be used in an Environmental Assessment prepared in conjunction with the final SAFE. These five scenarios, which are designed to provide a range of harvest alternatives that are likely to bracket the final TAC for 2011, are as follow (" $max F_{ABC}$ " refers to the maximum permissible value of F_{ABC} under Amendment 56):

Scenario 1: In all future years, F is set equal to $max F_{ABC}$. (Rationale: Historically, TAC has been constrained by ABC, so this scenario provides a likely upper limit on future TACs.)

Scenario 2: In all future years, F is set equal to a constant fraction of $max F_{ABC}$, where this fraction is equal to the ratio of the F_{ABC} value for 2015 recommended in the assessment to the $max F_{ABC}$ for 2015. (Rationale: When F_{ABC} is set at a value below $max F_{ABC}$, it is often set at the value recommended in the stock assessment.)

Scenario 3: In all future years, F is set equal to 50% of max F_{ABC} . (Rationale: This scenario provides a likely lower bound on F_{ABC} that still allows future harvest rates to be adjusted downward when stocks fall below reference levels.)

Scenario 4: In all future years, F is set equal to the 2009-2013 average F. (Rationale: For some stocks, TAC can be well below ABC, and recent average F may provide a better indicator of F_{TAC} than F_{ABC} .)

Scenario 5: In all future years, F is set equal to zero. (Rationale: In extreme cases, TAC may be set at a level close to zero.)

Two other scenarios are needed to satisfy the MSFCMA's requirement to determine whether the Pacific ocean perch stock is currently in an overfished condition or is approaching an overfished condition. These two scenarios are as follow (for Tier 3 stocks, the MSY level is defined as $B_{35\%}$):

Scenario 6: In all future years, F is set equal to F_{OFL} . (Rationale: This scenario determines whether a stock is overfished. If the stock is expected to be above 1) above its MSY level in 2014 or 2) above $\frac{1}{2}$ of its MSY level in 2014 and above its MSY level in 2014 under this scenario, then the stock is not overfished.)

Scenario 7: In 2015 and 2016, F is set equal to $max F_{ABC}$, and in all subsequent years F is set equal to F_{OFL} . (Rationale: This scenario determines whether a stock is approaching an overfished condition. If the stock is expected to be above its MSY level in 2027 under this scenario, then the stock is not approaching an overfished condition.)

The recommended F_{ABC} and the maximum F_{ABC} are equivalent in this assessment, and projections of the mean harvest and spawning stock biomass for the remaining six scenarios are shown in Table 14.

Status Determination

In addition to the seven standard harvest scenarios, Amendments 48/48 to the BSAI and GOA Groundfish Fishery Management Plans require projections of the likely OFL two years into the future. While Scenario 6 gives the best estimate of OFL for 2015, it does not provide the best estimate of OFL for 2016, because the mean 2015 catch under Scenario 6 is predicated on the 2015 catch being equal to the 2015 OFL, whereas the actual 2015 catch will likely be less than the 2015 OFL. The executive summary contains the appropriate one- and two-year ahead projections for both ABC and OFL.

Under the MSFCMA, the Secretary of Commerce is required to report on the status of each U.S. fishery with respect to overfishing. This report involves the answers to three questions: 1) Is the stock being subjected to overfishing? 2) Is the stock currently overfished? 3) Is the stock approaching an overfished condition?

Is the stock being subjected to overfishing? The official BSAI catch estimate for the most recent complete year (2013) is 31,393 t. This is less than the 2013 BSAI OFL of 41,900 t. Therefore, the stock is not being subjected to overfishing.

Harvest Scenarios #6 and #7 are intended to permit determination of the status of a stock with respect to its minimum stock size threshold (MSST). Any stock that is below its MSST is defined to be *overfished*. Any stock that is expected to fall below its MSST in the next two years is defined to be *approaching* an overfished condition. Harvest Scenarios #6 and #7 are used in these determinations as follows:

Is the stock currently overfished? This depends on the stock's estimated spawning biomass in 2014:

- a. If spawning biomass for 2014 is estimated to be below 1/2 B35%, the stock is below its MSST.
- b. If spawning biomass for 2014 is estimated to be above B35% the stock is above its MSST.
- c. If spawning biomass for 2014 is estimated to be above ½ *B35%* but below *B35%*, the stock's status relative to MSST is determined by referring to harvest Scenario #6 (Table 14). If the mean spawning biomass for 2024 is below *B35%*, the stock is below its MSST. Otherwise, the stock is above its MSST.

Is the stock approaching an overfished condition? This is determined by referring to harvest Scenario #7:

- a. If the mean spawning biomass for 2017 is below 1/2 *B*35%, the stock is approaching an overfished condition.
- b. If the mean spawning biomass for 2017 is above $B_{35\%}$, the stock is not approaching an overfished condition.
- c. If the mean spawning biomass for 2017 is above 1/2 *B35%* but below *B35%*, the determination depends on the mean spawning biomass for 2027. If the mean spawning biomass for 2027 is below *B35%*, the stock is approaching an overfished condition. Otherwise, the stock is not approaching an overfished condition.

The results of these two scenarios indicate that the BSAI POP stock is neither overfished nor approaching an overfished condition. With regard whether the stock is currently overfished, the expected stock size in the year 2014 of Scenario 6 is 1.66 times its $B_{35\%}$ value of 148,053 t. With regard to whether the BSAI POP stock is likely to be overfished in the future, the expected stock size in 2017 of Scenario 7 is 1.41 times the $B_{35\%}$ value.

Area Allocation of Harvests

The ABC of BSAI POP is currently partitioned into subarea ABCs based on the relative biomass from research surveys. The current method of obtaining the subarea is a weighted average was applied to the AI trawl surveys in order to compute the average biomass from each of the four subareas, and applies weights of 4, 6, and 9 to the 2010, 2012, and 2014 surveys. A weighted average was also applied to EBS slope survey estimates, with weights of 4, 6, and 9 applied to 2008, 2010, and 2012 surveys.

It is also of interest to estimate the area proportions using the random effects model. The survey averaging workgroup is evaluating the use of the random effects model to smooth survey time series for computing area apportionments, and its use for computing area proportion might logically be delayed until after the workgroup has completed their evaluations. However, a comparison between the two methods may be useful. A comparison between the biomass estimates and proportions is shown below, with the average biomass in the EBS area being the sum of the estimate from the AI SBWS area and the EBS slope area.

| | WAI | CAI | EAI | EBS |
|-----------------------------|---------|---------|---------|---------|
| Weighted average biomass | 326,939 | 269,931 | 282,471 | 266,753 |
| Proportion of biomass | 28.5% | 23.6% | 24.6% | 23.3% |
| | | | | |
| | | | | |
| Estimated 2014 biomass | | | | |
| (from random effects model) | 311,678 | 236,416 | 254,448 | 268,506 |
| Proportion of biomass | 29.1% | 22.1% | 23.8% | 25.1% |

The apportionments for the 2015 and 2016 ABC are shown below:

| | BSAI | WAI | CAI | EAI | EBS | Total |
|------------------------------|--------|--------|-------|-------|-------|--------|
| OFL (2015) | 42,558 | | | | | |
| ABC (2015, weighted average) | | 9,981 | 8,240 | 8,623 | 8,143 | 34,988 |
| ABC (2015, RE model) | | 10,182 | 7,723 | 8,312 | 8,771 | 34,988 |
| OFL (2016) | 40,809 | | | | | |
| ABC (2016, weighted average) | | 9,571 | 7,902 | 8,269 | 7,809 | 33,550 |
| ABC (2016, RE model) | | 9,763 | 7,406 | 7,970 | 8,411 | 33,550 |

Ecosystem Considerations

Ecosystem Effects on the stock

1) Prey availability/abundance trends

POP feed upon calanoid copepods, euphausids, myctophids, and other miscellaneous prey (Yang 2003). From a sample of 292 Aleutian Island specimens collected in 1997, calanoid copepods, euphausids, and myctophids contributed 70% of the total diet by weight. The diet of small POP was composed primarily of calanoid copepods (89% by weight), with euphausids and myctophids contributing approximately 35% and 10% of the diet, respectively, of larger POP. The availability and abundance trends of these prey species are unknown.

2) Predator population trends

POP are not commonly observed in field samples of stomach contents, although previous studies have identified sablefish, Pacific halibut, and sperm whales as predators (Major and Shippen 1970). The population trends of these predators can be found in separate chapters within this SAFE document.

3) Changes in habitat quality

POP appear to exhibit ontogenetic shifts in habitat use. Carlson and Straty (1981) used a submersible off southeast Alaska to observe juvenile red rockfish they believed to be POP at approximately 90-100 m in rugged habitat including boulder fields and rocky pinnacles. Kreiger (1993) also used a submersible to observe that the highest densities of small red rockfish in untrawlable rough habitat. As POP mature, they move into deeper and less rough habitats. Length frequencies of the Aleutian Islands survey data indicate that large POP (> 25 cm) are generally found at depths greater than 150 m. Brodeur (2001) also found that POP was associated with epibenthic sea pens and sea whips along the Bering Sea slope. There has been little information identifying how rockfish habitat quality has changed over time.

Fishery Effects on the ecosystem

Catch of prohibited species from 2003-2008 by fishery are available from the NMFS Regional Office. The rockfish fishery in the BSAI area, which consists only of the AI POP target fishery, contributed approximately 2% of the gold/brown king crab catch and approximately 1% of the halibut bycatch. For other prohibited species, the BSAI rockfish fisheries contributed much lower that 1% of the bycatch.

Estimates of non-target catches in the rockfish fishery are also available from the Catch Accounting System database maintained by the NMFS Regional Office. BSAI rockfish fisheries contribute mostly to the bycatch of coral, sponge, and polychaetes. From 2003 to 2008, the BSAI rockfish fisheries contributed 31% of the coral and bryozoan bycatch,18% of the sponge bycatch, 8% of the red tree coral bycatch, and 7% of the polychaete bycatch. The relative contribution was variable between years; for example, the annual relative contribution corals and bryozoans ranged from 5% in 2004 to 53% in 2003, and the other groups listed above show similar levels of variability.

The POP fishery is not likely to diminish the amount of POP available as prey due to its low selectivity for fish less than 27 cm. Additionally, the fishery is not suspected of affecting the size-structure of the population due to the relatively light fishing mortality, averaging 0.04 over the last 5 years. It is not known what effects the fishery may have on the maturity-at-age of POP.

Data Gaps and Research Priorities

Although Pacific ocean perch may be considered a "data-rich" species relative to other rockfish, little information is known regarding most aspects of their biology, including reproductive biology and the distribution, duration, and habitat requirements of various life-history stages. Given the relatively unusual reproductive biology of rockfish and its importance in establishing management reference points, data on reproductive capacity should be collected on a periodic basis.

References

- Archibald, C. P., W. Shaw, and B. M. Leaman. 1981. Growth and mortality estimates of rockfishes (Scorpaenidae) from B.C. coastal waters, 1977-79. Can. Tech. Rep. Fish. Aquat. Sci. 1048, 57 p.
- Chikuni, S. 1975. Biological study on the population of the Pacific ocean perch in the North Pacific. Bull. Far Seas Fish. Res. Lab. (Shimizu) 12:1-119.
- Chilton, D. E., and R. J. Beamish. 1982. Age determination methods for fishes studied by the Groundfish Program at the Pacific Biological Station. Can. Spec. Publ. Fish. Aquat. Sci. 60, 102 p.
- Dorn, M.W. 2002. Advice on west coast rockfish harvest rates from Bayesian meta-analysis of stock-recruitment relationships. N. Am. J. Fish. Aquat. Sci. 22:280-300.
- Gelman, A., J.B. Carlin, H.S. Stern, and D.A. Rubin. 1995. Bayesian data analysis. Chapman and Hall, New York. 552 pp.
- Gharrett, A.J., A.K. Gray, and J. Heifetz. 2001. Identification of rockfish (*Sebastes* spp.) by restriction site analysis of the mitochondrial ND-3/ND-4 and 12S/16S rDNA gene regions. Fish. Bull. 99:49-62.
- Gunderson, D.R. 1972. Evidence that Pacific ocean perch (*Sebastes alutus*) in Queen Charlotte Sound for aggregations that have different biological characteristics. J. Fish. Res Brd. Can. 29:1061-1070

- Gunderson, D. R. 1977. Population biology of Pacific ocean perch, *Sebastes alutus*, stocks in the Washington-Queen Charlotte Sound region, and their response to fishing. Fish. Bull., U.S. 75(2): 369-403.
- Ianelli, J. N., and D. H. Ito. 1991. Stock assessment of Pacific ocean perch (*Sebastes alutus*) using an explicit age structured model. *In* Stock assessment and fishery evaluation report for the groundfish resources of the Bering Sea/Aleutian Islands region as projected for 1992 (November 1991), 20 pp. North Pacific Fishery Management Council, P.O. Box 103136, Anchorage, AK 99510.
- Ianelli, J. N., and D. H. Ito. 1992. Pacific ocean perch. *In* Stock assessment and fishery evaluation report for the groundfish resources of the Bering Sea/Aleutian Islands region as projected for 1993 (November 1992), 36 pp. North Pacific Fishery Management Council, P.O. Box 103136, Anchorage, AK 99510.
- Ito, D. H. 1982. A cohort analysis of Pacific ocean perch stocks from the Gulf of Alaska and Bering Sea regions. NWAFC Processed Rep. 82-15, 157 p. Northwest and Alaska Fish. Cent., Natl. Mar. Fish. Serv., NOAA, 7600 Sand Point Way N.E., Bin C15700, Seattle, WA 98115.
- Ito, D. H. 1986. Pacific ocean perch. *In* R. G. Bakkala and L. L. Low (editors), Condition of groundfish resources of the eastern Bering Sea and Aleutian Islands region in 1985, p. 101-132. U.S. Dep. Commer., NOAA Tech. Memo. NMFS F/NWC-104.
- Kendall, A.W. Jr. 1991. Systematics and identification of larvae and juveniles of the genus *Sebastes*. Env. Biol. Fish. 30:173-190.
- Kimura, D. K., and J. J. Lyons. 1991. Between-reader bias and variability in the age-determination process. Fish. Bull., U.S. 89: 53-60.
- Krieger, K. J., and M. F. Sigler. 1996. Catchability coefficient for rockfish estimated from trawl and submersible surveys. Fish. Bull., U.S. 94: 282-288.
- Megrey, B.A. and V.G. Wespestad. 1990. Alaskan groundfish resources: 10 years of management under the Magnuson Fishery Conservation and Management Act. North American Journal of Fisheries Management 10:125-143.
- Park, L.K. and P. Moran. 1994. Developments in molecular genetic techniques in fisheries. Reviews in Fish Biology and Fisheries 4:272-299.
- Rocha-Olivares, A. 1998. Multiplex haplotype-specific PCR: a new approach for species identification of the early life stages of rockfishes of the species-rich genus *Sebastes* Cuvier. J. Exp. Mar. Biol. Ecol. 231:279-290.
- Seeb, L.W. and D.R. Gunderson. 1988. Genetic variation and population structure of of Pacific ocean perch (*Sebastes alutus*). Can J. Fish. Aquat. Sci. 45:78-88.
- Seeb, L.W. and A.W. Kendall, Jr. 1991. Allozyme polymorphisms permit the identification of larval and juvenile rockfishes of the genus *Sebastes*. Env. Biol. Fish. 30:191-201.
- Spencer, P.D., and J.N. Ianelli. 2012. Assessment of the Pacific ocean perch stock in the eastern Bering Sea and Aleutian Islands. In Stock assessment and fishery evaluation report for the groundfish resources of the Bering Sea/Aleutian Islands regions, pp. 1291-1348. North Pacific Fishery Management Council, 605 W. 4th Ave, suite 306. Anchorage, AK 99501.

- Stockhausen, W. and A. Hermann. 2007. Modeling larval dispersion of rockfish: A tool for marine reserve design? In: J. Heifetz, J. DiCosimo, A.J. Gharrett, M.S. Love, T. O'Connell, and R. Stanley (eds.), Biology, assessment, and management of North Pacific rockfishes, pp. 251-273. Alaska Sea Grant College Program, University of Alaska Fairbanks.
- Tagart, J.V. 1984. Comparison of final ages assigned to a common set of Pacific ocean perch otoliths. Washington Department of Fisheries Technical Report 81, 36 pp. Olympia, WA.
- Westrheim, S.J. 1970. Survey of rockfishes, especially of Pacific ocean perch, in the northeast Pacific ocean, 1963-66. J. Fish. Res. Brd. Can. 27:1781-1809.
- Westrheim, S.J. 1973. Age determination and growth of Pacific ocean perch (*Sebastes alutus*) in the northeast Pacific ocean. J. Fish. Res. Brd. Can. 30:235-247.
- Withler, R.E., T.D. Beacham, A.D. Schulze, L.J. Richards, and K.M. Miller. 2001. Co-existing populations of Pacific ocean perch, *Sebastes alutus*, in Queen Charlotte Sound, British Columbia. Mar. Biol. 139:1-12.
- Yang, M-S. 1996. Diets of the important groundfishes in the Aleutian Islands in summer 1991. U.S. Dep. Commer., NOAA Tech. Memo. NMFS-AFSC-60. 105 pp.
- Yang, M.S. 2003. Food habits of the important groundfishes in the Aleutian Islands in 1994 and 1997. U.S. Dep. Commer., AFSC Proc. Rep 2003-07. 233 pp.

Table 1. Total allowable catch (TAC), acceptable biological catch (ABC), and catch of the species groups used to manage Pacific ocean perch from 1977 to 2001 in the Aleutian Islands and the eastern Bering Sea. The "POP complex" includes the other red rockfish species (shortraker rockfish, rougheye rockfish, northern rockfish, and sharpchin rockfish) plus POP.

| | Aleutian Islands | | | | | | | Eastern Bo | ering Sea | |
|------|------------------|---------|---------|---------|-----------|-------------|---------|------------|-----------|-----------|
| | Management | | | | | Management | | | | |
| Year | Group | OFL (t) | ABC (t) | TAC (t) | Catch (t) | Group | OFL (t) | ABC (t) | TAC (t) | Catch (t) |
| 1977 | 7 POP | | | | 7927 | POP | | | | 2406 |
| 1978 | 3 POP | | | | 5286 | POP | | | | 2230 |
| 1979 | 9 POP | | | | 5486 | POP | | | | 1722 |
| 1980 |) POP | | | | 4010 | POP | | | | 959 |
| 1981 | 1 POP | | | | 3668 | POP | | | | 1186 |
| 1982 | 2 POP complex | | | | 979 | POP complex | | | | 205 |
| 1983 | 3 POP complex | | | | 471 | POP complex | | | | 192 |
| 1984 | 4 POP complex | | | | 564 | POP complex | | | | 315 |
| 1985 | 5 POP complex | | | | 216 | POP complex | | | | 61 |
| 1986 | 5 POP | | | 6800 | 302 | POP | | | 825 | 670 |
| 1987 | 7 POP | | | 8175 | 1055 | POP | | | 2850 | 1178 |
| 1988 | 3 POP | | 16600 | 6000 | 2024 | POP | | 6000 | 5000 | 1326 |
| 1989 | POP complex | | 16600 | 6000 | 2963 | POP complex | | 6000 | 5000 | 2533 |
| 1990 | POP complex | | 16600 | 6000 | 11826 | POP complex | | 6300 | 6300 | 6499 |
| 1991 | 1 POP | | 10775 | 10775 | 2785 | POP | | 4570 | 4570 | 5099 |
| 1992 | 2 POP | 11700 | 11700 | 11700 | 10280 | POP | 3540 | 3540 | 3540 | 3255 |
| 1993 | 3 POP | 16800 | 13900 | 13900 | 13376 | POP | 3750 | 3330 | 3330 | 3764 |
| 1994 | 4 POP | 16600 | 10900 | 10900 | 10866 | POP | 2920 | 1910 | 1910 | 1688 |
| 1995 | 5 POP | 15900 | 10500 | 10500 | 10304 | POP | 2910 | 1850 | 1850 | 1208 |
| 1996 | 5 POP | 25200 | 12100 | 12100 | 12827 | POP | 2860 | 1800 | 1800 | 2855 |
| | 7 POP | 25300 | 12800 | 12800 | 12648 | POP | 5400 | 2800 | 2800 | 681 |
| | 3 POP | 20700 | 12100 | 12100 | 9299 | POP | 3300 | 1400 | 1400 | 1022 |
| 1999 | 9 POP | 19100 | 13500 | 13500 | 12484 | POP | 3600 | 1900 | 1400 | 421 |
| |) POP | 14400 | 12300 | 12300 | 9328 | POP | 3100 | 2600 | 2600 | 451 |
| 2001 | 1 POP | 11800 | 10200 | 10200 | 8557 | POP | 2040 | 1730 | 1730 | 896 |

Table 2. Total allowable catch (TAC), acceptable biological catch (ABC), and catch for BSAI POP from 2002 to present. Catch data is through October 11, 2014, from NMFS Alaska Regional Office.

Bering Sea/Aleutian Islands

| | | 24. | 2.5 2 2 2 2 2 2 | - 00010011 101001 | 1000 |
|------|------------|---------|-----------------|-------------------|-----------|
| | Management | | | | |
| Year | Group | OFL (t) | ABC (t) | TAC (t) | Catch (t) |
| 2002 | 2 POP | 17500 | 14800 | 14800 | 11215 |
| 2003 | 3 POP | 18000 | 15100 | 14100 | 14744 |
| 2004 | 4 POP | 15800 | 13300 | 12580 | 11896 |
| 2005 | 5 POP | 17300 | 14600 | 12600 | 10427 |
| 2006 | 5 POP | 17600 | 14800 | 12600 | 12867 |
| 2007 | 7 POP | 26100 | 21900 | 19900 | 18451 |
| 2008 | 8 POP | 25700 | 21700 | 21700 | 17436 |
| 2009 | 9 POP | 22300 | 18800 | 18800 | 15347 |
| 2010 |) POP | 22400 | 18860 | 18860 | 17852 |
| 201 | 1 POP | 36300 | 24700 | 24700 | 24004 |
| 2012 | 2 POP | 35000 | 24700 | 24700 | 24143 |
| 2013 | 3 POP | 41900 | 35100 | 35100 | 31393 |
| 2014 | 4 POP | 39585 | 33122 | 33122 | 25889 |

Table 3. Total allowable catch (TAC), acceptable biological catch (ABC), and catch of POP by area and management group from 1977 to 2014.

| | East | tern Bering | Sea | Al | eutian Islar | nds | BSAI |
|-------|---------|-------------|-------|---------|--------------|--------|-------------|
| Year | Foreign | JVP | DAP | Foreign | JVP | DAP | Total catch |
| 1977 | 2,406 | 0 | | 7,927 | 0 | | 10,333 |
| 1978 | 2,230 | 0 | | 5,286 | 0 | | 7,516 |
| 1979 | 1,722 | 0 | | 5,486 | 0 | | 7,208 |
| 1980 | 907 | 52 | | 4,010 | 0 | | 4,969 |
| 1981 | 1,185 | 1 | | 3,668 | 0 | | 4,854 |
| 1982 | 186 | 19 | | 977 | 2 | | 1,183 |
| 1983 | 99 | 93 | | 463 | 8 | | 663 |
| 1984 | 172 | 142 | | 324 | 241 | | 879 |
| 1985 | 30 | 31 | | 0 | 216 | | 277 |
| 1986 | 18 | 103 | 549 | 0 | 163 | 139 | 972 |
| 1987 | 5 | 49 | 1,123 | 0 | 502 | 554 | 2,233 |
| 1988 | 0 | 46 | 1,280 | 0 | 1,512 | 512 | 3,350 |
| 1989 | 0 | 26 | 2,507 | 0 | 0 | 2,963 | 5,496 |
| 1990 | | | 6,499 | | | 11,826 | 18,324 |
| 1991 | | | 5,099 | | | 2,785 | 7,884 |
| 1992 | | | 3,255 | | | 10,280 | 13,534 |
| 1993 | | | 3,764 | | | 13,376 | 17,139 |
| 1994 | | | 1,688 | | | 10,866 | 12,554 |
| 1995 | | | 1,208 | | | 10,304 | 11,511 |
| 1996 | | | 2,855 | | | 12,827 | 15,681 |
| 1997 | | | 681 | | | 12,648 | 13,329 |
| 1998 | | | 1,022 | | | 9,299 | 10,320 |
| 1999 | | | 421 | | | 12,484 | 12,905 |
| 2000 | | | 451 | | | 9,328 | 9,780 |
| 2001 | | | 896 | | | 8,557 | 9,453 |
| 2002 | | | 639 | | | 10,575 | 11,215 |
| 2003 | | | 1,145 | | | 13,600 | 14,744 |
| 2004 | | | 731 | | | 11,165 | 11,896 |
| 2005 | | | 879 | | | 9,548 | 10,427 |
| 2006 | | | 1,041 | | | 11,826 | 12,867 |
| 2007 | | | 870 | | | 17,581 | 18,451 |
| 2008 | | | 513 | | | 16,923 | 17,436 |
| 2009 | | | 623 | | | 14,725 | 15,347 |
| 2010 | | | 3,547 | | | 14,304 | 17,852 |
| 2011 | | | 5,601 | | | 18,403 | 24,004 |
| 2012 | | | 5,589 | | | 18,554 | 24,143 |
| 2013 | | | 5,050 | | | 26,342 | 31,393 |
| 2014* | | | 1,842 | | | 24,047 | 25,889 |

*Estimated removals through October 11, 2014.

Table 4. Estimated retained and discarded catch (t), and percent discarded, of Pacific ocean perch from the eastern Bering Sea (EBS) and Aleutian Islands (AI) regions.

| | | EBS | | | AI | | | BSAI | |
|------|----------|-----------|-----------|----------|-----------|-----------|----------|---------|-----------|
| | | | Percent | | | Percent | | | Percent |
| Year | Retained | Discarded | Discarded | Retained | Discarded | Discarded | Retained | Discard | Discarded |
| 1990 | 5,069 | 1,275 | 20.1 | 10,288 | 1,551 | 13.1 | 15,357 | 2,826 | 15.54 |
| 199 | 1 4,126 | 972 | 19.07 | 1,815 | 970 | 34.82 | 5,941 | 1,942 | 24.63 |
| 1992 | 2 5,464 | 1044 | 16.05 | 17,332 | 3,227 | 15.7 | 22,797 | 4,271 | 15.78 |
| 1993 | 3 2,601 | 1163 | 30.9 | 11,479 | 1,896 | 14.18 | 14,080 | 3,059 | 17.85 |
| 1994 | 4 1,187 | 501 | 29.69 | 9,491 | 1,374 | 12.65 | 10,678 | 1,876 | 14.94 |
| 1995 | 5 839 | 368 | 30.49 | 8,603 | 1,701 | 16.51 | 9,442 | 2,069 | 17.97 |
| 1996 | 5 2,522 | 333 | 11.66 | 9,831 | 2,995 | 23.35 | 12,353 | 3,328 | 21.22 |
| 199 | 7 420 | 261 | 38.35 | 10,854 | 1,794 | 14.18 | 11,274 | 2,055 | 15.42 |
| 1998 | 821 | 200 | 19.62 | 8,282 | 1,017 | 10.93 | 9,103 | 1,217 | 11.79 |
| 1999 | 9 277 | 144 | 34.28 | 10,985 | 1,499 | 12.01 | 11,261 | 1,643 | 12.73 |
| 2000 | 230 | 221 | 49.01 | 8,586 | 743 | 7.96 | 8,816 | 964 | 9.85 |
| 200 | 1 399 | 497 | 55.45 | 7,195 | 1,362 | 15.92 | 7,594 | 1,859 | 19.66 |
| 2002 | 2 286 | 355 | 55.44 | 9,315 | 1,260 | 11.91 | 9,601 | 1,615 | 14.4 |
| 2003 | 3 549 | | | 10,720 | 2,042 | 16 | 11,269 | 2,668 | 19.14 |
| 2004 | 4 536 | | | 9,286 | 1,879 | 16.83 | 9,822 | 2,074 | 17.44 |
| 2003 | 5 627 | | | 8,100 | 1,448 | 15.16 | 8,727 | 1,700 | 16.31 |
| 2000 | 5 751 | 290 | | 9,869 | 1,957 | 16.55 | 10,620 | 2,246 | 17.46 |
| 200 | 7 508 | | | 15,051 | 2,530 | 14.39 | 15,558 | 2,893 | 15.68 |
| 2008 | 318 | | | 16,640 | 283 | 1.67 | 16,959 | 477 | 2.74 |
| 2009 | 9 463 | 160 | 25.67 | 14,011 | 713 | 4.84 | 14,474 | 873 | 5.69 |
| 2010 | 3347 | 200 | 5.64 | 13,988 | 316 | 2.21 | 17,335 | 516 | 2.89 |
| 201 | 1 5249 | 353 | 6.30 | 18,021 | 382 | 2.08 | 23,269 | 735 | 3.06 |
| 2012 | 5181 | 409 | | 18,162 | 392 | 2.12 | 23,343 | 801 | 3.32 |
| 2013 | 3 4746 | 304 | 6.03 | 26,094 | 249 | 0.94 | 30,840 | 553 | 1.76 |
| 2014 | * 1096 | 747 | 40.53 | 23,873 | 174 | 0.72 | 24,969 | 920 | 3.56 |

*Estimated removals through October 11, 2014. Source: NMFS Alaska Regional Office

Table 5. Percentage catch (by weight) of Aleutians Islands POP in the foreign/joint venture fisheries and the domestic fishery by depth.

| Depth Zone (m) | | | | | | | | | | | |
|----------------|------|----|-----|-----|-----|-----|-----|-----|-----------|--------|---------|
| | | | | | | | | | Observe I | | Percent |
| | | | | | | | | | d catch | total | sampled |
| _ | Year | 0 | 100 | 200 | 300 | 400 | 500 | 501 | (t) | catch | |
| | 1977 | 25 | 23 | 39 | 11 | 2 | 1 | 0 | 173 | 7,927 | 2 |
| | 1978 | 0 | 40 | 36 | 19 | 3 | 1 | 1 | 145 | 5,286 | 3 |
| | 1979 | 0 | 13 | 60 | 23 | 4 | 0 | 0 | 311 | 5,486 | 6 |
| | 1980 | 0 | 7 | 45 | 49 | 0 | 0 | 0 | 108 | 4,010 | 3 |
| | 1981 | 0 | 9 | 67 | 23 | 0 | 0 | 0 | 138 | 3,668 | 4 |
| | 1982 | 0 | 34 | 56 | 5 | 2 | 1 | 2 | 115 | 979 | 12 |
| | 1983 | 0 | 11 | 85 | 0 | 1 | 1 | 1 | 54 | 471 | 11 |
| | 1984 | 0 | 53 | 42 | 5 | 0 | 1 | 0 | 85 | 565 | 15 |
| | 1985 | 0 | 87 | 13 | 0 | 0 | 0 | 0 | 109 | 216 | 50 |
| | 1986 | 0 | 74 | 25 | 2 | 0 | 0 | 0 | 66 | 163 | 40 |
| | 1987 | 0 | 39 | 61 | 0 | 0 | 0 | 0 | 258 | 502 | 51 |
| | 1988 | 0 | 78 | 21 | 1 | 0 | 0 | 0 | 76 | 1,512 | 5 |
| | 1989 | | | | | | | | | | |
| | 1990 | 2 | 23 | 58 | 14 | 2 | 1 | 0 | 7,726 | 11,826 | 65 |
| | 1991 | 0 | 23 | 70 | 5 | 1 | 1 | 0 | 1,588 | 2,785 | 57 |
| | 1992 | 0 | 21 | 71 | 8 | 0 | 0 | 0 | 6,785 | 10,280 | 66 |
| | 1993 | 0 | 20 | 77 | 3 | 0 | 0 | 0 | 8,867 | 13,375 | 66 |
| | 1994 | 0 | 20 | 69 | 11 | 0 | 0 | 0 | 7,562 | 10,866 | 70 |
| | 1995 | 0 | 15 | 68 | 14 | 2 | 0 | 0 | 6,154 | 10,303 | 60 |
| | 1996 | 0 | 17 | 54 | 26 | 2 | 1 | 0 | 8,547 | 12,827 | 67 |
| | 1997 | 0 | 13 | 66 | 21 | 0 | 0 | 0 | 9,320 | 12,648 | 74 |
| | 1998 | 0 | 21 | 72 | 7 | 0 | 0 | 0 | 7,380 | 9,299 | 79 |
| | 1999 | 0 | 30 | 63 | 7 | 0 | 0 | 0 | 10,369 | 12,483 | 83 |
| | 2000 | 0 | 21 | 63 | 15 | 0 | 0 | 0 | 7,456 | 9,328 | 80 |
| | 2001 | 0 | 29 | 61 | 10 | 0 | 0 | 0 | 5,679 | 8,557 | 66 |
| | 2002 | 2 | 36 | 57 | 5 | 1 | 0 | 0 | 8,124 | 10,575 | 77 |
| | 2003 | 0 | 26 | 70 | 3 | 0 | 0 | 0 | 11,266 | 13,600 | 83 |
| | 2004 | 1 | 26 | 65 | 7 | 1 | 0 | 0 | 10,083 | 11,165 | 90 |
| | 2005 | 2 | 36 | 55 | 6 | 1 | 0 | 0 | 7,403 | 9,548 | 78 |
| | 2006 | 1 | 33 | 61 | 5 | 0 | 0 | 0 | 9,895 | 11,826 | 84 |
| | 2007 | 0 | 23 | 68 | 7 | 1 | 0 | 0 | 15,551 | 17,581 | 88 |
| | 2008 | 1 | 20 | 74 | 5 | 0 | 0 | 0 | 16,682 | 16,923 | 99 |
| | 2009 | 1 | 26 | 65 | 8 | 1 | 0 | 1 | 14,495 | 14,724 | 98 |
| | 2010 | 1 | 21 | 71 | 7 | 1 | 0 | 0 | 14,299 | 14,304 | 100 |
| | 2011 | 0 | 13 | 78 | 7 | 1 | 0 | 0 | 18,391 | 18,403 | 100 |
| | 2012 | 0 | 21 | 67 | 11 | 1 | 0 | 0 | 18,553 | 18,554 | 100 |
| | 2013 | 0 | 12 | 76 | 11 | 1 | 0 | 0 | 26,297 | 26,342 | 100 |

Table 6. Proportional catch (by weight) of Aleutians Islands POP in the foreign and joint venture fisheries and the domestic fishery by management area.

| | | Area | | | | |
|------|-----|------|-----|--------------------|--------|---------|
| | | | | Observed Estimated | | Percent |
| | | | | catch (t) | total | sampled |
| | 541 | 542 | 543 | | catch | |
| 1977 | 17 | 22 | 61 | 173 | 7,927 | 2 |
| 1978 | 30 | 36 | 35 | 145 | 5,286 | 3 |
| 1979 | 21 | 25 | 55 | 311 | 5,486 | 6 |
| 1980 | 11 | 42 | 47 | 108 | 4,010 | 3 |
| 1981 | 42 | 40 | 17 | 138 | 3,668 | 4 |
| 1982 | 42 | 38 | 20 | 115 | 979 | 12 |
| 1983 | 85 | 8 | 7 | 54 | 471 | 11 |
| 1984 | 84 | 8 | 7 | 85 | 565 | 15 |
| 1985 | 66 | 34 | 0 | 109 | 216 | 50 |
| 1986 | 99 | 1 | 0 | 66 | 163 | 40 |
| 1987 | 94 | 6 | 0 | 258 | 502 | 51 |
| 1988 | 6 | 94 | 0 | 76 | 1,512 | 5 |
| 1989 | | | | | | |
| 1990 | 63 | 16 | 21 | 7,726 | 11,826 | 65 |
| 1991 | 27 | 57 | 16 | 1,588 | 2,785 | 57 |
| 1992 | 81 | 15 | 3 | 6,785 | 10,280 | 66 |
| 1993 | 67 | 22 | 11 | 8,867 | 13,375 | 66 |
| 1994 | 64 | 31 | 5 | 7,562 | 10,866 | 70 |
| 1995 | 70 | 25 | 5 | 6,154 | 10,303 | 60 |
| 1996 | 27 | 20 | 54 | 8,547 | 12,827 | 67 |
| 1997 | 20 | 23 | 57 | 9,320 | 12,648 | 74 |
| 1998 | 21 | 27 | 52 | 7,380 | 9,299 | 79 |
| 1999 | 22 | 23 | 56 | 10,369 | 12,483 | 83 |
| 2000 | 22 | 24 | 54 | 7,456 | 9,328 | 80 |
| 2001 | 27 | 25 | 48 | 5,679 | 8,557 | 66 |
| 2002 | 24 | 28 | 48 | 8,124 | 10,575 | 77 |
| 2003 | 30 | 22 | 48 | 11,266 | 13,600 | 83 |
| 2004 | 24 | 27 | 49 | 10,083 | 11,165 | 90 |
| 2005 | 23 | 24 | 52 | 7,403 | 9,548 | 78 |
| 2006 | 24 | 28 | 48 | 9,895 | 11,826 | 84 |
| 2007 | 30 | 26 | 45 | 15,551 | 17,581 | 88 |
| 2008 | 28 | 28 | 44 | 16,682 | 16,923 | 99 |
| 2009 | 27 | 28 | 44 | 14,495 | 14,724 | 98 |
| 2010 | 28 | 28 | 44 | 14,299 | 14,304 | 100 |
| 2011 | 30 | 26 | 44 | 18,391 | 18,403 | 100 |
| 2012 | 30 | 26 | 44 | 18,553 | 18,554 | 100 |
| 2013 | 36 | 26 | 38 | 26,297 | 26,342 | 100 |

Table 7. Length measurements from the EBS and AI POP fisheries during 1964-1972, from Chikuni (1975)

| Year | EBS | AI | Total |
|------|--------|--------|---------|
| 1964 | 24,150 | 55,599 | 79,749 |
| 1965 | 14,935 | 66,120 | 81,055 |
| 1966 | 26,458 | 25,502 | 51,960 |
| 1967 | 48,027 | 59,576 | 107,603 |
| 1968 | 38,370 | 36,734 | 75,104 |
| 1969 | 28,774 | 27,206 | 55,980 |
| 1970 | 11,299 | 27,508 | 38,807 |
| 1971 | 14,045 | 18,926 | 32,971 |
| 1972 | 10,996 | 18,926 | 29,922 |

Table 8. Length measurements and otoliths read from the EBS and AI POP fisheries, from the NORPAC Observer database.

| | Fish | | | Otoliths read | | |
|--------------|------------|----------------|---------|-----------------|-----|-------|
| | lengths | | | O tolkilis read | | |
| Year | EBS | AI | Total | EBS | AI | Total |
| 1973 | 1 | | 1** | | | |
| 1974 | 84 | | 84** | 84 | | 84** |
| 1975 | 271 | | 271** | 125 | | 125** |
| 1976 | 633 | | 633** | 114 | 19 | 133** |
| 1977 | 1,059 | 9,318 | | 139 | 404 | 543 |
| 1978 | 7,926 | 7,283 | 10,377* | 583 | 641 | 1,224 |
| | 1,045 | | 15,209* | | | |
| 1979 | 1,043 | 10,921 | 11,966* | 248 | 353 | 601 |
| 1980 | | 3,995 | 3,995* | | 398 | 398 |
| 1981 | 1,502 | 7,167 | 8,669* | 78 | 432 | 510 |
| 1982 | | 4,902 | 4,902* | | 222 | 222 |
| 1983 | 232 | 441 | 673 | | | |
| 1984 | 1,194 | 1,210 | 2,404 | 72 | | 72** |
| 1985 | 300 | | 300** | 160 | | 160** |
| 1986 | | 100 | 100** | | 99 | 99** |
| 1987 | 11 | 384 | 395 | | | |
| 1988 | 306 | 1,366 | 1,672 | | | |
| 1989 | 957 | 91 | 1,048 | | | |
| 1990 | 22,228 | 47,198 | 69,426 | 144 | 184 | 328 |
| 1991 | 8,247 | 8,221 | 16,468 | | | |
| 1992 | 13,077 | 24,932 | 38,009 | | | |
| 1993 | 8,379 | 26,433 | 34,812 | | | |
| 1994 | 2,654 | 11,546 | 14,200 | | | |
| 1995 | 272 | 11,452 | 11,724 | | | |
| 1996 | 2,967 | 13,146 | 16,113 | | | |
| 1997 | 143 | 10,402 | 10,545 | | 000 | 022 |
| 1998 | 989 | 11,106 | 12,095 | | 823 | 823 |
| 1999 2000 | 289 284 | 3,839 3,382 | 4,128 | | 487 | 487 |
| | | | 3,666* | | | |
| 2001 | 327 | 2,388 | 2,715* | 11 | 524 | 524 |
| 2002 | 78 | 3,671 | 3,749* | 11 | 455 | 466 |
| 2003 | 247 | 4,681 | 4,928* | 11 | 386 | 397 |
| 2004 | 135 | 3,270 | 3,405* | 30 | 754 | 784 |
| 2005 | 237 | 2,243 | 2,480* | 42 | 539 | 581 |
| 2006 | 274 | 3,757 | 4,031* | 25 | 424 | 449 |
| 2007 | 74 | 5,629 | 5,703* | 11 | 664 | 675 |
| 2008 | 250 | 7,001 | 7,251* | 17 | 555 | 572 |
| 2009 | 460 | 5,593 | 6,053* | 49 | 670 | 719 |
| 2010 | 2,584 | 5,384 | 7,968 | | | |
| 2011 | 4,144 | 7,965 | 12,109* | 316 | 616 | 932 |
| 2012 | 5,686 | 7,896 | 13,582 | | | |
| 2013 | 3,897 | 13,082 | 16,979* | 233 | 810 | 1043 |
| 2014 | 754 | 7,910 | 8,664 | | | |
| | | / | , ' | 44 | | |

^{*}Used to create age composition. **Not used.

Table 9. Pacific ocean perch estimated biomass (t) from the eastern Bering Sea slope and Aleutian Islands trawl surveys (by management area).

| | South | ern Bering S | Sea | Ale | utian Islands | S | Total Aleut | ian Islands S | Survey | EBS | Slope surv | ey |
|--------------|--------|--------------|-------|------------|---------------|--------|-------------|---------------|--------|-----------------|------------|------|
| Year | Mean | SD | CV | Mean | SD | CV | Mean | SD | CV | Mean | SD | CV |
| 1979 | | | | | | | | | | | | |
| 1980 | 5,833 | 5,658 | 97% | 76,545 | 45,686 | 60% | 82,378 | 46,035 | 56% | | | |
| 1981 | | | | | | | | | | | | |
| 1982 | | | | | | | | | | | | |
| 1983 | 90,622 | 72,317 | 80% | 141,261 | 37,075 | 26% | 231,883 | 81,267 | 35% | | | |
| 1984 | | | | | | | | | | | | |
| 1985 | | | | | | | | | | | | |
| 1986 | 26,784 | 13,031 | 49% | 197,656 | 42,463 | 21% | 224,440 | 44,418 | 20% | | | |
| 1987 | | | | | | | | | | | | |
| 1988 | | | | | | | | | | | | |
| 1989 | | | | | | | | | | | | |
| 1990 | | | | | | | | | | | | |
| 1991 | 1,501 | 758 | 51% | 342,785 | 70,773 | 21% | 344,286 | 70,777 | 21% | | | |
| 1992 | | | | | | | | | | | | |
| 1993 | | | | | | | | | | | | |
| 1994 | 18,217 | 11,685 | 64% | 369,699 | 88,327 | 24% | 387,916 | 89,096 | 23% | | | |
| 1995 | | | | | | | | | | | | |
| 1996 | 42.000 | = 000 | #00¢ | # < # OO # | 0.4.50.4 | 4.50/ | | 04.044 | 4.50/ | | | |
| 1997 | 12,099 | 7,008 | 58% | 565,885 | 84,524 | 15% | 577,984 | 84,814 | 15% | | | |
| 1998 1999 | | | | | | | | | | | | |
| 2000 | 10 070 | 10,150 | 54% | 500,118 | 91.099 | 18% | 518,988 | 91,662 | 18% | | | |
| 2000 | 18,870 | 10,130 | 34% | 300,118 | 91,099 | 18% | 310,900 | 91,002 | 18% | | | |
| 2001 | 16,311 | 6,637 | 41% | 446,860 | 77,841 | 17% | 463,171 | 78,123 | 17% | 72,665 | 38,586 | 53% |
| 2002 | 10,511 | 0,037 | 41 /0 | 440,800 | 77,041 | 1 / /0 | 403,171 | 70,123 | 1 / /0 | 72,003 | 30,300 | 3370 |
| 2003 | 74,208 | 33,397 | 45% | 503,228 | 64,592 | 13% | 577,436 | 72,715 | 13% | 112,273 | 42,681 | 38% |
| 2005 | 74,200 | 33,371 | 7370 | 303,220 | 04,372 | 1370 | 377,430 | 72,713 | 1370 | 112,273 | 42,001 | 3070 |
| 2006 | 23,701 | 11,194 | 47% | 623,549 | 90,482 | 15% | 647,250 | 91,172 | 14% | | | |
| 2007 | 25,701 | 11,17 | .,,0 | 020,019 | 70,.02 | 1570 | 017,250 | 71,172 | 11,0 | | | |
| 2008 | | | | | | | | | | 107,886 | 43,711 | 41% |
| 2009 | | | | | | | | | | , | - /- | |
| 2010 | 87,794 | 47,952 | 55% | 884,241 | 104,840 | 12% | 972,035 | 115,286 | 12% | 203,421 | 78,235 | 38% |
| 2011 | | , | | • | | | | | | , in the second | | |
| 2012 | 38,658 | 24,190 | 63% | 863,741 | 153,111 | 18% | 902,398 | 155,010 | 17% | 231,383 | 75,235 | 33% |
| 2013 | - | • | | • | • | | | • | | • | • | |
| 2014 | 83,409 | 41,568 | 50% | 887,559 | 180,704 | 20% | 970,968 | 185,423 | 19% | | | |

Table 10. Pacific ocean perch biomass estimates (t) from the 1991-2014 triennial trawl surveys for the three management sub-areas in the Aleutian Islands region.

| | Aleutian Islands Management | | | |
|------------------|-----------------------------|---------|---------|--|
| | Sub-Areas | | | |
| Year | Western | Central | Eastern | |
| 1991 | 208,465 | 78,776 | 55,545 | |
| 1994 | 184,703 | 84,411 | 100,585 | |
| 1997 | 178,437 | 166,816 | 220,633 | |
| 2000 | 229,850 | 129,740 | 140,528 | |
| 2002 | 196,704 | 140,361 | 109,795 | |
| 2004 | 212,639 | 153,477 | 137,112 | |
| 2006 | 281,946 | 150,851 | 190,752 | |
| 2010 | 395,944 | 221,700 | 266,607 | |
| 2012 | 263,661 | 233,666 | 366,413 | |
| 2014 | 338,455 | 315,544 | 233,560 | |
| Weighted Average | | | | |
| (2010-2014) | 326,939 | 269,931 | 282,471 | |
| Percentage | 37.18% | 30.70% | 32.12% | |

Table 11. Length measurements and otoliths read from the Aleutian Islands surveys.

| Year | Length measurements | Otoliths read |
|------|---------------------|---------------|
| 1980 | 20,796 | 890 |
| 1983 | 22,873 | 2,495 |
| 1986 | 14,804 | 1,860 |
| 1991 | 14,262 | 1,015 |
| 1994 | 18,922 | 849 |
| 1997 | 22,823 | 1,224 |
| 2000 | 21,972 | 1,238 |
| 2002 | 20,284 | 337 |
| 2004 | 24,949 | 1,031 |
| 2006 | 19,737 | 462 |
| 2010 | 22,725 | 951 |
| 2012 | 31,450 | 1,140 |
| 2014 | 30,206 | |

Table 12. Negative log likelihoods, and several measures of model fits, for the evaluated models for BSAI POP.

| | Model 0 | Model 0.1 | Model 1 | Model 2 | Model 3 |
|--|---------|-----------|---------|---------|---------|
| Negative log-likelihood | | | | | |
| Data components | | | | | |
| AI survey biomass | 9.97 | 7.67 | 9.49 | 20.92 | 8.52 |
| CPUE | 23.74 | 24.89 | 33.59 | 22.16 | 26.28 |
| Catch biomass | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| Fishery age comp | 41.38 | 43.80 | 209.27 | 225.20 | 226.03 |
| Fishery length comp | 326.17 | 313.63 | 416.90 | 467.54 | 358.98 |
| AI survey age comp | 120.78 | 52.96 | 261.75 | 177.14 | 150.23 |
| AI survey lengths com | p 6.74 | 6.48 | 13.32 | 23.27 | 10.54 |
| Maturity | 2.71 | 2.71 | 2.71 | 2.71 | 2.71 |
| Priors and penalties | | | | | |
| Recruitment | 8.29 | 6.98 | 10.48 | 13.40 | 11.95 |
| Prior on survey q | 14.69 | 14.38 | 16.46 | 0.00 | 9.89 |
| Prior on M | 0.19 | 0.21 | 0.01 | 1.30 | 0.30 |
| Fishery selectivity | 14.32 | 13.51 | 23.48 | 67.66 | 142.86 |
| Total negative log-likelihood | 565.00 | 483.02 | 988.71 | 1018.15 | 942.41 |
| Parameters | 121 | | 145 | 121 | 137 |
| Number of data points | | | 1024 | 1024 | 1024 |
| BIC | | | 2982.49 | 2875.01 | 2834.43 |
| AIC | | | 2267.43 | 2278.30 | 2158.81 |
| Effective sample size | | | | | |
| Fishery age comp | 266 | 266 | 324 | 269 | 199 |
| Fishery length comp | 189 | 215 | 144 | 138 | 164 |
| AI survey age comp | 101 | 120 | 158 | 173 | 210 |
| AI survey lengths com | p 88 | 93 | 41 | 24 | 52 |
| Sample weights | | | | | |
| Fishery age comp | 24 | 24 | 153 | 153 | 153 |
| Fishery length comp | 147 | 147 | 168 | 168 | 168 |
| AI survey age comp | 33 | 30 | 224 | 224 | 224 |
| AI survey lengths com | p 174 | 174 | 175 | 175 | 175 |
| Root mean square error | | | | | |
| AI survey biomass | 0.245 | 0.198 | 0.227 | 0.351 | 0.222 |
| CPUE | 0.778 | | | 0.761 | 0.804 |
| Recruitment | 0.763 | | | 0.832 | 0.813 |
| Fishery age comp | 0.013 | | | | |
| Fishery length comp | 0.025 | | | | |
| AI survey age comp | 0.020 | | | | |
| AI survey lengths com | | | | | |
| Standard Deviation of Normalized Residuals | | | | | |
| AI survey biomass | 1.24 | | 1.38 | 2.04 | 1.30 |
| CPUE | 2.59 | 2.63 | 2.92 | | |
| Fishery age comp | 0.47 | | | 1.05 | 1.00 |
| Fishery length comp | 1.05 | 1.03 | 1.16 | 1.20 | 1.00 |
| AI survey age comp | 1.42 | | | | |
| AI survey lengths com | p 0.80 | 0.78 | 1.14 | 1.51 | 1.00 |

Table 13. Estimated time series of POP total biomass (t), spawning biomass (t), and recruitment (thousands) for each region.

| Total B | iomass | Spawning | Spawning Biomass | | | |
|--------------|---------|----------|------------------|-----------------|---------|--|
| (ages | (ages | | Recruitmen | (0) | | |
| Assessme | | Assessme | | Assessment Year | | |
| Year 2014 | 2012 | 2014 | 2012 | 2014 | 2012 | |
| 1977 238,630 | 124,658 | 99,038 | 40,191 | 25,581 | 26,044 | |
| 1978 226,774 | 122,886 | 93,438 | 39,443 | 23,814 | 35,844 | |
| 1979 221,931 | 128,670 | 89,477 | 39,942 | 65,542 | 84,341 | |
| 1980 218,921 | 137,002 | 86,494 | 41,018 | 57,894 | 85,676 | |
| 1981 218,423 | 152,553 | 84,678 | 42,955 | 38,725 | 113,845 | |
| 1982 221,874 | 165,729 | 83,672 | 45,547 | 72,287 | 52,563 | |
| 1983 231,810 | 185,825 | 84,372 | 49,941 | 78,598 | 80,592 | |
| 1984 250,389 | 216,581 | 86,102 | 55,711 | 145,384 | 176,053 | |
| 1985 267,617 | 241,708 | 89,001 | 63,190 | 76,490 | 67,223 | |
| 1986 285,047 | 267,573 | 93,376 | 72,676 | 51,875 | 67,491 | |
| 1987 324,368 | 316,811 | 98,946 | 83,575 | 311,366 | 329,776 | |
| 1988 349,846 | 356,870 | 105,631 | 95,212 | 45,770 | 140,299 | |
| 1989 381,769 | 396,998 | 113,713 | 107,429 | 140,206 | 131,150 | |
| 1990 406,537 | 431,007 | 121,809 | 118,965 | 59,069 | 84,057 | |
| 1991 427,526 | 464,005 | 128,713 | 128,974 | 192,205 | 236,082 | |
| 1992 454,119 | 500,992 | 139,920 | 143,481 | 92,298 | 123,119 | |
| 1993 470,393 | 525,969 | 150,480 | 157,834 | 52,015 | 68,419 | |
| 1994 477,652 | 541,359 | 160,910 | 172,697 | 26,405 | 46,977 | |
| 1995 486,685 | 557,206 | 172,343 | 188,861 | 43,833 | 51,764 | |
| 1996 492,300 | 569,967 | 182,568 | 203,668 | 33,146 | 56,337 | |
| 1997 498,090 | 581,483 | 190,338 | 215,652 | 127,572 | 133,733 | |
| 1998 503,929 | 594,244 | 197,447 | 226,765 | 92,073 | 125,087 | |
| 1999 522,414 | 617,922 | 202,870 | 235,761 | 201,659 | 209,310 | |
| 2000 529,217 | 630,674 | 205,025 | 240,999 | 48,880 | 84,719 | |
| 2001 549,735 | 654,299 | 206,720 | 245,314 | 188,650 | 181,061 | |
| 2002 563,361 | 673,980 | 208,269 | 248,953 | 59,844 | 113,471 | |
| 2003 597,042 | 702,078 | 210,198 | 252,743 | 331,236 | 228,612 | |
| 2004 611,396 | 713,418 | 213,591 | 257,607 | 38,212 | 48,243 | |
| 2005 629,571 | 723,660 | 220,422 | 265,526 | 81,642 | 43,243 | |
| 2006 643,005 | 730,659 | 229,350 | 274,802 | 29,979 | 40,220 | |
| 2007 654,833 | 731,429 | 238,001 | 282,981 | 94,791 | 48,897 | |
| 2008 657,981 | 722,295 | 245,839 | 288,902 | 78,861 | 46,016 | |
| 2009 656,210 | 710,471 | 254,326 | 293,817 | 36,897 | 43,792 | |
| 2010 653,665 | 702,831 | 261,815 | 296,276 | 45,917 | | |
| 2011 646,482 | 692,564 | 264,704 | 293,527 | 59,397 | | |
| 2012 633,007 | 676,409 | 262,190 | 285,289 | | | |
| 2013 618,851 | 661,440 | 256,200 | 273,683 | | | |
| 2014 597,506 | | 246,104 | | | | |
| 2015 577,967 | | | | | | |

Table 14. Projections of BSAI spawning biomass (t), catch (t), and fishing mortality rate for each of the several scenarios. The values of $B_{40\%}$ and $B_{35\%}$ are 169,203 t and 148,053 t, respectively.

| Catch | Scenario 1 | Scenario 2 | Scenario 3 | Scenario 4 | Scenario 5 | Scenario 6 | Scenario 7 |
|---------|------------|------------|------------|------------|------------|------------|------------|
| | | | | | | | |
| 2014 | 31,162 | 31,162 | 31,162 | 31,162 | 31,162 | 31,162 | 31,162 |
| 2015 | 34,988 | 34,988 | 17,819 | | 0 | 42,558 | 34,988 |
| 2016 | 33,194 | 33,194 | 17,534 | 12,280 | 0 | 39,716 | 33,194 |
| 2017 | 31,413 | 31,413 | 17,187 | 12,170 | 0 | 36,996 | 38,210 |
| 2018 | 29,768 | 29,768 | 16,838 | 12,049 | 0 | 34,544 | 35,625 |
| 2019 | 28,299 | 28,299 | 16,507 | 11,929 | 0 | 32,398 | 33,354 |
| 2020 | 27,088 | 27,088 | 16,247 | 11,848 | 0 | 30,554 | 31,484 |
| 2021 | 26,131 | 26,131 | 16,064 | | 0 | 28,391 | 29,574 |
| 2022 | 25,348 | , | 15,976 | | 0 | 26,840 | 27,831 |
| 2023 | 24,634 | | 15,966 | | 0 | | |
| 2024 | 24,133 | 24,133 | 16,020 | | 0 | | 25,887 |
| 2025 | 23,837 | | 16,123 | | 0 | | |
| 2026 | 23,676 | | 16,247 | | 0 | 24,792 | 25,184 |
| 2027 | 23,605 | 23,605 | 16,381 | 12,458 | 0 | | 25,054 |
| Sp. | | Scenario 2 | Scenario 3 | Scenario 4 | Scenario 5 | Scenario 6 | Scenario 7 |
| Biomass | | | | | | | |
| 2014 | 246,105 | | 246,105 | | 246,105 | | 246,105 |
| 2015 | 234,426 | | 236,594 | | 238,782 | , | 234,426 |
| 2016 | 221,197 | | 231,347 | | 241,995 | | 221,197 |
| 2017 | 208,987 | | 226,189 | | 244,930 | , | 208,118 |
| 2018 | 198,647 | | 222,008 | | 248,429 | | 194,699 |
| 2019 | 190,249 | | 218,933 | | 252,576 | | 183,739 |
| 2020 | 183,936 | | 217,242 | | 257,706 | | 175,298 |
| 2021 | 179,344 | | 216,654 | | 263,543 | | 168,988 |
| 2022 | 176,272 | | 217,096 | | 270,096 | | |
| 2023 | 174,328 | | 218,199 | | 277,021 | | 161,943 |
| 2024 | 173,230 | | 219,732 | | 284,150 | | 160,261 |
| 2025 | 172,627 | | 221,405 | | 291,195 | | 159,205 |
| 2026 | 172,341 | 172,341 | 223,123 | | 298,094 | | 158,553 |
| 2027 | 172,283 | | 224,857 | 246,386 | 304,831 | 157,113 | 158,192 |
| F | Scenario 1 | Scenario 2 | Scenario 3 | Scenario 4 | Scenario 5 | Scenario 6 | Scenario 7 |
| 2014 | 0.076 | 0.076 | 0.076 | 0.076 | 0.076 | 0.076 | 0.076 |
| 2015 | 0.089 | 0.089 | 0.045 | 0.031 | 0 | 0.109 | 0.089 |
| 2016 | 0.089 | 0.089 | 0.045 | 0.031 | 0 | 0.109 | 0.089 |
| 2017 | 0.089 | 0.089 | 0.045 | 0.031 | 0 | 0.109 | 0.109 |
| 2018 | 0.089 | 0.089 | 0.045 | 0.031 | 0 | 0.109 | 0.109 |
| 2019 | 0.089 | 0.089 | 0.045 | 0.031 | 0 | 0.109 | 0.109 |
| 2020 | 0.089 | 0.089 | 0.045 | 0.031 | 0 | 0.109 | 0.109 |
| 2021 | 0.089 | 0.089 | 0.045 | 0.031 | 0 | 0.106 | 0.108 |
| 2022 | 0.089 | 0.089 | 0.045 | 0.031 | 0 | 0.103 | 0.105 |
| 2023 | 0.088 | 0.088 | 0.045 | 0.031 | 0 | 0.102 | 0.103 |
| 2024 | 0.087 | 0.087 | 0.045 | 0.031 | 0 | 0.100 | 0.102 |
| 2025 | 0.086 | 0.086 | 0.045 | 0.031 | 0 | 0.100 | 0.101 |
| 2026 | 0.086 | 0.086 | 0.045 | 0.031 | 0 | 0.099 | 0.100 |
| 2027 | 0.086 | 0.086 | 0.045 | 0.031 | 0 | 0.099 | 0.100 |

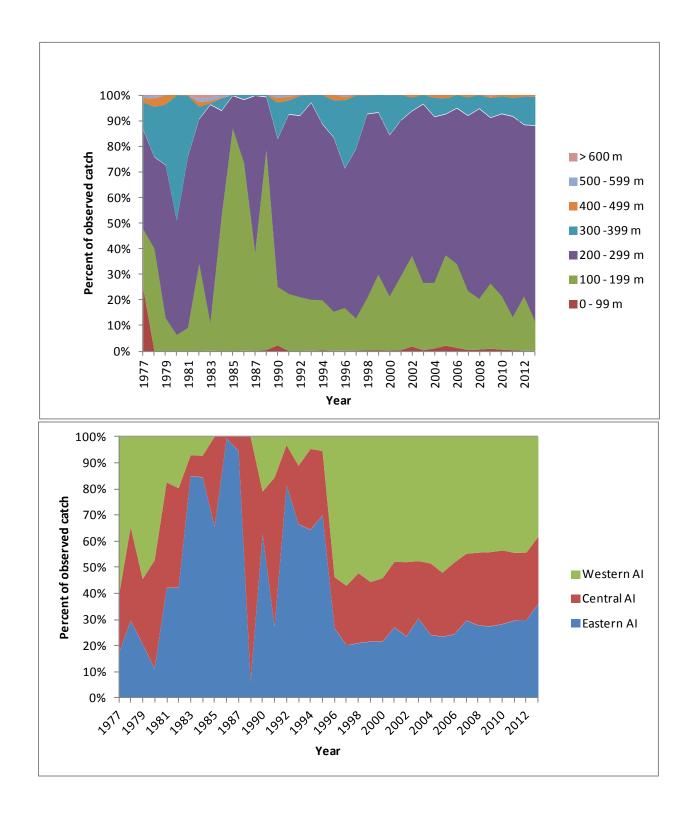


Figure 1. Distribution of observed Aleutian Islands Pacific ocean perch catch (from North Pacific Groundfish Observer Program) by depth zone (top panel) and AI subarea (bottom panel) from 1977 to 2013.

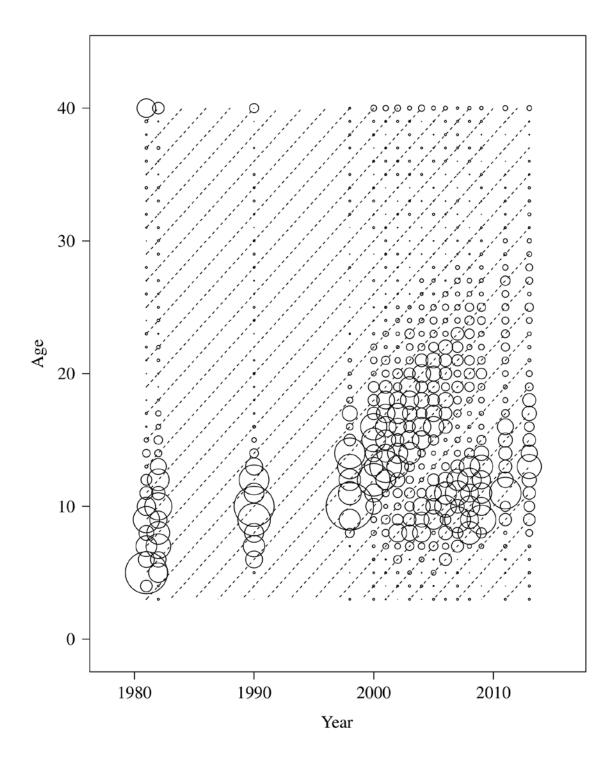


Figure 2. Fishery age composition data for the BSAI POP; The diameter of the circles are scaled within each year of samples, and dashed lines denote cohorts.

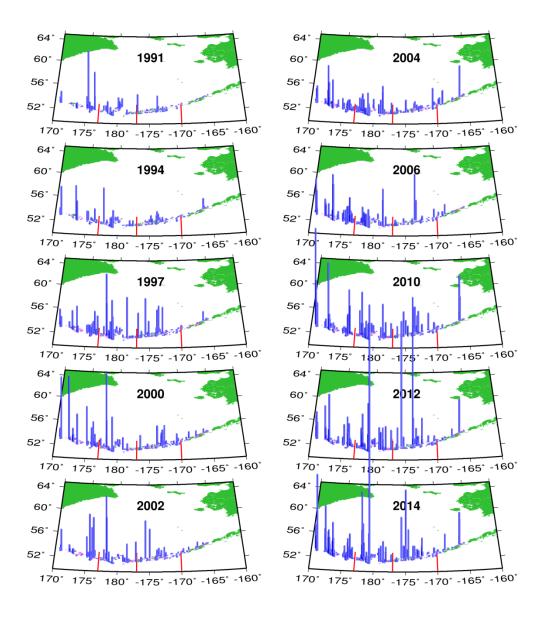


Figure 3. AI survey POP CPUE (kg/km^2) from 1991-2014; the symbol \times denotes tows with no catch. The red lines indicate boundaries between the WAI, CAI, EAI, and EBS areas.

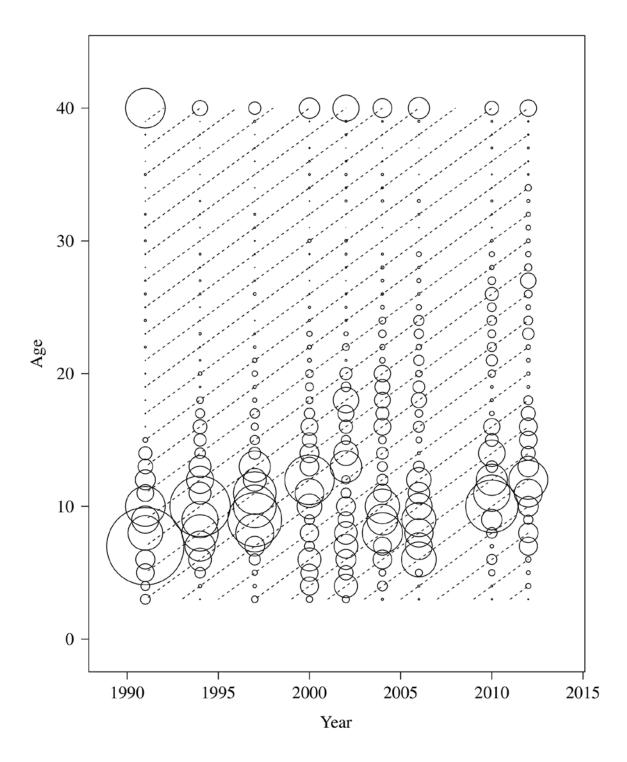


Figure 4. Age composition data from the Aleutian Islands trawl survey; bubbles are scaled within each year of samples; and dashed lines denote cohorts.

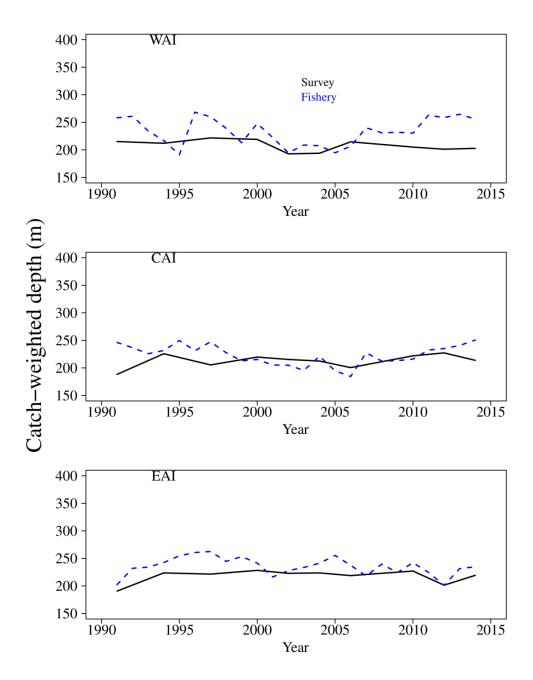


Figure 5. Catch-weighted (by numbers) depth of capture for Pacific ocean perch in the fishery and AI survey by AI subarea from 1991 to 2014.

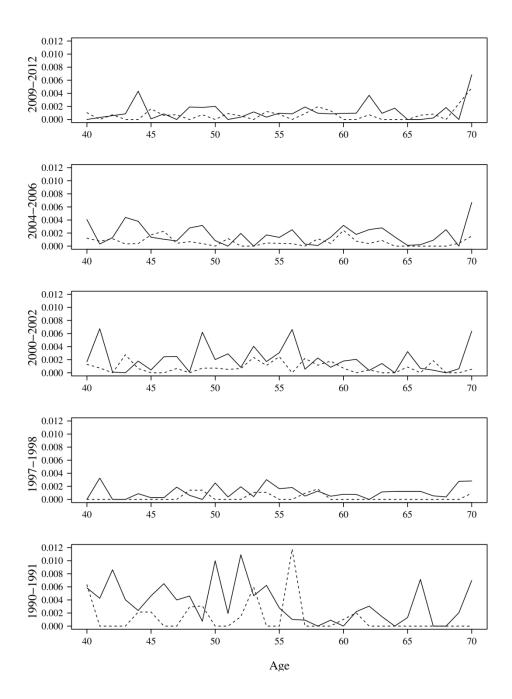


Figure 6. Age compositions in the Aleutian Islands survey (solid line) and fishery (dashed line) for ages 40 to 70+ for five time periods.

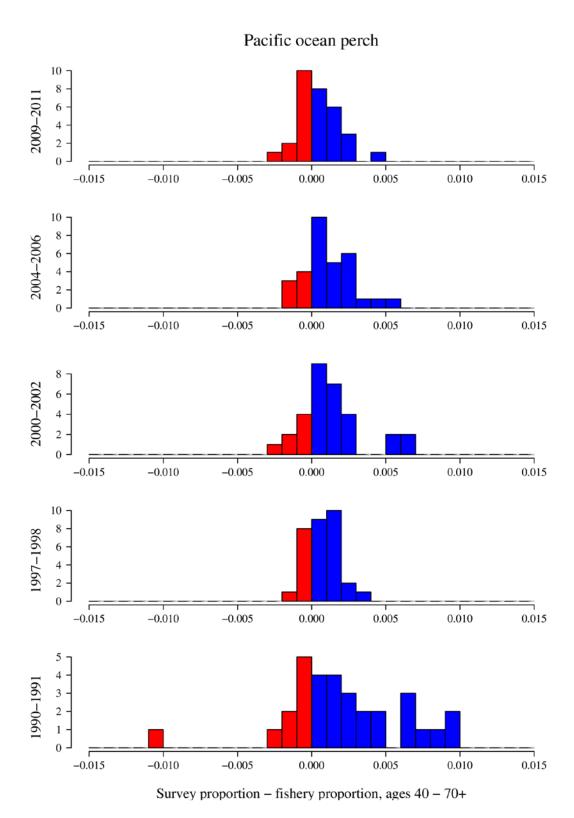


Figure 7. Histograms of the difference (survey proportion – fishery proportion) for ages 40 to 70+ for five time periods.

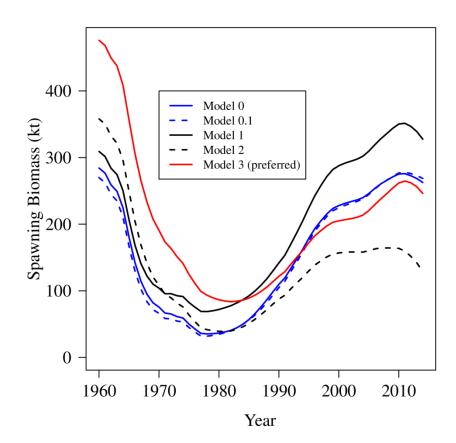


Figure 8. Estimated time series of spawning stock biomass across the models.

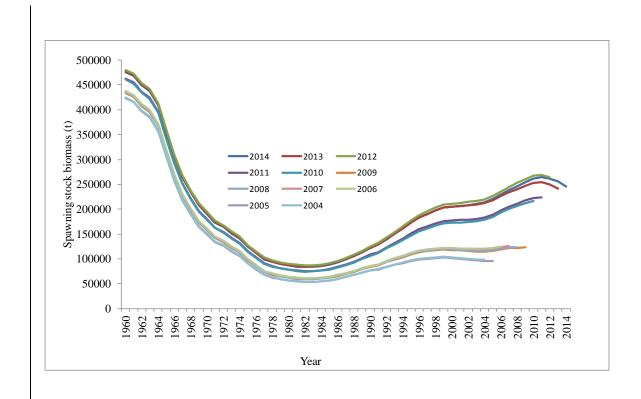


Figure 9. Retrospective estimates of spawning stock biomass for model runs with end years of 2004 to 2014.

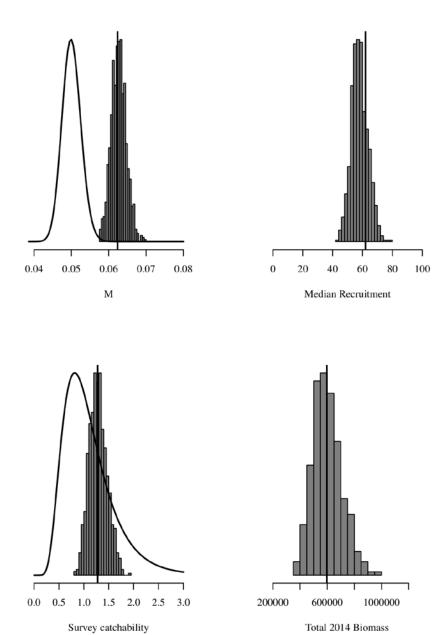


Figure 10. Posterior distributions for key model quantities M, survey catchability, median recruitment, and 2014 total biomass. For M and survey catchability, the prior distributions are also shown in the solid lines. The MLE estimates are indicated by the vertical lines.

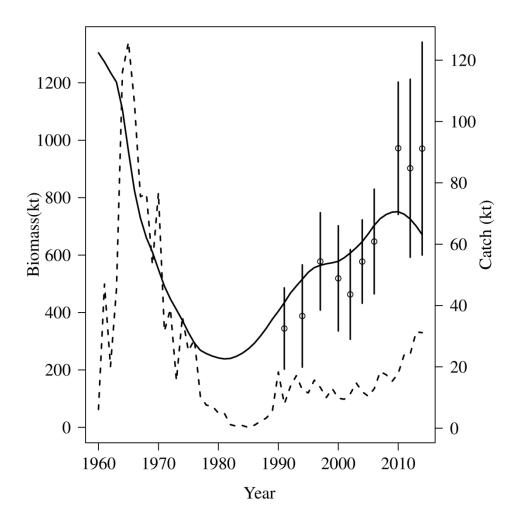


Figure 11. Observed AI survey biomass (data points, +/- 2 standard deviations), predicted survey biomass(solid line), and BSAI harvest (dashed line).

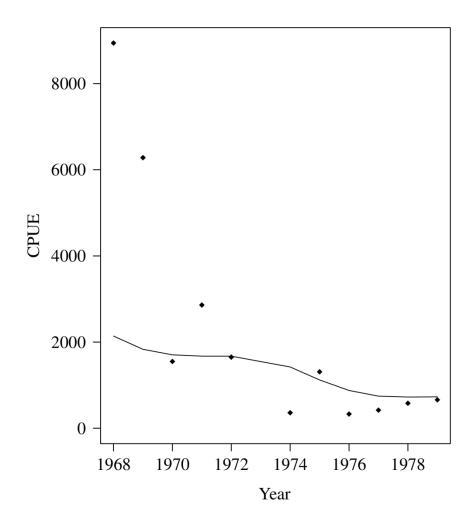


Figure 12. Observed AI CPUE (data points) and predicted CPUE (solid line) for BSAI POP.

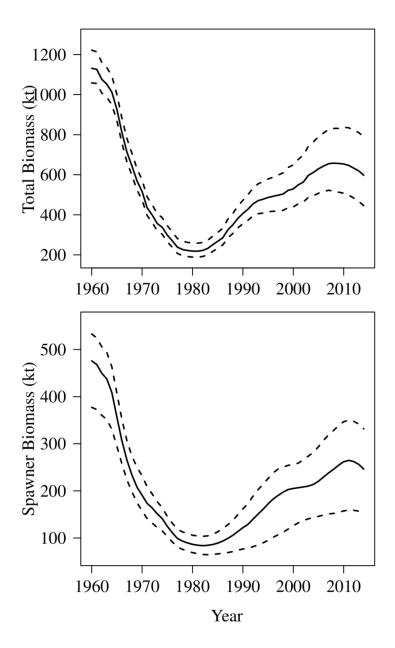


Figure 13. Total and spawner biomass for BSAI Pacific ocean perch, with 95% confidence intervals from MCMC integration.

Proportion

Figure 14. Model fits (dots) to fishery age composition data (columns) for Aleutian Islands Pacific ocean perch, 1981-2013. Colors correspond to cohorts (except for the 40+ group).

Fishery length composition data

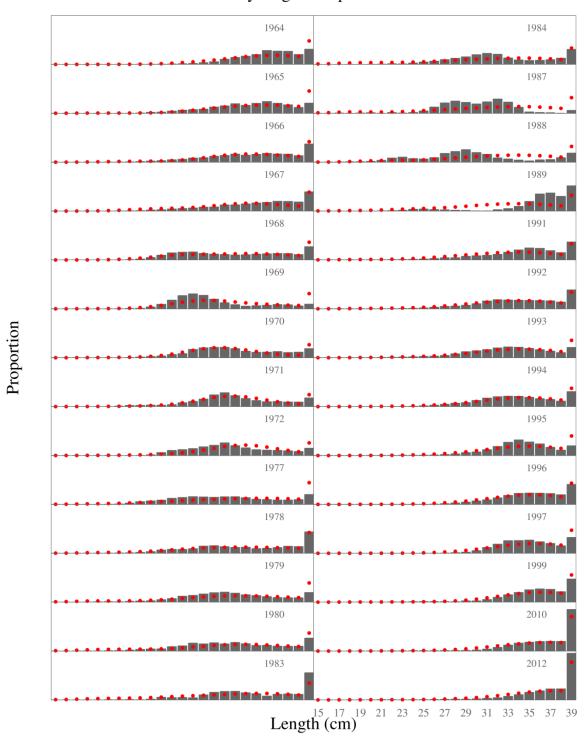


Figure 15. Model fits (dots) to fishery length composition data (columns) for Aleutian Islands Pacific ocean perch, 1964-2012.

Survey age composition data

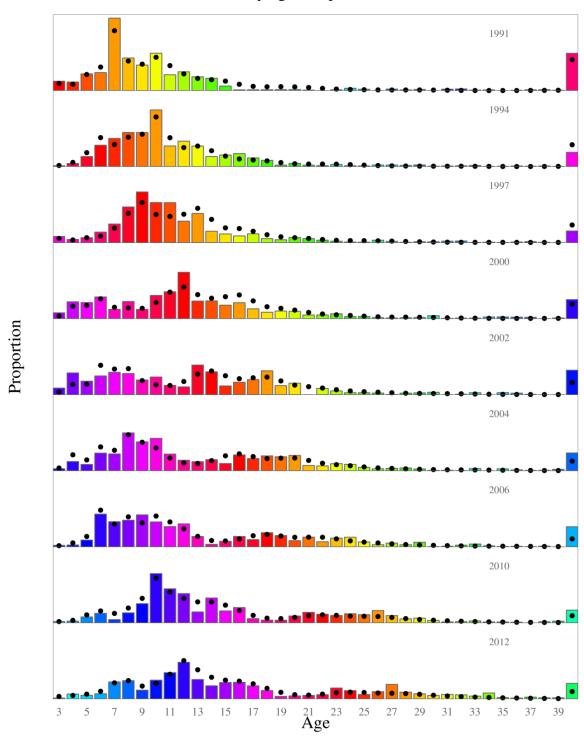
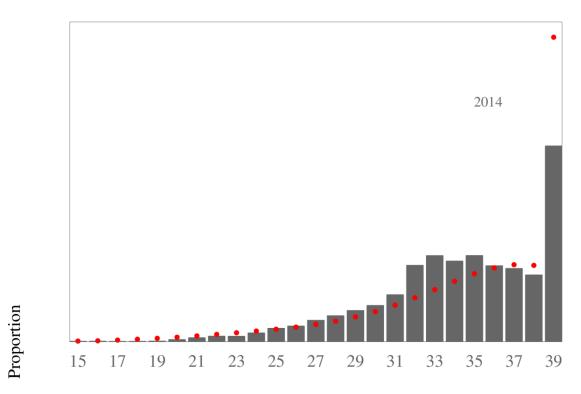


Figure 16. Model fits (dots) to survey age composition data (columns) for Aleutian Islands Pacific ocean perch, 1991-2012. Colors correspond to cohorts (except for the 40+ group).

Survey length composition data



Length (cm)

Figure 17. Model fits (dots) to 2014 survey length composition data (columns) for Aleutian Islands Pacific ocean perch.

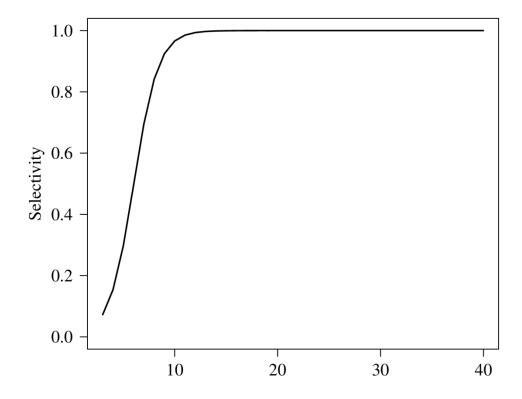


Figure 18. Estimated survey selectivity curve for BSAI POP

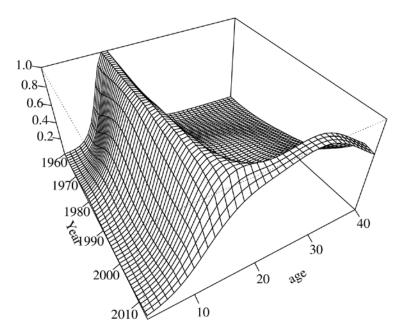


Figure 19. Estimated fishery selectivity from 1960-2014.

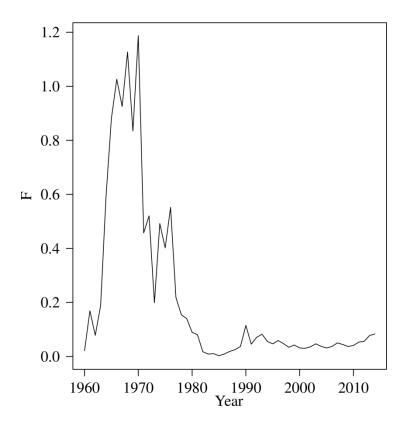


Figure 20. Estimated fully selected fishing mortality for BSAI POP.

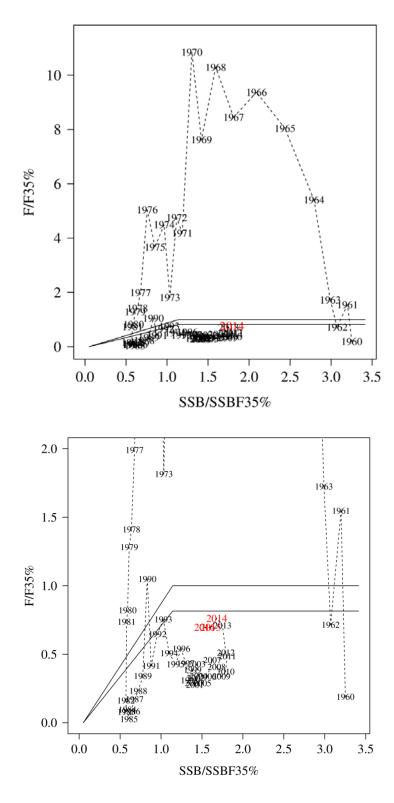


Figure 21. (Top panel) Estimated fishing mortality and SSB in reference to OFL (upper line) and ABC (lower line) harvest control rules, with 2014 shown in red. The bottom panel shows a reduced vertical scale, and the projected F and stock size for 2015 and 2016.

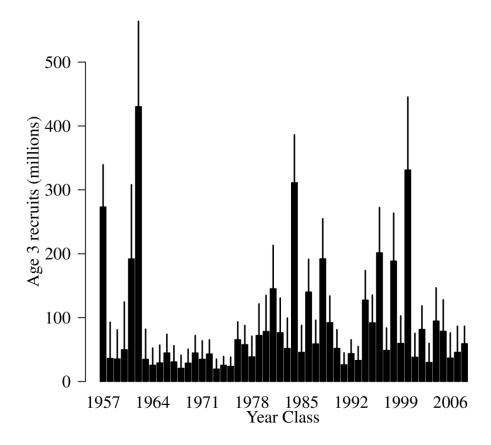


Figure 22. Estimated recruitment (age 3) of BSAI POP, with 95% CI limits obtained from MCMC

integration.

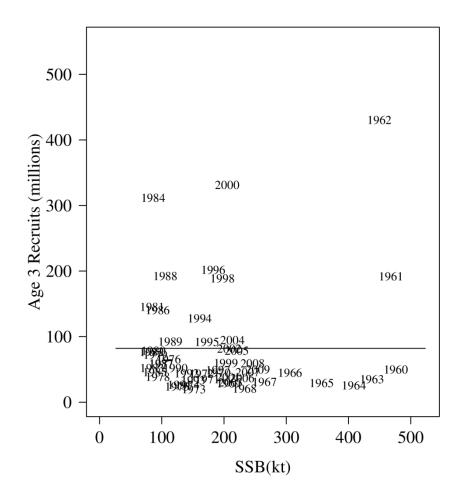


Figure 23. Scatterplot of BSAI POP spawner-recruit data; label is year class.

Appendix A. Supplemental Catch Data

In order to comply with the Annual Catch Limit (ACL) requirements, non-commercial removals that do not occur during directed groundfish fishing activities are reported (Table A1). This includes removals incurred during research, subsistence, personal use, recreational, and exempted fishing permit activities, but does not include removals taken in fisheries other than those managed under the groundfish FMP. These estimates represent additional sources of removals to the existing Catch Accounting System estimates. For BSAI POP, these estimates can be compared to the trawl research removals reported in previous assessments. POP research removals are small relative to the fishery catch. The majority of removals are taken by the Alaska Fisheries Science Center's (AFSC) biennial bottom trawl survey which is the primary research survey used for assessing the population status of BSAI POP. The amount of POP captured in research longline gear has typically been less than 0.1 t. There was no recorded recreational harvest or harvest that was non-research related in 2010 and 2011. Total removals of POP ranged between 8 and 267 t between 2010 and 2014, which was were 267 t in 2010 and 3 t in 2011, and did not exceed 1.4 of the ABC for these years.

Appendix Table A1. Removals of BSAI POP from activities other than groundfish fishing (t). Trawl and longline include research survey and occasional short-term projects. "Other" is recreational, personal use, and subsistence harvest.

| Year Source | | Trawl | Longline | Other |
|-------------|----------------------------|---------|----------|----------|
| 1977 | • | 0.008 | | |
| 1978 | } | 0.144 | | |
| 1979 |) | 3.083 | | |
| 1980 |) | 71.474 | | |
| 1981 | | 13.982 | | |
| 1982 | | 14.250 | | |
| 1983 | | 133.461 | | |
| 1984 | | 0.000 | | |
| 1985 | | 98.567 | | |
| 1986 | • | 164.541 | | |
| 1987 | , | 0.014 | | |
| 1988 | } | 10.428 | | |
| 1989 | 1 | 0.003 | | |
| 1990 |) | 0.031 | | |
| 1991 | | 76.327 | | |
| 1992 | NMFS-AFSC survey databases | 0.383 | | |
| 1993 | | 0.011 | | |
| 1994 | survey databases | 112.815 | | |
| 1995 | | 0.023 | | |
| 1996 | • | 1.179 | 0.015 | í |
| 1997 | , | 178.820 | | |
| 1998 | } | 0.006 | 0.003 | } |
| 1999 | | 0.192 | 0.014 | |
| 2000 | | 164.166 | 0.019 |) |
| 2001 | | 0.114 | 0.015 | i |
| 2002 | , | 143.795 | 0.026 |) |
| 2003 | | 7.595 | | 2 |
| 2004 | | 180.928 | 0.029 |) |
| 2005 | | 10.682 | | |
| 2006 | | 168.609 | 0.043 | |
| 2007 | | 0.063 | 0.036 |) |
| 2008 | | 21.087 | 0.037 | ' |
| 2009 | | 1.436 | | |
| 2010 | | 266.674 | | |
| 2011 | | 104.409 | | |
| | AKFIN database | 285.773 | | |
| 2013 | | 8.496 | | |
| 2014 | | 247.860 | 0.000 | <u>)</u> |